VARIATION DUE TO DIRECT AND MATERNAL GENETIC EFFECTS IN CANADIAN DAIRY GOATS

A Thesis Submitted to The Faculty of Graduate Studies and Research In partial fulfilment of requirements of the degree of Master of Science

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0-612-50904-4

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ABSTRACT

Variation due to direct genetic and maternal genetic effects in Canadian dairy goats

Derivative free restricted maximum likelihood (DFREML) was used to calculate variance and covariance components for cumulative milk yield, cumulative fat yield and cumulative protein yield, of Alpine, Toggenburg, Saanen and Nubian dairy goat breeds. First lactation records of 691 Alpine, 641 Toggenburg, 439 Saanen and 433 Nubian goats freshening between 1986 and 1995, and a combined data set of all breeds, were analysed. Two basic models, either including or excluding coefficients for phantom groups, besides the random and fixed effects, were fitted for each trait and breed. Included fixed effects were age at first kidding and flock-year for the single breed data sets and an additional breed effect for the multiple breed data set. Random effects in sub-models were fitted to estimate direct effects (model 1), direct and maternal effects (model 2) and direct effects, maternal effects and the covariance between direct and maternal genetic effects (model 3).

Solutions for fixed effects were obtained, and estimates of heritabilities (h^2) for the different breeds were between 0.17 and 0.30 for cumulative milk yield, 0.09 and 0.44 for cumulative fat yield and 0.04 and 0.25 for cumulative protein yield. There were marked differences between breeds for h^2 . Values of the maternal genetic variances (m^2) were small and statistically non-significant for all traits for Alpine, Toggenburg and Saanen goats. For Nubian the results for m^2 were very high and statistically significant (p< 0.05) for milk and fat yield and (p< 0.025) for protein yield. Maternal effects estimated for the all breed data set were not significant for milk yield, but were significant (p< 0.05) for fat and protein yield. Covariances between direct and maternal effects for all breeds and traits were not statistically significant. Fitting phantom groups in the model resulted in generally higher estimates. Best models to describe the data sets were those only fitting direct genetic effects in addition to fixed effects. The F-test revealed that fitting phantom groups had no effect in the single breed analyses, but had an effect in the multiple breed analysis.

Variations de la production laitière dues aux effets génétiques directs et maternels chez les caprins au Canada

L'utilisation du maximum de vraisemblance restreint et d'un algorithme sans dérivation (MVRSD) a été utilisé pour calculer les composantes de la variance et de la covariance des rendements cumulatifs en lait, matières grasses et protéines chez les chèvres de races Alpine, Toggenburg, Saanen et Nubian. Les analyses ont été réalisées sur les données de 691 Alpines, 641 Toggenburgs, 439 Saanens et 433 Nubians entrées en première lactation entre 1986 et 1995, ainsi que sur un ensemble de données de toutes les races combinées. Deux modèles de base prenant en compte les effets aléatoires ainsi que fixés et, incluant ou n'incluant pas les coefficients pour les groupes "fantômes", ont été ajustés pour chaque caractère et pour chaque race étudiés. Les effets fixés pour les données par races étaient l'age des individus et l'année du troupeau. De plus, pour les données combinées les effets fixés comprenaient l'effet race. Les effets aléatoires des sous-modèles ont été ajustés pour l'estimation des effets génétiques directs (modèle 1), directs et maternels (modèle 2) et, directs, maternels, avec la covariance entre les 2 (modèle 3).

Les solutions aux effets fixés ont été calculées et, les valeurs de l'estimation de l'héritabilité (h^2) pour les différentes races étaient comprises entre 0.17 et 0.30 pour le rendement cumulé en lait, 0.09 et 0.44 pour le rendement cumulatif en matières grasses, 0.04 et 0.25 pour le rendement cumulatif en protéines. Il existe donc des différences marquées de h^2 entre les races. Les valeurs des variances de l'effet génétique maternel (m^2) pour tous les caractères étudiés étaient minimes et non statistiquement significatives pour les chèvres Alpines, Toggenburg et Saanens. Pour les Nubians les résultats des m^2 étaient statistiquement significatifs pour le rendement en lait et en matières grasses ($p \le 0.5$) et pour le rendement en protéines ($p \le 0.25$), mais les estimés de m^2 étaient très élevés. Les effets maternels estimés pour l'ensemble des données combinées n'étaient pas significatifs pour le rendement en lait, mais étaient significatifs ($p \le 0.05$) pour le rendement en lait, mais étaient significatifs ($p \le 0.05$) pour le rendement en lait, mais étaient significatifs ($p \le 0.05$) pour le rendement en lait, mais étaient significatifs ($p \le 0.05$) pour le rendement en lait, mais étaient significatifs ($p \le 0.05$) pour le rendement en matières grasses et en protéines. Les covariances entre les effets directs et maternels n'étaient pas significatives pour aucun des caractères ni aucune des races. L'ajout des groupes "fantomes" au modèle a globalement augmenté les estimés. Les

meilleurs modèles pour décrire lènsemble des données étaient ceux incluants seulement les effets génétiques directs en plus des effets fixés. Le test-F a révélé que l'ajout des groupes "fantomes" n'avait pas d'effet sur les analyses portant sur une seule race, par contre il avait un effet sur les analyses portant sur plusieures races.

ACKNOWLEDGEMENTS

I would like to thank all the people who helped and supported me during my project. First of all I want to thank my supervisor J. Flannan Hayes for his support over the whole project, he gave me valuable advice and suggestions. Further more I want to acknowledge my other committee members Roger I. Cue for helping me in the theoretical aspects of genetic statistics and for the advice in the statistical model I have chosen and Robert K. Moore for transferring and preparing the test day record data and his always helpful answers, when I needed them. I would also like to thank the Quebec Dairy Herd Analysis Service for providing the dairy goat data. Special thanks to Brian Sullivan, from the Canadian Centre for Swine Improvement, for his help with the pedigree file, and to Karin Meyer from the University of New England, for providing the program 'DFREML'. Whenever I had questions they always replied fast and gave me useful information. I would like to acknowledge the financial support from my father who made it possible for me to do these studies. I would like to express my gratefulness to all graduate students at the Department of Animal Science. Special thanks for Mark Strasser, Diederik Pietersma and Ainsley Archer in the neighbouring lab for their patience and help when I needed someone to solve heavy computer problems and the friendship and patience of my three roommates Regina Holzbauer, Pilar Schneider and Josée Chicoine, who helped me with everything when I needed them. Finally I will acknowledge my friend Peter Baustädter for his support and encouragement.

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<u>1. INTRODUCTION</u>

The market situation for dairy goats in Canada is not promotional. However, there is a great potential for expansion of production to satisfy the current market. To select genetically superior animals, genetic evaluation is available to participants in milk recording programs. Replacement does are selected mainly according to dams milk production (63 %), litter size (60 %), conformation (55 %), body size at weaning (42 %), dams maternal ability (35 %). Bucks are selected according to conformation (69 %), dam/daughter milk (45 %), litter size (38 %) and other reasons such as average daily gain, sire's performance for growth, extended pedigree and breed characteristics (Nadarajah, 1998).

Dual-purpose goat breeds are dominant in many countries. In Canada, imported European breeds such as Alpine, Saanen, Toggenburg, and Nubian have been developed as single purpose breeds for milk production. In dairy goats, as in dairy cattle, profitability is dependent on the efficiency of milk production. For the development of effective selection plans, knowledge of genotypic and environmental parameters is necessary. The phenotype of an animal is the product of genetic and environmental effects. The genetic variance (effect) itself is composed of additive, maternal, dominance and epistatic genetic variance. The dam influences the phenotypic value of her offspring in two ways; first by contributing a sample half of her genes, and second by providing a maternal environment through mothering and supply of nutrients. Sire, on the other hand, contributes to the phenotype of the offspring through a sample half of his genes. The biometrical aspects of maternal effects in terms of linear genetic models were developed around 1950. Maternal effects, nowadays, are known to be present in beef cattle and sheep. Studies (e.g. Meyer, 1992, Robinson, 1996) show that between 0 to 20 % of the genetic variance is contributed by the maternal effect and that there are differences between breeds.

The development of computer programs allows the use of restricted maximum likelihood (REML) with animal models to estimate variance components (e.g. Meyer, 1989). The objective of this study was to estimate variance and covariance components due to direct and maternal genetic effects and environmental effects for milk production traits

(cumulative milk, fat and protein yield) using REML procedures under single trait animal models with data from Alpine, Toggenburg, Saanen and Nubian goats.

2. LITERATURE REVIEW

Very little research is published on goat breeding and genetics. As in dairy cattle, the main interest is to increase performance in production, reproduction and conformation traits. The literature shows that many factors, environmental and genetic, play a role in affecting lactation traits; milk yield, fat yield (milk fat content) and protein yield (milk protein content). Most environmental influences are difficult to quantify, but there are known environmental effects. They are so consistent in influencing traits, that researchers have developed mathematical adjustment factors or adjustment procedures to account for them (Bourdon, 1997).

 Table 1. Examples of Environmental Effects for which Mathematical Adjustment Procedures and
 (or) Adjustment Factors are available [from Bourdon, 1997]

Species	Trait	Environmental Effect		
Cattle (dairy)	Milk yield	Length of lactation		
	Milk vield	Milking per day		
	Milk yield	Age at calving		
	Fat yield	Length of lactation		
	Fat vield	Milking per day		
	Fat yield	Age at calving		
	Fat corrected milk yield	Fat yield		

The purpose of such adjustments is 1) to reduce sampling variance, 2) to remove systematic biases from comparison of animals or animal groups and 3) to estimate what a specific record would have been if taken at a standard age, season, length of lactation etc. (Iloeje et al, 1980).

2.1 Environmental and Genetic Causes of Variation in Production Traits

Several genetic and environmental factors affect production in dairy animals. These factors which may cause variation in productivity of dairy goats should be considered in the evaluation and calculation of genetic components. One of the most important is the herdyear-season, which includes all factors like flock environment (such as nutrition), flock management, year of birth and season of production. Year of birth is important, because there are existing phenotypic time trends, which could be genetic and/or environmental. Another important factor is age of doe at first kidding (parity) and the breed used. Studies in the United States indicate that milk yield of dairy goats is highly variable (Iloeje et al, 1981; Shelton, 1978). Effects of age, parity, and season of kidding on milk and fat yield were examined (Gipson et al, 1989; Kennedy et al, 1981), with the purpose of developing correction factors for use in genetic evaluation of breeding animals. A review of factors affecting production in goats is given by Iloeje et al, 1978.

2.1.1 Nutrition

Nutrition is the easiest to influence and the fastest responding cause of variation. All lactation traits are affected by changing food supply. Comparing dairy goats and dairy cattle on the basis of the metabolic bodyweight, the calculated requirements for energy and protein for maintenance and production are very similar to those for dairy cattle (Sutton, 1990). Studies show that a well fed doe has a higher milk production. Research conducted in 1974 with Don goats showed that milk production was phenotypically significantly correlated with body weight (Orlyanskii et al, 1974). The correlation between milk production in the first lactation and body weight was 0.33 and in the second lactation 0.43. In the literature the correlation between body weight and milk production ranged from 0.19 to 0.43. Gall (1973) showed that about 60 % of the variation in milk yield could be attributed to body size, rumen volume, skeletal size, muscle volume, and fat.

After kidding, high producing does need a high amount of energy for milk synthesis and secretion in early lactation. Does can lose up to 6 kg live weight, consisting of protein as well as fat in early lactation, slowly regaining this weight later. In early lactation stage it is very important to supply the animals with the needed requirements, otherwise the doe will not be able to produce as much milk as her genotype would allow.

If the genetic potential of a doe is to produce 4 kg of milk per day, but she only receives enough nutrients to produce 2 kg, she will only produce 2 kg. Even if just one necessary nutrient is missing she will only produce as much milk as the supply of the restricted nutrient allows. This is referred to as 'the principle of the first limiting nutrient' (Agriculture Canada, 1989).

Morand-Fehr and Sauvant (1978) stated that energy intake is the most important dietary factor influencing milk production regardless of the stage of lactation. The amount of energy really consumed by the goat appeared to be the most positively correlated factor with milk production whatever the composition of the diet may be (Morand-Fehr et al., 1978).

Table 2.Correlation between intake of metabolisable energy and dairy goat performance

	Contention between marke of methodistore energy and			
	Milk yield	Fat (%)	Protein (%)	
1st-8th week of lactation	0.752*	+ 0.010	- 0.190	
9 th –18 th week of lactation	0.794*	- 0.121	+ 0.188	
19th-28th week of lactation	0. 87 3•	- 0.157	+ 0.123	

Correlation between intake of metabolicable energy and

[from Morand-Fehr et al. 1978]

Ruminants obtain energy primarily from fibre, carbohydrates and fat in their ration. Studies show that improved energy efficiency from dietary fat increases milk production (Teh et al, 1994). The correlations between intake of metabolisable energy and trait yields shown in Table 2, indicate that milk yield increases progressively with advance in lactation and more energy is needed. In mid and late lactation fat percentage generally decreases whereas protein percentage increases.

2.1.2 Stage and length of lactation

The stage and length of lactation is important for genetic evaluation based on cumulative milk yield, because with increasing length, the yields increase too. The cumulative milk, fat and protein yield is the actual amount of milk that a doe gives in her lactation up to a certain point and the longer a goat is in lactation the higher is her milk yield per lactation, all other things being equal. Within species and within breed, it is the stage of lactation that has the greatest influence on milk composition (Haenlein, 1995). To get unbiased results it is important to make sure that only goats with a complete lactation participate in the evaluation. The lactation curve of goats looks similar to that of dairy cattle and dairy

sheep. The lactation starts with a high increase in milk yield and reaches a peak after 30 days for sheep and around 45 days for goats and then declines until next kidding. (LKV, 1993; Sutton, 1990).



Figure 1. Average lactation curves of German sheep and German 'Alpine' goat populations [from LKV, 1993]

In literature reports, the averages of lactation length for dairy goats, depending on breed, are from 211 to 238 days (Ali et al., 1983). Ali et al (1983) and Kennedy et al (1982) reported a strong relationship between lactation length and lactation yield.

Sullivan (1988) concluded that if lactation length is determined mainly by environmental factors, there would be a large bias by failing to adjust for differences in lactation length and if lactation length is heritable, it would be inappropriate to simply adjust yield for lactation length.



Figure 2. Protein and fat content of goats and sheep milk during lactation [from LKV, 1993]

Many components, especially fat and protein percentages, are high in ewe and goat colostrum, much lower thereafter in milk, and they rise again at the end of lactation (Anifantakis et al., 1980). Fat content in goat milk changed from 2.7 % in mid lactation to 4.6 % during the last week (42) of lactation, and protein content from 3.0 % to 4.2 % (Voutsinas et al, 1990). Similar trends were observed from the LKV Germany, (1993), which is illustrated in Figure 2.

2.1.3 Season of kidding

In Canada goats are like sheep in that they are highly seasonal in their breeding. Singh, Acharay and Biswas (1970) have found season of kidding to be a very important source of variation affecting milk production. Strong relationships between reproduction and production indicate that photo period may have an indirect effect on production traits (Ali et al., 1983; Kennedy et al., 1982). In the U.S. goats seem to be seasonal breeders (Shelton, 1978; Ali, et al., 1983; Mohammad, 1984; Amoah et al., 1996) and, as in other septentrional countries, related to length of photo period. Light introduces the fertile oestrous cycles, which generally begin 10 weeks after the longest day of the year (Agriculture Canada, 1989). This was also reported by BonDurant et al (1981) where the goats showed an endogenous annual rhythm of biological activity, which was responding to day length changes. Seasonal variation in photo period increased with increasing distance from the equator leading to the conclusion that breeding in dairy goats is more seasonal at northern than at southern latitudes because of differences in length of photo period (Ali et al, 1983; Mohammad, 1984).

On the other hand, some studies showed that the effect of latitude on reproductive performance is minor (Mohammad et al, 1984) and that even in the most northern regions of U.S., breeding in goats is not highly seasonal. The results suggest that factors other than length of photo period may be causing the seasonality of breeding of goats in the U.S., in particular, the interaction of feeding regimes, reproduction and management. The effects of season of kidding, lactation yield, temperature and humidity (climate) could cause variation in fat and protein content. It is reported that temperature (Johnson, H.D., 1965; McDonald et al, 1958) and humidity (Ingraham et al., 1979; Johnson J.C. et al.,

1962) influence the animal directly and cause adjustments in their behaviour which result eventually in a decline in milk energy output (Iloeje et al, 1980). However, artificial manipulation of photo period alters seasonality of reproduction and has been known for years as a way of controlling the breeding season (Henderson D., 1985).

Generally breeding season begins in early September and ends in February or March in the northern hemisphere. Amoah (1983) observed that there was negligible ovarian cycling



and reproductive activity in does between April and August. This observation corresponds with the breeding season, reported in the UK (Henderson D., 1985) and the US. Mohammad et al (1984) observed that the median month of conception was October and the five breeds (Alpine, Toggenburg, Saanen, Nubian and LaMancha) kidded most frequently in March. Similar findings are reported by Grossman and Wiggans (1980).

Results on seven dairy goat breeds (Amoha et al., 1996) show that most breeds start their seasonal breeding, approximately in June, reaching a peak in September to November. A small number of goats, about 5 % will cycle 1-2 months earlier or later than their flock mates (Agriculture Canada, 1989). Good nutrition will allow these animals to express this tendency, but this trait may also be genetic. Differences between breeds are also observed. Alpine and Toggenburg appear to be the most seasonal breeds and Nubians are less seasonal than Alpine, Saanen or Toggenburg (Mohammad et al., 1984). Another study also found that Nubian goats have an extended (8 to 11 month) breeding season (Amoah et al., 1996) which could be due to the lack of sensitivity to climate changes through the year, because of their origin in Africa. Further, it appears that interactions between age and season are significant for fat yield and approached significance for milk (Alderson,

1980) and that year of kidding, month of kidding and their interaction have a significant effect on milk production and lactation length (Mavrogenis et al., 1984). The season of freshening affects younger does more than their older flock mates (Iloeje et al., 1980). Studies showed that does which kidded earlier in the season (January through March) produced a higher milk yield than their flock mates that freshened later (Steine, 1975; Iloeje et al., 1980).

There are also clear seasonal differences in milk composition of the major and minor components (Renner, 1983), but these are confounded with climatic and dietary effects. Winter climate can affect milk yield and composition, and both are negatively correlated. Winter feeding usually provides different proportions and qualities of grazing, hay, silage and supplements, which influence milk composition considerably (Haenlein, 1995).

2.1.4 Age at first kidding

The fertile oestrus, or heat, of a doe starts as early as in her fourth month of age. Research on dairy cattle concluded (Iloeje, 1980) that age has no direct biological effect on production, but body weight does. In other words, age is an indicator of maturity and maturing has a biological effect on production since a growing animal needs a part of its energy intake for growth and development.

Many goat breeders recommend a body weight at first kidding of at least 32 kg. Does may reach this weight by 7-9 months (Sutton, 1990; Agriculture Canada, 1989). To some extent, breeding maturity (32 kg bodyweight) is dependent on genetic factors, but to a larger extent on management and feeding. Bred within 7-9 months, does would freshen, after a 5 month gestation period (~ 150 days), at 12 - 14 month of age. If does continue to grow during their first lactation, their milk production will be steady, but not as high as the milk production of fully grown does. For example, does bred at 18 months of age and freshening for the first time at 2 years of age milk less heavily compared to a second freshening two years old, but compared to a first freshening 1-year-old doe, the older doe will have a higher milk production (Agriculture Canada, 1989). Research in dairy goats, found that age is a very important source of variation affecting milk production. Mavrogenis (1984) demonstrated that age of dam at kidding had a quadratic relationship with all traits studied except for lactation length. Iloeje (1978) and Rathore (1970) have shown that milk yield increases linearly with age until about the third to fifth lactation and tends to decrease in the following lactations. Expressed in age, does reach peak production at 24 to 50 months of age.

Results of a study made by Sullivan (1988) match with these from the literature (Finley et al. 1984; Kennedy et al, 1981), except that the quadratic effect of age was not significant. Solutions indicate that the effect of age on all traits in first lactation was initially large and positive, but quickly diminishes and eventually becomes negative after two years of age. A plot of the age solutions for first lactation milk yield from the study made by Sullivan (1988) is given in Figure 4.



Figure 4. Age solution curve (solid line) and unadjusted means (•) by age subclasses for first lactation milk yield. [from Sullivan, 1988]

2.1.5 Parity

Within a lifetime production a doe has a certain number of parities and her maximum production occurs between 24 and 50 months, which equals second to fourth parity, with a mean of about 40 months. Parity effects on production in dairy goats are large and significant (Ali et al., 1983; Finley at al., 1984; Kennedy et al., 1981). Kennedy et al (1981) calculated a set of factors for four groupings and showed that parity affected production. Differences between first and second parities of animals of the same age were as high as 217 kg of milk and 7.2 kg of fat. These amounts are about 20 % of the production for first parity. Differences among later parities were less.

Differences due to parity, lactation number or age of animal can be significant in gross milk composition, but this is also confounded with milk yield level. In sheep, the fat content of ewe milk changed linearly from the 1^{st} to the 6^{th} parity from 6.8 to 7.4 % and total protein content from 5.8 to 6.2 % (Casoli et al, 1989). Similar trends can be expected for goats. The cause of parity differences and the negative relationship between age and production within parity is unclear and could be genetic or environmental (Kennedy et al., 1982). Finley et al. (1984) recommends that the large effects such as the one between first and second parities should not be ignored.

2.1.6 Breed comparison for milk production

In modern animal breeding, a "breed is conceived as a defined population, in which pure breeding is the rule and breeding animals are registered by a breeding organisation" (Gall, 1996). Within goat breeds, different types of goats have been developed or have evolved to serve specific functions or for adaptation to specific production or environmental conditions (Shelton, 1978). Some breeds developed for dairy purposes, others to produce milk and meat or just meat (Table 3).

Туре	Breed or genotype	Localisation	Source		
Dairy	Saanen	Temperate regions (world wide)	Guss (1975), Colby et al.		
	Toggenburg	Europe and U.S.	Guss (1975), Colby et al.		
	Anglo-Nubian	World-wide	Guss (1975), Colby et al.		
	Alpine	Europe and U.S.	Guss (1975), Colby et al.		
	LaMancha	U.S.	•		
Meat and milk	Nubian	Sudan	Devendra, (1975)		
	Damascus (Shami)	Syria and adjoining areas	Choveiri, (1973)		
	Jamnapari	India and Pakistan	Singh and Singh. (1974)		
	Barbari	India and Pakistan	Mittal and Pandey, (1971)		
Meat	Boer	South Africa	Skinner. (1972)		

Table 3. Types and breeds of goats [from Sheiton, 1978]

Genetic differences of production yields within breeds have a wide range and numbers for milk, fat and protein yields are presented as mean values of the breeds as reported in the literature.

The most popular goat breeds used in Canada for milk production are Alpine, Toggenburg, Saanen and Nubian. All dairy breeds (Alpine, Saanen, and Toggenburg) have nearly the same milk yield per lactation, Nubians are lower in production. A possible explanation might be that this breed is a result of cross breeding where dual-purpose breeds were involved. Goat milk composition can differ greatly among breeds. For example, for milk fat from 2.3 % to 6.9 % with an average of 3.3 %; for milk protein from 2.2 % to 5.1 % with an average of 3.4. % (Juarez et al., 1986). A major portion of this variation is caused by negative correlation's between milk yield and composition, which means that low yields have higher contents and vice versa. (Haenlein, 1995).

The *Alpine* is a breed of goat that originated in the Swiss / French Alps. It is a medium to large goat and more variable in size than the Swiss breeds. Alpine females are reported as excellent milkers; milk yield is around 950 kg with 3,5 % fat (~33.5 kg) in a 259 days lactation period. (Sambraus, 1994; Gall, 1996)

The *Toggenburg* is a Swiss dairy goat from the Toggenburg valley of Switzerland. They are thought to be oldest known dairy breed (since 1802). This breed is of medium size and slightly smaller than other Alpine breeds. They give their best performance in cooler conditions and they are noted for high milk production. Milk yield of 700–1000 kg with 3,3 % fat (~21.2 kg) in 267 days is common. (Sambraus, 1994; Gall, 1996; Haenlein, 1996)

Saanen is probably the most developed dairy breed. Among goat breeds it occupies the place that the Holstein-Friesian has among cattle breeds. The Saanen dairy goat originated in Switzerland in the Saanen valley. It is of medium to large size with a milk yield of 750 to 1000 kg and 24.4 kg fat in 262 days. (Sambraus, 1994; Gall, 1996; Haenlein, 1996)

The *Nubian* goat breed developed in England by crossing British and Swiss goats with Zaraibi, Jamnapari and Chitral goats from Pakistan. They are known as all-purpose goats, useful for milk, meat and hide production. It is a relatively large goat and described

as a not heavy milk producer (774 kg), but this breed has a very high fat yield with 4-5% (~35.4 kg). The average lactation length is 237 days, but, as mentioned earlier, their breeding season is much longer than that of the Swiss breeds and so it is possible to produce milk year round. (Sambraus, 1994; Gall, 1996; Haenlein, 1996)

Breeds differences are well reported under different conditions. Knowls and Watkins (1938) tabulated some breed comparisons under English conditions. Garcia, Castillo and Gado (1972) reported lactation yields under Venezuelan conditions. Gill and Dev (1972) reported lactation yields for Alpine and Nubian goats under Indian conditions and Dickinson and King summarised milk yield and fat content of U.S. dairy goats (Table 4.)

	Current world	England *		<u>.</u>	<u>U.S.</u> ^b		<u>Vene-</u> zuela [°]	India ^d
Breed	record	Milk	Fat	Milk		Fat	milk	milk
	(kg)	(kg)	(%)	(kg)		(%)	(kg)	(kg)
Saanen	3430	1188	4.0	979		3.6	294.2	-
Alpine	2194	1136	4.2	9 7 0		3.5	232.2	310.6
Toggenburg	2613	1087	4.5	921		3.3	283.4	
Nubian	2009	839	5.6	817		4.5	154.7	289.5

^a Knowles and Watkins, 1938

^b Dickinson and King, 1977

^c Garcia et al., 1972

^d Gill and Dev. 1972

In all four regions, Saanen had the highest milk production. Alpine and Toggenburg had similar high yields. Nubian does were lower in milk production, but had the highest fat percentages in the milk. Sutton (1990) and Iloeje (1980) made a detailed analysis of breed differences in milk yield, fat yield and fat %. Geerts (1975) summarised the performance of does of the five major dairy breeds.

The description of breeds of goats and their origin can be found in Haenlein (1981) and on the Oklahoma University Animal Science homepage. Later publications of the status of dairy goats in the U.S.A. can be viewed in Haenlein (1996).

2.1.7 Others

A healthy dairy goat population will give a reasonable milk yield. Diseases caused by infection, parasites, nutritional and other reasons endanger the milk production of does. This variation appears because the animals are not able to make full use of the nutrients given in the ration (Agriculture Canada, 1989; Waite et al., 1963). Fortunately these factors, which influence the production, are reported in milk recording data, and it is possible to adjust genetic evaluations for environmental factors which are known. As it happens, there are further effects, but when not recorded, we can not account for them in a statistical analysis. Besides the already mentioned causes of variation, there is, for example, in sheep evidence that within the limits imposed by inherent ability and level of nutrition, milk production and the shape of lactation curve are affected by the suckling stimulus of the lamb or lambs. (Peart, 1968; Zygoyiannis and Katsaunis, 1984; Zygoyiannis, 1994). Also the number of kids born by a doe has been investigated. Results for a Research Institute flock indicate that mothers of twins gave proportionately 0.27 more milk than mothers of singles and mothers of triplets 0.47 more, and it was suggested that multiple foetuses led to a higher lactogenic activity, that results in higher milk yield post partum (Hayden, Thomas and Forsyth, 1979). More recent evidence from Spain also shows raised milk yields after multiple births (Subines, Lava, Ferrando and Boza, 1988). Williams, 1993 on the other hand could not find any evidence that the number of kids born affected either scale or persistency, which contradicts the work reviewed earlier. These results might be due to the husbandry of the goats surveyed, because housing and conserved forage and concentrate food were offered year around.

Some researchers mention that the number of milkings per day have a significant effect on milk production. Henderson et al (1983,1985) reported that thrice-daily milking increases milk yield immediately and in the long term. The immediate increase is believed to result from the more frequent removal of a chemical feedback inhibitor, which is present in the milk (Henderson and Peaker, 1984) and the long term increase seems to be caused by either growth or reduced regression of the thrice-daily milking gland (Henderson et al., 1985).

2.2 Components of an Individuals Performance

For animal breeders, the most important economical characteristic of an animal is its production or performance. For each performance trait we can observe or measure a certain level, which is called the phenotype. An animal has as many phenotypes as there are traits to be observed or measured on this animal. In selection it is important to increase the phenotypic value which is possible through genetic and / or environmental improvement. The basic mathematical formula for expressing the phenotype is:

$$\mathbf{P} = \mathbf{G} + \mathbf{E}$$

where **P** stands for an animal's phenotype, **G** represents it's genotype and **E** the environmental effects (Bourdon, 1997). Through measuring the phenotype and adjusting for environmental effects, (these are all non-genetic effects) we can estimate the genotype. Accurate information on environmental factors helps obtain accurate estimation of the genotypic variance.

2.2.1 Variance Components

Variation is the source of genetic change. If there is little variation in a trait, selection will be slow, because no animal is really genetically superior to another. To make genetic progress it is important for a population to have a wide range of variation. The amount of variation is measured and expressed as the variance. The total variance is the phenotypic variance (V_P) which is the sum of the genotypic (V_G) and the environmental (V_E) variance. The genotypic variance itself can be divided into additive (V_A) , dominance (V_D) and epistatic (V_I) variance. The total variance is then:

$$\mathbf{V}_{\mathbf{P}} = \mathbf{V}_{\mathbf{G}} + \mathbf{V}_{\mathbf{E}}$$
$$= \mathbf{V}_{\mathbf{A}} + \mathbf{V}_{\mathbf{D}} + \mathbf{V}_{\mathbf{I}} + \mathbf{V}_{\mathbf{E}}$$

Additive variance is the main cause of resemblance between relatives and therefore the chief determinant of the observable genetic properties of the population and of the response of the population to selection. It is also referred to as the direct genetic variance.

Estimates of additive variance are needed to estimate genetic parameters such as heritability, repeatability etc.

Dominance variance is caused by within locus interaction. For an accurate estimate, a complete pedigree with many different relationships is needed.

Epistatic variance (V_I) occurs if the genotypes at different loci show epistatic interaction, but interactions involving large numbers of loci usually contribute so little variance that they can be ignored.

Environmental variance is defined as all variation caused by non-genetic factors. It can have many sources and its nature depends very much on the character and on the animal studied. In other words, environmental variance is a source of errors that reduces precision in genetic studies. Most common external causes are nutritional, managemental and climatic factors (Falconer and Mackay, 1996).

Partitioning the variance into its components allows us to estimate the relative importance of various determinants of the phenotype. In other words, we can determine the role of the components in heredity and environment. Estimates of phenotypic and additive genetic variances are needed for designing effective breeding programs. Iloeje et al (1981) reported variances for flock, year-season, sire, doe, and residual for five US dairy breeds (Alpine, Saanen, Toggenburg, LaMancha, and Nubian). Flock effects accounted for 22 to 31 % of the phenotypic variation for milk yield, 24 to 25 % for fat yield and 15 to 25 % for fat percentage. These results indicated large differences in production from feeding and overall flock management. Year-season accounted for 8 to 13 %, 8 to 14 % and 10 to 14 % of the total variation for milk yield, fat yield and fat percentage respectively. Sire components contribute 8 to 11 % of variation in these three traits. Because of the usually small sizes of goat flocks and only a few does per sire, it is hard to distinguish between effects of sire and year-seasons; therefore, some variation attributed to sire may be from confounding of sire effect with other effects. Doe component was 16 to 25 % of the total variation in milk yield, fat yield and fat percentage, which suggests that sire component was overestimated. Residual variances were about 34 % in milk and fat yield and 40 % for fat percentage. Sullivan (1988) working with Canadian dairy goats used two methods,

Schaeffer's and Van Raden's, to estimate variance components. The results are given in Table 5.

SCHAEFFER'S METHOD

TABLE 5.	Estimates of Flock-year-season, sire, and residual components of variance
	(from Sullivan, 1988)

Component	Milk	Fat	Fat	Protein	Protein	Lactation	
(N)*	yield	yield	%	yield	%	Length	
Flock-year-	9837	13.7	.031	9.7	.0062	997	
scason (730)	(2098)	(3.1)	(.009)	(1.9)	(.0026)	(150)	
Sire	598 0	6.1	.032	4.7	.0050	443	
(933)	(1709)	(2.3)	(.008)	(1.4)	(.0017)	(168)	
Residual	46810	62.8	.206	44.1	.0630	3816	
(1979)	(2249)	(2.7)	(.010)	(1.6)	(.0016)	(149)	
VAN DADEN'S METHOD							
Component	Milk yield	Fat yield	Fat %	Protein	Protein	Lactation	
<u>(N)</u>				yield	<u>%</u>	Length	
Flock-year-	9626	14.3	.034	9.8	.0063	1119	
season (730)	(1817)	(3.0)	(.008)	(1.8)	(.0023)	(139)	
Sire	3919	3.8	.037	2.9	.0060	383	
(933)	(1687)	(2.7)	(.007)	(1.3)	(.0017)	(169)	
Residual	48140	64.0	.202	45.2	.0621	3805	
(1979)	(2123)	(2.7)	(.010)	(1.7)	(.0014)	(141)	

* N= number of levels of flock-year-seasons or sires or number of residual degrees of freedom

The flock-year-season component of variance ranged from 8 to 20 % of the total variance, which was similar to the results of Iloeje (1981). In Sullivan's study, percentage traits were less variable across flock-year season than were yield traits, but this varied depending on breed. In general, fat percent was more variable than milk and fat yields across year-season.

2.2.2 Heritability

Each somatic cell in a goat's body contains 30 pairs of chromosomes; One half of each pair is inherited from its sire and the other half from its dam. Each chromosome carries a certain number of genes and the nature and position are controlling genetic characteristics. During formation of the ovum or sperm cells this genetic material is resorted. This gives each germ cell 30 chromosomes, each with it's required number of genes, but in a combination, random from the original pair. These random combinations make genetic prediction incredibly complicated and we cannot predict the goat's characteristic with total confidence.

A superior doe can produce good quantities of milk persistently through her total lactation, and for many lactations. Such doe can pass these qualities to her offspring. Some traits are more easily improved through breeding than others, and are spoken of as having high heritabilities (Agriculture Canada, 1989). Heritability can be defined as "the fraction of total phenotypic variation attributable to genetic differences, and measure the accuracy of phenotype as an expression of genotype" (Cue, 1998). In algebraic terms, if $\mathbf{p} = \mu + \mathbf{g} + \mathbf{e}$, where \mathbf{p} is a phenotypic record on an animal, μ is the population mean, \mathbf{g} is the effect of genotype and \mathbf{e} is the effect of environment such that $\sigma_{\mathbf{p}}^2 = \sigma_{\mathbf{g}}^2 + \sigma_{\mathbf{e}}^2$, then heritability (\mathbf{h}^2) is defined as:

$$h^2 = \sigma_g^2 / \sigma_p^2$$
 assuming no covariance between g and e.

Heritability can be estimated in a broad and in a narrow sense. In the narrow sense, σ_g^2 contains only additive genetic variance. In addition to the additive genetic variance, heritability in the broad sense, σ_g^2 contains dominance and epistatic genetic effects. Theoretical limits of heritability estimates are $0 \leq h^2 \geq 1$. Results on heritabilities of milk yield are in a similar range for dairy cattle and goats (Auran, 1976; Iloeje et al., 1978; Steine, 1975).

Table 6. Heritability	estimates f	'or dairy goa	t breeds
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Traits	Breed	h ²	Method	Source
Milk yield	ASTNL*	.25	BLUP	Boldman et al, 1984
Average yield per operational year	Norwegian	.55		Ronningen, 1965
for milk yield	-			
Butterfat %		.22		
Milk yield (morning yield)		.40		
Butter fat % (daily yield)		.10		
1 st lactation	Indian Beetal goat	.32		Prakesh et al. 1971
2 nd lactation		.29		
3 rd lactation		.32		
4 th lactation		.28		
5 th lactation		.16		
Milk production	ASTN *	.17 ± .20		Garcia, 1971
Fat production		.22 ± .20		
milk yield in first lactation	Indian Beetal goat	.25 ± .08		Singh et al, 1970
Unadjusted records	ASTN •	AST* N*	MINQUE	Kennedy et al, 1982
Milk yield		.69 .30		
Fat yield		.62 .48		
Fat %		.52 1.10		
Adjusted records	ASTN •	AST* N*	MINQUE	Kennedy et al. 1982
Milk yield		.68 .35		
Fat yield		.61 .54		
Fat %		.54 1.09		
Milk yield (per lactation)	Alpine	.60		Bouillon et al, 1976
Protein yield (milk)		.47		
Protein %		.58		
Fat yield		.47		
Fat %		.48		
Milk yield per lactation	Dairy breeds	.36		Ronningen, 1967
Fat yield		.30		
Fat %		.32		
Single-trait analysis	Murciano-		DFREML	Analla et al, 1996
Milk yield	Granadina	.18		
Fat content		.16		
Protein content		.25		
Multiple-trait analysis	Murciano-		DFREML	Analla et al, 1996
Milk yield	Granadina	.17		
Fat content		.14		
Protein content		.22		

• A=Alpine, S=Saanen, T=Toggenburgs, N=Nubian, L=Lamancha

Heritabilities of dairy goats are summarised by Shelton (1978) and in Table 6. Estimates of heritabilities of traits can vary significantly from study to study. This depends on breed, population sampled, environmental and management conditions and error, both random and systematic, in the estimation procedures. The wide range of heritabilities reported would suggest that milk production has a medium heritability and fat and protein percentage have a high one.

2.2.3 Direct and Maternal Effects

Production traits such as milk yield, birth weight and early growth rate are determined not only by the animals' own genetic potential, but also by maternal effects. The maternal effect of a dam on her offspring is mostly environmental (with respect to the offspring) and represents mainly the dam's milk production and mothering ability, though effects of the uterine environment and extra chromosomal inheritance may contribute (Meyer, 1992). However, these abilities such as milk production and mothering are expressed according to genotype. The phenotypic differences among dams for the maternal effects are expressed in the phenotypic values of their young; i.e. offspring from good mothers have a better development and produce better later on.

A maternal effect defined by Willham (1972) is "a phenotypic value of a dam measurable only as a component part of her offspring's phenotypic value'. These influences of a dam measurable on her offspring are recognised as a special case of the joint action of genotype and environment (Robison, 1981).



Figure 5. A path coefficient diagram showing the biometrical relationship when a direct and maternal effect are involved in the phenotypic expression of a trait [from Willham, 1972]. G_{0X} : genotypic value of X for direct effects; E_{0X} : environmental value of X for direct effects; G_{MW} : genotypic values of W for the maternal effects; E_{MW} : environmental values of W for the maternal effects; G_{0X} : genotypic value of X for direct effects; G_{MW} :

Dickerson (1947), Koch and Clark (1955) and Kempthorne (1955) developed the biometrical aspects of maternal effects. Later, Willham (1963) put the biometrical aspects

in terms of linear genetic models. Falconer (1965) developed a genetic model in which the maternal effect was linearly related to the phenotype of the dam. Van Vleck (1971) devised selection index procedures for direct and maternal genetic components of traits (Willham, 1972).

The genetic principles behind the maternal effects are easiest to describe with a path coefficient diagram developed by Willham, 1972 (Figure 5):

 P_X represents the phenotypic value of individual X. For this model we have to assume only additive genetic effects and that covariances between genetic and environmental effects and between environmental effects (E_{MW} , E_{OX}) are zero. Under this assumption P_X is just influenced by the genotypic value of X and by the genotypic value of its dam (W). Effects denoted by O are the direct effects and denoted by M are maternal effects. Then

$P_x = P_{Ox} + P_{MW}$

For single traits, the phenotypic value is composed of the sum of genotypic variance and environmental variance accordingly

Here G_{0x} and E_{0x} are the genotypic and environmental values for the direct effects and G_{MW} and E_{MW} are the genotypic and environmental values for the maternal effects. The total variance of P_x is then composed of:

$\mathbf{V}(\mathbf{P}_{\mathbf{X}}) = \sigma^2_{\mathbf{Go}} + \sigma_{\mathbf{GoGm}} + \sigma^2_{\mathbf{Gm}} + \sigma^2_{\mathbf{Eo}} + \sigma^2_{\mathbf{Em}}$

Unlike sires, which contribute just through transmitting genes to the offspring, dams make a 'direct' genetic contribution to the offspring through contributing genes and an 'indirect' genetic contribution through the environment she provides for the young (milk quality and quantity, nest structure, care of young etc.). This indirect contribution is genetic to the extent that these maternal qualities are heritable in the mother, even though experienced by her offspring as an environmental effect (Riska et al, 1985).

Research on maternal effects has been conducted mostly on beef cattle, meat sheep, swine or mice, where researchers and producers are convinced that maternal effects play a role. The little work done on dairy cattle gave contradictory results.

Meyer (1992) estimated variance components for birth, weaning, yearling and final weight in Australian Hereford, Angus and Zebu cross cattle by Restricted Maximum Likelihood. Using six models, significant maternal effects were found in all analyses except for final weight in Angus. Fitting a permanent environmental effect increased model fit markedly and identified a significant maternal effect contributing 8.3 (Angus) to 10.1 % (Hereford) of the total variation in birth weight. Estimates of the direct, maternal and total heritabilities were well within the range of estimates reported in the literature summarised by Meyer, 1992. Maternal genetic effects accounted for approximately 4 % (final weight) to 14 % (weaning weight). Genetic covariances between direct and maternal effect were essentially zero to very small (0.128 to 0.758 kg²) for birth weight, -203.1 to 14.7 kg² for weaning weight depending on breed, with Angus having a positive value. Covariance for yearling weight was -66.4 to 45.6 kg², where again Angus had the positive value. For final weight, covariance's between direct and maternal effects were slightly lower and ranged from -6.7, 3.0 to 41.2 kg² for Hereford, Angus and Zebu Crosses. Variance components estimated by Robinson (1996) were slightly smaller. Maternal genetic effects were 7, 9 and 6 % for birth, weaning and yearling weights, respectively. Maternal effect for final weight was not significant. Van Vleck et al (1996) estimated direct and maternal effects for weaning weight of calves of dams at 2, 3 years of age and older with a derivative-free REML algorithm. For the three analyses for pure breeds, average proportions of phenotypic variance were 34, 31, and 27 % for direct genetic; 16, 15, and 12 % for maternal genetic; and 18, 20, and 17 % for maternal environmental effects. For composite breeds, average proportions of phenotypic variance were 44, 46, and 36 % for direct genetic; 6, 6 and 5 % for maternal genetic; and 16, 14, and 14 % for maternal environmental effects. A study by Eler et al. (1995) on growth traits of Nelore Cattle in Brazil, showed that maternal variance accounted for 12, 13 and 10 % of the total phenotypic variance and direct additive genetic variance accounted for 22, 14, 16 % of the total variance for univariate analysis of birth, weaning and yearling weight. Covariances between genetic and maternal effects were -0.49 for birth weight, 46.38 for weaning weight and 75.45 for yearlings' weight.

Research, using least square procedures, was undertaken to investigate maternal abilities on cross breeding sheep. A study made by Vesely et al. (1977) was designed to estimate heterosis, general specific combining abilities, maternal and sex-linked effects based on a diallel mating system among four breeds of sheep.

Maternal effects were measured only among progeny of crossbreeding type of mating and were found to be an important factor (p<0.01) influencing body weight, weaning weight, weight-per-day-of-age and final weight.

Research on crossbreeding animals (two-breed cross, back cross and 3 breed cross progeny) by Holtmann et al. (1969) showed an interaction between mating system and maternal effects attained significance with weights at 28 and 120 days and daily gain. Lambs raised by 3-year, 4-year and 5-year old ewes had highest body weight at 28 and 120 days of age, indicating that these age groups appeared to be at optimum in terms of the mothering ability of a ewe.

A similar study (Rastogi et al., 1982) also showed that maternal ability is a significant source of variation for birth weight, preweaning weight, preweaning average daily gain and weaning weight.

Direct and maternal effects for growth traits of Romanov sheep were estimated by Maria et al., (1993) using restricted maximum likelihood with an animal model, which included fixed effects for year x season, sex, rearing type and litter size, and random effects for the direct genetic effect (h^2) , the maternal genetic effect (m^2) , the permanent environmental effect (c^2) and random residual errors. The estimates for h^2 , m^2 and c^2 were 0.04, 0.22 and 0.10 for birth weight; 0.34, 0.25 and 0.0 for weaning weight; 0.09, 0.01, and 0.07 for 90 day weight; 0.26, 0.17, and 0.02 for preweaning daily gain (birth to wearing); and 0.15, 0.01, and 0.03 for postweaning daily gain (weaning to 90 days) respectively. These estimates with sheep are similar to those with beef cattle and indicate that important maternal effects in dairy cattle have been studied by Van Vleck and Bradford (1966), Van Vleck and Bradford stated that the apparent pattern for heritability estimates for deviations is that the paternal half-sib correlation is about the same for all lactations but that daughter-dam

estimates are higher for first lactation records, with a gradual drop in second and third lactation to the level of estimates derived from paternal half-sib analyses. This result suggested a sizeable genetic maternal effect in first lactation, which drops in second lactation and disappears by third lactation. On the other hand Van Vleck and Hart (1966) regressed the actual covariance on the coefficient of additive direct genetic variance which gave a correlation of 0.92. These results suggested that in this research only additive genetic effects were important for first lactation. Gipson and Russel (1978) showed deviations of Ayrshire and Jersey from Holstein of 1574 and 1080 kg of milk. Of this, 61 and 72 % respectively were due to maternal effects. Robison et al (1981) analysed crossbreeding data, in which the model fitted included breed additive direct, heterotic and breed maternal effects. The deviation of Swiss and Ayrshire from Holstein were 1333 and 1366 kg, respectively. Approximately 36 and 44 % of these deviations were due to breed maternal effects. These results suggest that maternal effects play a large role in breed differences. No research on maternal effects with dairy goats has been reported. An understanding of the genetic variation in maternal effects and the relationship between direct genetic effects and maternal effects is essential for formulating optimum breeding programs (Robison, 1981).
<u>3. MATERIAL AND METHOD</u>

The final cumulative milk, fat and protein yields of the first lactation of Alpine, Saanen, Toggenburg, and Nubian does were used to estimate:

- 1. Basic statistical parameters
- 2. Variance components such as additive genetic effects, maternal genetic effects, and covariance between additive genetic and maternal genetic effects
- 3. Resulting genetic parameters

Basic statistical parameters were calculated using SAS for OS/2, release 6.12 and variance components were estimated using the DFREML package version 2.1 for the derivative-free REML algorithm (Meyer, 1993).

3.1 Data Source

A total of 88965 test day records from all registered breeds in Canada, recorded from 1985 to 1995, were available through the Quebec Dairy Herd Analysis Service (QDHAS). As with dairy cattle, monthly measurements of milk yield are taken, and analysed for fat and protein content. These monthly measurements are used to estimate cumulative milk, fat and protein yields. The information included in the test day records is: flock, QDHAS goat number, testing program (official or self testing), breed, goat identification number, nip letter of goat, sire identification number, dam identification number, nip letter of dam, date of birth, lactation number, date of kidding, date of testing, lactation code, test day milk, test day fat, test day protein, 305 days milk yield, 305 days fat yield, 305 days protein yield, cumulative milk yield, cumulative fat yield, cumulative protein yield, handling code.

Genetic evaluation for dairy goats in Canada is undertaken by the Canadian Centre for Swine Improvement from which the pedigree file, containing registered goats from all breeds, was received. The file contained 108545 records. Information included in the pedigree was goat, sire and dam identification number, sex of the animal, name, foreign identification number and date of birth. The pedigree was used to extract ancestors and build up the relationship matrix.

3.2 Data Editing

3.2.1 Test day records

Each animal, tested by the QDHAS, had an average number of 8 to 10 monthly measurements (or test day records) per lactation. For a separate analysis of Alpine, Saanen, Toggenburg and Nubian data, the test day records were divided into records by breed. Alpine had 24265 records on all lactations, Toggenburg had 23364 records, Saanen had 17797 records and the number for Nubian test day records for all lactations was 15611. The rest, counting 7928 records, belonged to breeds like LaMancha, Boer etc., which were excluded from the analysis. Data from each breed was edited separately. The number of records and reasons for editing are summarised for all four breeds in Table 7.

	Alpine	Toggenburg	Saanen	Nubian
TEST DAY RECORDS	24265	23364	17797	15611
Editing	# of records edited			
Reg. # \neq 0, Sire = 0. Dam = 0	562	138	624	594
Not first lactation records	14049	14465	10514	9351
Differ. from kidding to first-test-day >365	270	641	189	149
Animals with just one record	36	24	13	35
Animals which have not finished lactation	1626	649	1059	835
Other reasons	628	589	600	184
Clean test day records	7094	6858	4798	4463
NUMBER OF ANIMALS	786	711	501	534
Cumulative milk yield less than 50 kg	8	4	6	13
Single granddaughters	69	37	44	74
Single animal flocks	18	29	12	14
Final Number	691	641	439	433

Table 7. Edits per breed

Test day records missing sire or dam registration number were discarded. The evaluation in this study was based on first lactation records only. After eliminating second and higher lactation records, the following edits were based on first lactation test day records: Records which had a difference larger than 365 days between date of kidding and first day of testing were considered to be wrong records and also excluded from the evaluation. Animals with just one record on their first lactation were dropped to avoid any wrong data and resulting bias. All animals, which had not finished their lactation, because of selling, or nursing reasons etc. were excluded.

Other reasons to edit records included animals, which were declared to be in first lactation but had a lactation continuing far beyond 305 days. By manually checking the records, they showed extreme variation considering the lactation curve, such as increasing yield over a few months and after the peak the yield decreased until again an increase occurred. Furthermore, animals which were younger than 10 months or older than 30 months for age at first kidding were excluded, because does younger than 10 months would have been bred at 5 months of age and that is quite unlikely, and after 30 months a doe is probably in her in second lactation. Animals with cumulative milk yield less than 50 kg, single granddaughter animals and animals which were the only animal in a flock were excluded from this study. For animals without siblings or cousins, or single animal flocks, the flock fixed effects can not be separated from the sire effect, which means these effects are confounded and those animals have to be deleted.

3.2.1.1 Time period of test day records

The numbers of records for breed and year of animals' date of birth, which passed editing, are given in Table 8. Also given are the numbers of animals for breed and year the production was recorded. Animals included in the analysis were born between 1984 and 1995, and the time period in which the production traits for the animals are recorded is from 1986 to 1996.

BREED	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	1996
Year of birth													
Alpine	13	52	86	69	55	61	86	109	57	42	52	9	•
Saanen	5	34	56	50	56	43	63	41	43	26	20	2	-
Togganburg	2	33	62	37	59	87	111	94	79	71	6	0	-
Nubian	1	31	56	48	38	31	69	63	49	21	25	1	-
Year of recording	3												
Alpine	-	-	35	94	73	54	60	69	108	73	49	61	15
Saanen	-	-	18	58	57	41	64	49	49	43	25	32	3
Togganburg	-	-	14	6 2	44	40	75	113	72	111	68	40	2
Nubian	-	-	10	54	53	39	37	43	65	71	27	25	9

Table 8. Number of records per year and breeds for animals' date of birth and year of recording

3.2.1.2 Season of kidding

The months of kidding in the data set used are similar to the seasons described in the literature. Most of the animals included in the analysis were born between January and June (Figure 6). After reaching sexual maturity, these animals were bred, around September, October or November and gave birth five month later in January to April.



Figure 6. Date of birth distribution of does for first lactation across 12 month

It is interesting that the shape of Figure 6, which is the graph of the number of does born in the different month, is similar to that of Figure 7 which represents the number of does kidding in the different months of the year.



Figure 7. Kidding distribution of does for first lactation across 12 month

Raw means by breed and months of kidding for cumulative milk, fat and protein yield of first parity does are given in Table 9 to 12.

month	Ν	cum. Milk (kg)	cum. Fat (kg)	cum. Protein (kg)
January	25	541.80	18.32	16.14
February	89	746.60	26.75	22.97
March	190	750.39	25.44	22.58
April	141	796 .33	27.71	24.05
May	105	656.54	22.44	19.73
June	46	862.20	30.50	25.94
July	21	743.81	26.47	22.61
August	14	642.57	25.91	21.15
September	25	753.68	27.26	24.47
October	14	602.00	23.81	19.78
November	15	696.27	23.00	20.67
December	6	928.17	28.81	28.4 9
Total	691	740.01	25.74	22.48

Table 9. Number of observations and raw means by month of kidding for lactation traits (ALPINE)

For all breeds, most of the animals were born in March to May. Almost 63 % of the Alpine does, 55 % of the Toggenburg does, 56 % of the Saanen does and 63 % of the Nubian does gave birth in these three months.

month	Ν	cum. Milk (kg)	cum. Fat (kg)	cum. Protein (kg)
January	27	827.48	27.33	24.98
February	77	836.35	26.17	23.55
March	172	880.64	27.81	24.70
April	103	676.28	21.38	19.07
May	79	770.56	24.70	22.39
June	47	656.15	21.12	18.68
July	29	801.69	26.57	23.55
August	12	913. 5 0	28.39	25.40
September	18	519.17	16.69	14.52
October	44	797.66	27.93	22.91
November	18	824.94	26.40	23.26
December	15	1155.87	34.55	32.79
Total	641	796.29	25.46	22.64

Table 10. Number of observations and raw means by month of kidding for lactation traits (TOGGENBURG)

month	N	cum. Milk (kg)	cum. Fat (kg)	cum. Protein (kg)
January	31	860.06	30.22	26.07
February	58	953.31	30.58	27.64
March	99	829.94	26.54	24.63
April	82	758.12	23.82	22.73
May	64	755.16	23.95	22.71
June	27	817.96	26.36	24.53
July	9	808.78	24.98	24.83
August	6	767.33	27.51	22.43
September	13	769.31	27.32	23.55
October	22	961.73	30.28	28.59
November	16	856.69	27.52	25.40
December	12	752.25	29.32	24.43
Total	439	825.68	26.74	24.65

Table 11. Number of observations and raw means by month of kidding for lactation traits

(SAANEN)

Observing raw means, Toggenburg does which gave birth in September had a significant drop in milk, fat and protein yield. It was not obvious that Nubians were more or less seasonal than the other breeds. The highest milk, fat and protein yields were for animals, which gave birth in September, and the lowest for animals, which were kidding in July.

month	Ν	cum. Milk (kg)	cum. Fat (kg)	cum. Protein (kg)
January	31	548.61	26.39	20.29
February	56	484.43	22.96	17.13
March	93	529.74	25.07	19.58
April	104	487.41	22.49	17.48
May	74	498.66	23.55	18.95
June	26	395.08	20.54	15.22
July	9	334.56	15.12	11.74
August	4	496.50	25.82	18.37
September	13	741.00	37.65	29.33
October	12	502.08	23.08	19.07
November	6	430.17	20.14	16.04
December	5	452.60	20.63	17.73
Total	433	500.61	23.74	18.48

Table 12. Number of observations and raw means by month of kidding for lactation traits (NUBIAN)



Figure 8. Least square means for month of kidding for cumulative milk, fat and protein yield

Using the combined data set of all four breed, the SAS procedure 'proc glm' was used to fit a fixed effects model with age, breed, month of kidding and flock-year, to estimate least-squares-means for cumulative milk, fat and protein yields (Figure 8).

The graph shows just a small difference between months for cumulative milk yield and almost no change in cumulative fat and protein yield. This led to the decision that season has almost no effect on cumulative milk, fat and protein yield and just flock-year instead of flock-year-season was fitted later in the final model and analysis for all four breeds.

3.2.1.3 Age at first kidding

Unadjusted means by kidding age for lactation traits for first parity does for Alpine, Toggenburg, Saanen, and Nubian are given in Tables 13 to 16. Some animals are listed to be 9 months old, despite the fact, that it is an editing criterion. These animals were manually checked and reintegrated in the process, because they were found to be below 10 months (9 months and 20 days old). For Alpine, relatively few does were kidding between the ages of 16 to 20 months. This is explainable since most does born between February and May would have to conceive outside of the natural breeding season in order to kid within these ages (Sullivan, 1988).

Age	N	Cumulative	Cumulative	Cumulative
9	7	444.71	13.43	12.34
10	17	533.24	18.50	16.10
11	45	630.38	21.27	18.75
12	116	685.42	24.11	20.86
13	123	768.83	26.35	23.33
14	55	670.47	23.67	20.35
15	49	751.53	26.87	23.54
16	34	843.38	28.62	25.82
17	37	828.00	29.51	26.23
18	25	731.12	27.62	23.10
19	17	731.76	24.55	21.53
20	16	722.06	23.47	21.32
21	19	700.74	23.74	20.45
22	35	850.29	30.47	25.67
23	48	768.31	27.02	22.98
24	25	780.68	26.26	23.76
25	13	870.46	29.08	25.56
26	4	1296.50	45.61	39.48
27	3	997.00	33.29	28.93
28	2	868.50	28.70	25.00
29	1	743.00	29.78	25.72

Table 13. Number of observations and unadjusted means by age at kidding for lactation traits. (Alpine)

Most of the animals for each breed gave birth at 12 or 13 months of age. The general pattern for all breeds was an increase in cumulative milk yield, cumulative fat yield, and cumulative protein yield with age. However, for Alpine a drop in milk yield, fat yield and protein yield was observed at 14 / 15 months of age.

Age (months)	N	cum. milk yield (kg)	cum. fat vield (kg)	cum. protein vield (kg)
9	2	807.00	25.48	21.60
10	7	639.57	19.7 2	19.30
11	30	632 .10	20.66	18.62
12	48	633.85	19.57	17.64
13	65	598.28	18.95	16.91
14	56	675.95	22.00	19.48
15	44	912.70	29.24	25.96
16	45	812.07	25.14	22.57
17	39	857.95	27.63	24.21
18	33	832.33	27.45	23.66
19	36	890.33	30.21	25.76
20	27	788.48	25.83	22.44
21	32	932.47	29.53	26.58
22	36	1024.25	31.15	29.32
23	46	844.33	26.59	23.75
24	51	857.71	27.47	24.58
25	19	1031.95	34.00	29.12
2 6	11	688.09	21.99	19.75
27	8	767.13	24.51	21.57
28	4	773.00	23.88	21.88
29	1	892.00	27.96	22.58
30	1	640.00	20.84	19.28

Table 14. Numbers of observations and unadjusted means by age at kidding for lactation traits (TOGGENBURG).

The trend that in some months the number of kiddings is reduced was not so clear with Toggenburg, but the drop of milk, fat and protein yields could also be observed for this breed. It occurred at the age of 13 and 18 months.

Age (months)	N	cum. milk yield (kg)	cum. fat vield (kg)	cum. protein yield (kg)
9	4	637.75	19.62	18.56
10	16	727.69	22.81	21.12
11	36	668.39	21.29	20.10
12	71	692.17	21.00	20.47
13	51	792.47	25.77	24.01
14	54	755.81	23.79	22.38
15	31	839.32	28.62	25.57
16	18	926.83	32.63	27.54
17	26	1030.92	33.73	30.73
18	15	896.80	29.18	26.95
19	8	681.50	22.29	21.17
20	19	942.95	29.18	27.47
21	18	875.44	28.10	25.53
22	24	1034.42	34.38	30.96
23	24	963.42	33.15	29.00
24	12	1050.33	36.00	31.56
25	9	984.22	31.22	29.09
26	2	553.00	21.01	17.59
30	1	1245.00	35.45	43.18

Table 15. Numbers of observations and unadjusted means by age at kidding for lactation traits (SAANEN)

Saanen also had a reduced number of does kidding between 16 and 21 months. A drop in the milk, fat and protein yield could be observed for 12 and 20 months of age.

Age (months)	N	cum. milk yield (kg)	cum. fat vield (kg)	cum. protein vield (kg)
10	4	595.25	28.66	21.71
11	20	485.55	22.12	17.11
12	58	508.14	24.11	18.60
13	64	454.05	22 .07	17.07
14	52	476.58	23.40	17.80
15	28	471.89	22.99	17.83
16	20	537.20	25.09	19.70
17	17	464.24	23.45	18.27
18	13	503.31	24.37	19.28
19	15	575.13	27.05	20.83
20	20	674.50	30.38	24.26
21	25	435.08	18.61	15.57
22	28	471.86	22.85	17.31
23	27	573.96	27.30	21.48
24	17	455.88	21.24	16.27
25	9	573.89	28.46	20.16
26	8	496.63	24.18	18.96
27	5	570.60	21.75	20.84
29	3	507.67	20.11	18.25

Table 16. Numbers of observations and unadjusted means by age at kidding for lactation traits (NUBIAN).

The milk, fat and protein yields for 19 month old Nubians, were 32 % less than those of animals which were 18 month old when they gave birth and even 42 % less compared to animals which were 20 month of age. The reduced number of kiddings between 16 and 20 months of age was not so clear as with Alpine or Saanen.

To check whether age has a significant effect on milk yield or not, SAS was used to fit a fixed effects model including age and flock-year, to estimate least-square-means of cumulative milk yield for age at first kidding.

SAS results (Figure 9) showed a clear trend upward, despite some fluctuations, leading to the conclusion that age had to be fitted in the final model of analysis.



Figure 9. Least square means for age of first kidding for cumulative milk yield

<u>Animals</u>

Table 17 contains the number and size of flocks per breed. Of 58 Alpine flocks, 18 were smaller than five animals per flock with a minimum size of 2 animals. This equals 31 % of Alpine flocks. Three of the flocks had 51, 68, and 80 animals, which equals 29 % of all Alpine does. Toggenburg animals were distributed in 39 flocks and 45 % of the flocks were smaller than 5 animals. One flock had 217 animals, which is almost 34 % of the whole number of animals (641). 50 % of the total number of Saanen flocks (50) were smaller than 5 animals and the biggest flock contained 53 animals. Nubian flocks were relatively small. Most of them (78 %) had 2 to 10 animals and just a few had higher numbers of animals with a maximum size of 35 animals. The average flock size was 11.91 animals for Alpine, 16.44 animals for Toggenburg, 8.78 animals for Saanen and 8.02 animals for Nubian.

Flocks with	Alpine	Toggenburg	Saanen	Nubian
< 5 animals	18	17	25	22
5-10 animals	19	6	14	20
11-15 animals	3	5	3	5
16-20 animals	2	4	1	3
21-30 animals	2	3	5	2
31-50 animals	3	3	1	2
51-100 animals	3	-	1	•
101-200 animals	-	-	-	-
>200 animals	-	1	-	-
Total number of flocks	58	39	50	54
Average # of animals per flock	11.91	16.44	8.78	8.02

Table 17. Number of flocks and different sizes per breed

Sire

The frequencies of sires and dams per flock are shown in Table 18 and 19 respectively. The average number for Alpine was 3.10, for Toggenburg 3.92, for Saanen 3.15 and for Nubian 2.78 sires per flock. Two Alpine flocks had 18 and 28 different sires respectively. Toggenburg had one flock with 32 sires and for Saanen the maximum number of sires per flock was 15. The maximum number of sires for a Nubian flock was 13 animals.

Number of sires	Alpine	Toggenburg	Saanen	Nubian
1	9	4	9	7
2	11	12	12	14
3	14	5	8	11
4	6	3	8	5
5	7	3	2	8
6	1	3	2	2
7	2	1	2	0
8	3	2	3	0
9	0	2	0	1
10	2	1	1	2
11-20	1	1	2	3
21-30	1	1	0	0
Total # of flocks	58	39	50	54
Average # of sire per flock	3.10	3.92	3.15	2.78

Table 18. Frequency distribution of sires per flock

<u>Dam</u>

Most of the flocks had two to three dams. The average number for Alpine was 8.91, for Toggenburg 11.67, 6.56 for Saanen and 5.39 for Nubian (Table 19).

Number of dams	Alpine	Toggenburg	Saanen	Nubian
2	18	8	13	16
3	7	7	11	7
4	3	5	4	4
5	6	2	5	6
6	7	3	5	8
7	3	1	0	2
8	0	0	1	0
9	l	2	1	3
10	0	0	1	1
11-20	5	6	5	3
21-50	6	3	3	3
>50	1	1	0	0
Total # of flocks	58	39	50	54
Average # of dams per flock	8.91	11.67	6.56	5.93

 Table 19.
 Frequency distribution of dams per flock

3.2.1.5 Offspring per sire

Table 20 shows the frequency distribution of numbers of daughters per sire. 46 Alpine bucks had just one daughter, which equals almost 26 % of all sires. 10 of the alpine sires had from 11 up to 39 daughters. 21 % of Toggenburg sires had one daughter, and 12 of the total number of 153 sires had 11 to 31 daughters. For Saanen, 30 sires, which equals 24 % had one daughter. Six sires had more than 10 daughters. 44 Nubian sires had just one daughter, which equals 29 %, and just three had more than 10 daughters (11, 14, and 16). The average number of daughters per sire was 3.84 for Alpine, 4.19 for Toggenburg, 3.57 for Saanen and 2.89 for Nubian.

Number of daughters	Alpine	Toggenburg	Saanen	Nubian
1	46	32	30	44
2	45	39	34	47
3	30	22	17	18
4	16	19	10	16
5	10	7	13	11
6	7	8	7	5
7	5	7	3	2
8	5	2	2	0
9	2	3	1	3
10	4	2	0	1
11	2	2	0	l
> 11	8	10	6	2
Total # of sire	180	153	123	150
Average # of daughters per sire	3.84	4.19	3.57	2.89

Table 20. Frequency distribution of numbers of daughters per sire.

3.2.1.6 Offspring per dam

The frequency distribution of numbers of daughters per dam is summarised in Table 21. Out of 517 Alpine dams, 389 had just one daughter. The maximum number of daughters was 6. Of 455 Toggenburg dams 323 had one daughter, (71 %) and 1 dam had 6 daughters. Saanen had a total number of 328 dams, 76 % of these dams had one daughter. Like the previous breeds, just one dam had more than 5 daughters (6). The maximum number of daughters per Nubian dam was 5. In total 320 dams were known and 75 % of these dams had one daughter. The average number of daughters per dam is 1.37 for Alpine, 1.41 for Toggenburg, 1.34 for Saanen and 1.35 for Nubian.

Number of daughters	Alpine	Toggenburg	Saanen	Nubian
1	389	323	248	240
2	95	94	59	52
3	23	27	14	24
4	8	7	5	3
5	1	3	1	1
>5	1	1	I	0
Total # of dams	517	455	328	320
Average # of daughters per dam	1.37	1.41	1.34	1.35

Table 21. Frequency distribution of numbers of daughters per dam.

3.2.2 Pedigree

The pedigree file received from the Canadian Centre for Swine Improvement was used to build up an unique pedigree for each breed. No editing concerning incorrect records had to be done. Steps for building up the pedigrees for each breed is summarised in Table 22.

Table 22. Steps for building up the pedigr	tor building up the pedigre	angiy :	IOL	Steps	LL .	I able
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ALPINE - PEDIGREE	Records	Zero Birth date
Animals with records	691	0
Extracted ancestors	2408	1101
Phantom parents	772	772
Total records	3871	1873
TOGGENBURG - PEDIGREE	Records	Zero Birth date
Animals with records	641	0
Extracted ancestors	1530	657
Phantom parents	471	471
Total records	2642	1128
SAANEN - PEDIGREE	Records	Zero Birth date
Animals with records	439	0
Extracted ancestors	1501	702
Phantom parents	532	532
Total records	2472	1234
NUBIAN - PEDIGREE	Records	Zero Birth date
Animals with records	433	0
Extracted ancestors	1805	839
Phantom parents	663	703
Total records	2901	1542
COMBINED ALL BREED - PEDIGREE	Records	Zero Birth date
Animals with records	2204	0
Extracted ancestors	6774	2936
Phantom parents	2034	2034
Total records	11012	4970

Difficulties in identifying animals occurred because in the test day records the animal's identification is based on an eight-digit number, but in the pedigree file, country-breed codes are used in addition to the eight-digit number. Matching with just the number was impossible, because some numbers were used more than once in the pedigree. What made the animals unique in the pedigree file was the combination of country-breed code and that eight-digit number. The solution was matching test day record and pedigree animals with all their criteria (goat, sire and dam identification and date of birth) and assigning all test day record animals the country-breed code found in the pedigree file. Over 95 % of the animals were successfully renumbered without possible mix-up.

For the pedigree file all animals, even those with missing parents or date of birth were kept. Using FORTRAN programs, ancestors of all breeds were extracted from the pedigree file and for missing dams, sires or birth dates phantom parents and calculated birthdates were assigned. Data handling for the pedigree is summarised in Table 22. For the use in DFREML, the pedigree input file was sorted in logical order (oldest animals in the beginning followed by their offspring etc.) and renumbered from 1 to N.

3.2.2.1 Genetic grouping

Under the assumption that knowing relationships (pedigree) account for the effect of selection on the related ancestors, there is no need to assign groups when both parents are known. In case of having many animals with one or both parents unknown genetic groups must be assigned to account for genetic selection and trends. For example, migration of animals into a population results in a need for grouping to account for genetic merit of the migrant, or selection on parents for which records are not available (Westell et al, 1988). Pollak and Quaas (1983) demonstrated that the need for grouping decreases with increasing completeness of the genetic relationship information. To group animals, phantom parents for all animals with missing sire or dam information were assigned to replace each of the unknown animals. Phantom parents are assumed to be average representatives of the genetic group of animals selected to be parents at the same time. The concept of grouping by their year of birth for genetic evaluation was introduced by

C.R. Henderson (Henderson, 1973). An alternative approach, developed by Westell (1984), is to group only those animals missing one or both parents. Animals can be grouped by different criteria. Gianola and Fernando (1986) and Howson and Urbach (1989) formed genetic groups using non-data-based (diffuse) prior information. Genetic grouping based on birth period or selection criteria are proposed by Quaas and Pollak (1981), Robinson (1986), Quaas (1988), Westell et al. (1988) and Van Vleck (1990). Principles of assigning phantom parents are developed by Westell, 1984.

3.2.2.2 Example of assigning phantom parents [from Westell et al. 1988]:

Consider a case of having a pedigree where not all relatives' information is available (Figure 10).



Figure 10. Sample Pedigree [from Westell et al, 1988]

Both parents are unknown for S_1 , D_1 , D_2 ; one parent is unknown for D_3 , D_4 , D_5 , and S_2 and both parents of S_3 and S_4 are known. Figure 11 shows the phantom parents assigned in the example of Figure 10 and one half of the effect of the phantom parent genetic group is attributed to its progeny.

GENERATION



Figure 11. Assignment of phantom parents [from Westell et al, 1998]

Table 23 shows the assigned phantom groups and the resulting coefficient matrix, which will be included in the model to account for missing relatives and resulting bias.

Pedigree information			Coefficient Matrix							
Animal	Sire	Dam	Generation	PS1	PD1	PS2	PD2	PS3	PD3	Sum
DI	0	0	1	0.5	0.5	0	0	0	0	1
S1	0	0	1	0.5	0.5	0	0	0	0	1
D2	0	0	2	0	0	0.5	0.5	0	0	1
S 3	S 1	DI	2	0.5	0.5	0	0	0	0	1
S4	S 3	D2	3	0.25	0.25	0.25	0.25	0	0	1
D5	0	D2	3	0	0	0.25	0.25	0.5	0	1
D4	S 3	0	3	0.25	0.25	0	0	0	0.5	1
S2	S 1	0	2	0.25	0.25	0	0.5	0	0	1
D3	<u>S2</u>	0	3	0.125	0.125	0	0.25	0	0.5	1

Table 23. Pedigree and resulting coefficient matrix for example of Figure 13.

Principles of assigning phantom groups, equivalent mixed model equations and rules for calculating coefficients due to groups can be reviewed in Westell et al, 1988 and Westell, 1984.

3.2.2.3 Phantom groups

Selection differential may be different between phantom male and phantom females, so it is useful to set at least two parallel groups representing phantom sire and phantom dams.

The decision was made to assign four groups (born before 1966, 1966-1970, 1971-1975 and born after 1975) for phantom sire and four groups for phantom dams by estimated date of birth, ending up with eight different genetic groups. The numbers of animals per group are given in Table A1 (Appendix). To assign missing birth dates to animals, the average generation interval for the four selection pathways (phantom sire of buck, phantom sire of doe, phantom dam of buck and phantom dam of doe) were estimated (Table 24) of each breed and sex and subtracted from the birth date of the oldest offspring.

	Path of selection							
Breed	Dams to produce does	Dams to produce bucks	Sire to Produce does	Sire to produce bucks				
Alpine	3.28	4.64	2.65	3.05				
Toggenburg	3.15	4.46	2.81	3.22				
Saanen	3.48	4.93	2.89	3.32				
Nubian	2.90	4.10	2.94	3.37				
All breeds	3.20	4.53	2.82	3.24				

Table 24. Generation interval in years

3.4 Statistical Analysis

SAS for OS/2, release 6.12 was used for to test the data sets for normality. The data sets were found to be slightly skewed.

The lactation length was quite similar for Alpine, Saanen and Toggenburg and was in the range reported in the literature. The longest milking period had Toggenburg with 255 days followed by Alpine and Saanen. With a length of 172 days, the Nubians are significantly shorter in lactation than the other breeds (Table 25).

The number of flock-years was similar for all breeds. Nubians had the highest number with 157 flock years, followed by Alpine 155 and Saanen 146. Toggenburg had the lowest number of flock-years with 133. Maternal genetic effects can be estimated in different ways with different relationships of the animals in the data set. Probably the most important ones are the daughter-dam pairs. These pairs show how many animals have a dam with records in the data set. Alpine had the lowest number and percentage of

daughter-dam pairs and Toggenburg had the highest number. The mean of daughter-dam pairs for all breeds was around 18.4 % of all animals in the data set (Table 25).

Breed	Lactation length	Year-seasons	Daughter- Dam pairs Daughter- Dam pairs			
	(days)	(N)	(N)	(%)		
Alpine	247	155	106	15.3		
Toggenburg	255	133	146	22.8		
Saanen	246	146	78	17.8		
Nubian	172	157	75	17.3		
All breeds	230	485	405	18.4		

Table 25. Lactation length and number of flock-year-seasons per breed.

3.4.1 Mixed Model analysis

Computations of genetic and phenotypic parameters of cumulative first lactation traits were carried out using derivative-free REML for estimating variance and covariance components for univariate animal models by restricted maximum likelihood (Meyer, 1989). An animal mixed model including additive genetic effects and maternal genetic effects as additional random effects was fitted for data sets of the four breeds. The mixed model included effects of flock-year, age at first kidding and regressions on phantom groups as fixed effects, and additive, maternal and residual components as random effects. All known additive genetic relationships amongst animals in the pedigree were used to construct the relationship matrix.

3.4.2 The Models

The following basic models (1) and (2) were used for single breed data sets:

$$\mathbf{y}_{ijklmn} = \boldsymbol{\mu} + \mathbf{h}\mathbf{y}_i + \mathbf{age}_j + \mathbf{r.e.} + \mathbf{e}_{ij(klm)n}$$
(BM1)

$$y_{ijklmnop} = \mu + hy_i + age_j + \sum_{k=1, 4} b_k PS_k + \sum_{l=1, 4} b_l PD_l + r.e. + e_{ijkl(mno)p} (BM2)$$

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For the multiple breed data sets the following basic models (1) and (2) were used:

$$\mathbf{y}_{\text{hijklmn}} = \boldsymbol{\mu} + \mathbf{breed}_{h} + \mathbf{h}\mathbf{y}_{i} + \mathbf{age}_{j} + \mathbf{r.e.} + \mathbf{e}_{\text{hij}(\text{klm})n}$$
(BM1)

$$\mathbf{y}_{\text{hijklmnop}} = \mu + \mathbf{breed}_{\mathbf{h}} + \mathbf{h}\mathbf{y}_{\mathbf{i}} + \mathbf{age}_{\mathbf{j}} + \sum_{\mathbf{k}=1,4} \mathbf{b}_{\mathbf{k}} \mathbf{PS}_{\mathbf{k}} + \sum_{\mathbf{l}=1,4} \mathbf{b}_{\mathbf{l}} \mathbf{PD}_{\mathbf{l}} + \mathbf{r.e.} + \mathbf{e}_{\text{hijkl(mno)p}} (BM2)$$

With three sub-models depending on which random effects were included in the basic model, 6 different models were fitted for each data set.

r.e. = Random effects: Model 1 : a_m Model 2 : $a_m + m_n$ Model 3 : $a_m + m_n$ [c₀]

where:

= the hijklmnopth first lactation cumulative milk yield, fat yield or where, Yhijkimpop protein yield = the fixed effect of the h^{th} breed (4 levels) breed = the fixed effect of the i^{th} flock-year of first kidding hy_i (Alpine: 155 levels, Toggenburg: 133 levels, Saanen 146 levels, Nubian: 157 levels, all breeds 485 levels) = the fixed effect of age at first kidding j, in month age; (Alpine: 21 levels, Toggenburg: 22 levels, Saanen 19 levels, Nubian: 19 levels, all breeds 22 levels) =linear regression coefficients for y on the sum of phantom groups of sires (n = 4)k=1. n linear regression coefficients for y on the sum of phantom groups of **b**_iPD_i dams (n = 4)l≈1.n = the mth random direct genetic effect, assumed to be normally 8. distributed with mean 0 and variance σ^2 .

m,	the n th random maternal genetic effect, assumed to be normally
	distributed with mean 0 and variance σ^2_{m} ,
Cij kimn op	the random residual associated with the doe ijklmnop, assumed to
	be normally distributed with mean 0 and variance σ^2_{c}

 c_o refers to the covariance between direct genetic and maternal genetic effects, which was included when fitting Model 3 only.

3.4.3 DFREML procedure

The DFREML 2.1 package by Karin Meyer (1993) was obtained via ftp from the University of New England. The DFREML package is a Fortran program using subroutines for estimating genetic parameters using derivative-free restricted maximum likelihood. Theoretical steps and main computational steps can be summarised as follows:

The general mixed model is

$$\mathbf{Y} = \mathbf{X}\mathbf{b} + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

where

Y is the vector of N observations
b is the vector of NF fixed effects (including regression coefficients)
u is the vector of NR random effects
e is the vector of residuals
X is the NxNF incidence matrix pertaining to b
Z is the NxNR incidence matrix pertaining to u

It is assumed that:

 V(u) = G, E(u) = 0

 $V(e) = I\sigma^2_e = R,$ E(e) = 0

 Cov(u,e) = 0, E(y) = Xb

 V(y) = ZGZ' + R

Then:

The mixed model equations are:

$$\begin{bmatrix} \mathbf{X}^{*}\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}^{*}\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}^{*}\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}^{*}\mathbf{R}^{-1}\mathbf{Z}+\mathbf{G}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{b} \\ \mathbf{\hat{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{X}^{*}\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}^{*}\mathbf{R}^{-1}\mathbf{y} \end{bmatrix}$$

Under an animal model, u always contains the vector of animal's direct additive genetic effects, a. Maternal additive effects are taken into account by fitting vectors of maternal genetic effects (m), if traits analysed are subject to maternal effects (Meyer, 1996).

$$\mathbf{u} = \begin{pmatrix} \mathbf{a} \\ \mathbf{m} \end{pmatrix}$$

with corresponding partitioning $Z = (Z_A | Z_M)$. Further let A denote the numerator relationship matrix between animals in the univariate case,

$$\mathbf{G} = \mathbf{V} \quad \begin{pmatrix} \mathbf{a} \\ \\ \\ \mathbf{m} \end{pmatrix} = \begin{pmatrix} \sigma^2_{\mathbf{A}} \mathbf{A} & \sigma_{\mathbf{A}\mathbf{M}} \mathbf{A} \\ \\ \sigma_{\mathbf{A}\mathbf{M}} \mathbf{A} & \sigma^2_{\mathbf{M}} \mathbf{A} \end{pmatrix}$$

The likelihood is:

$$L = \frac{1}{(2\pi)^{\frac{r_2}{2}} |V|^{\frac{r_2}{2}}} e^{-\frac{1}{2} (y - Xb)' V''(y - Xb)}$$
$$\ln L = -\frac{n}{2} \ln 2\pi - \frac{1}{2} \ln |V| - \frac{1}{2} (y - Xb)' V''(y - Xb)$$

where $-\frac{n}{2} \ln 2\pi$ is a constant and V is the determinant of the variance, covariance matrix.

Restricted Maximum likelihood theory can be reviewed in Meyer (1983), Meyer (1985), Meyer (1986), Meyer (1989) and Meyer (1993).

For maximising the likelihood function, an initial estimate of the optimum must be used, and corresponding function values obtained. Iteration must be done until the point of maximum is reached, or, in other words, the aim of each iteration is to replace the worst point, e.g. for a maximising problem the point with the lowest function value. The next point, which is defining the next simplex, is chosen in a direction away from the discarded point. The procedure allows the simplex to rescale itself automatically in each iteration, changing shape and size according to the landscape of the surface being searched. This adaptability is achieved by a combination of so-called reflections, expansion and contraction steps. Iterations are repeated until the simplex finds the optimum (Perotto, 1992)

3.4.4 Likelihood ratio test

The likelihood ratio test evaluates the significance of a model containing one or more additional parameters, compared to another identical model, except for the omission of those parameters. The numerical values of the maximum of the likelihood function are required under both, the null (H₀) and the alternative hypothesis (H_A). Minus twice the difference in the two logL asymptotically has a χ^2 distribution with degrees of freedom equal to the number of parameters tested, i.e. it can be compared to tabulated χ^2 values in order to decide whether to accept or reject the H₀. (Meyer, 1992) Manual DFREML

3.4.5 F-test

The F-test was used to evaluate both basic models as to whether phantom groups have an effect and the null hypothesis can be accepted or must be rejected. Animals were seen to be nested within phantom groups and the following equation was used to calculate F_{c} -values.

$$F_{c} = \frac{SSR(R) - SSR(F) / N}{\sigma_{E}^{2}(F) + \sigma_{A}^{2}(F)}$$

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where: SSR (R) = Sums of square for residual from the reduced model SSR (F) = Sums of square for residual from the full model $\sigma^2_E(F)$ = Error Variance of full model $\sigma^2_A(F)$ = Additive Variance of full model

N = Number of phantom groups fitted

4. RESULTS AND DISCUSSION

4.1 Basic statistics

Means, standard deviations, minima, and maxima were calculated. The data given by breed for the lactation traits, of first parity does are given in Table 26.

Table 26.	Number of observations (N), unadjusted means, standard deviation, minimum a	nd
	maximum for the different breeds	

Breed	Trait	N	Mean	S. D.	Minimum	Maximum
Alpine	Cumulative milk yield (kg)	691	740.01	371.06	50.00	2182.00
	Cumulative fat yield (kg)		25.74	13.81	1.52	85.92
	Cumulative protein yield (kg)	_	22.48	11.43	1.40	65.61
Toggenburg	Cumulative milk yield (kg)	641	796.29	427.67	53.00	1984.00
	Cumulative fat yield (kg)		25.46	14.00	2.36	69.08
	Cumulative protein yield (kg)		22.64	12.11	1.93	58.20
Saanen	Cumulative milk yield (kg)	439	825.68	442.95	54.00	2081.00
	Cumulative fat yield (kg)		26.74	15.31	1.94	103.30
	Cumulative protein yield (kg)		24.65	13.26	1.78	68.98
Nubian	Cumulative milk yield (kg)	433	500.61	279.67	53.00	1520.00
	Cumulative fat yield (kg)		23.74	14.32	2.43	79.79
	Cumulative protein yield (kg)		18.48	10.66	2.20	57.89
All breeds	Cumulative milk yield (kg)	2204	726.41	405.06	50.00	2182.00
	Cumulative fat yield (kg)		25.46	14.29	1.52	103.30
	Cumulative protein yield (kg)		22.17	12.03	1.40	68.98

There were large differences between breeds. Nubian does had a much lower milk yield than that of the other breeds, but they had higher fat and protein percentages (4.7 and 3.7 % respectively). Saanen does had the highest milk, fat and protein yields, Alpine and Toggenburg had slightly smaller yields than Saanen. Milk components of these three breeds were also similar with 3.5 % fat and 3.0 % protein for Alpine, 3.2 % fat and 3.0 % protein for Saanen and 3.2 % fat and 2.8 % protein for Toggenburg. The milk yield differences between Nubian and the other three breeds are explainable through the fact that Nubian does had a 31 % shorter lactation length (average of 172 days) than Alpine (247), Toggenburg (255) and Saanen (246 days). The short lactation length of Nubian does agrees with findings in the literature, but in this study the lactation length was much shorter then reported (237 days), (Sambraus, 1994).

The phenotypic means for cumulative milk, fat and protein yield calculated in this study are close to those reported in the literature (Sambraus, 1994; Gall, 1996; Haenlein, 1996). Alpine had slightly smaller milk yields (740 kg) than reported (~950 kg). Saanen often described as the most developed dairy breed had the highest milk yield at 826 kg. Nubians are described in the literature as not heavy milkers, but with a very high fat yield. This could also be observed in this study. Raw means have to be compared carefully, because not each flock contained all breeds and these breed differences might be due to different flock management/ environment.

The highest variation or standard deviation (S.D.) for all traits had the Saanen breed, followed by Toggenburg and Alpine. Nubians had the smallest standard deviation of all breeds for milk and protein yield, but had a very high standard deviation for cumulative fat yield compared to the other breeds.

4.2 Breeding season

In Canada goats are seasonal breeders (Figure 6 and 7). Most of the kiddings occurred between January and July. That means they were mated between September and December to kid 5 month later, which corresponds with the reported breeding season for goats in the northern hemisphere (Shelton, 1978; Ali, 1983; Mohammad, 1984 and Amoha, 1996). Singh, Acharay and Biswas (1970) found season to be a very important source of variation affecting milk production, but this was not further investigated in this study.

4.3 Distribution of data

The distribution of all traits in each of the data sets was found to be slightly skewed. In statistical analyses, an important assumption is that the data has to be normally distributed otherwise results are biased, but previous studies using REML methodology revealed that heritabilities and genetic and phenotypic correlation of log transformed traits were unchanged from the untransformed traits (Jairath, 1992). Various authors (Harville, 1977;

Banks et al., 1985) indicated that REML estimates may be an appropriate choice even if normality does not hold. Based on these findings, the data sets were not transformed.

4.4 Fixed effects included in the Model

The most limiting factor in this study was the small number of animals with data on cumulative milk, fat and protein yields per breed. Working with such small data sets (small number per flock-season and age at first kidding groups) gives almost no alternative than to accept that the results will be biased. A common fitted effect for dairy cows is the combined effect of flock-year-season (f-y-s), but the numbers of records for each level of f-y-s was so small that the decision had to be made either to accept the small numbers or to fit just flock-year instead of flock-year-season. The least-squares means showed that there was a small difference for milk yield between months and almost no difference for fat and protein yield. Since the final model used was supposed to be fitted for all traits, it seamed appropriate to include just flock-year as a fixed effect. The differences for age at first kidding were clear for all breeds and were fitted as a second fixed effect.

For the combined data set (all four breeds), just breed as an additional fixed effect was fitted. If there would have been more than one observation per animal, breed nested within flock, must have been fitted.

The relatively small flock size for dairy goats, the few sires used in the average flock and the small number of different flocks a sire was used in might result in the possibility that differences between sires are confounded with flock effect.

4.5 Generalised Least Squares (GLS) - solutions for age at first kidding

DFREML obtained solutions for all effects fitted for arbitrary values of the (co) variance by direct inversion of the coefficient matrix, after absorbing random effects.

When estimates of environmental effects (in this case fixed effects), such as flock, breed of animal and age of doe at first kidding are plotted over time, they reveal environmental trend-changes in the mean performance of a population.

Solutions for age at first kidding for cumulative milk yield per breed are shown in Figure 12a to 12d and for cumulative fat and protein yield per breed in Figure 13a to 13d.

All solutions are presented as deviations from the first level, which is 9 and 10 months respectively. The first age effect level is set to zero, for a better comparison. Age differences in production are given in kg for milk, fat and protein yield.

Generalised least squares solutions plotted for age at first kidding show for all traits and breeds a slight increase in production with increasing age. Sullivan (1988) had similar results, that increasing ages at first kidding caused an initial increase in yield and after a peak (26 months of age) the milk, fat and protein yield decreased with age.



Figure 12a-12d.GLS- solutions for age at first kidding effect on milk yield per breed



Figure 13a-13d. GLS- Solutions for age at first kidding effect on cumulative fat and protein yield per breed

Fixed effects solutions for age at first kidding must be interpreted carefully in this study, due to the small number of animals considered to be a population. Alpine, Toggenburg and Nubian solutions show a high initial increase in cumulative milk, fat and protein yield and all breeds show a final high peak in production, in the last month of age at first kidding

Solutions reveal environmental trend changes in the mean performance of a population, meaning you get the most representative results for fixed effects, when the number of animals is equally distributed through age groups. In this study, the early age group

(9months) and the final age group (26-30 months) solutions are estimated with maximum 10 animals, which might result in bias.

The final high peak around 26-28 months, might also be less comparable, caused by animals, which are actually in second lactation, but were listed to be in first. Following the literature (Sullivan, 1988) an age span between 9 and 30 months of age was chosen, but maybe a closer range (12 to 24 months of age at first kidding) would have been more appropriate.



Figure 14a,14b. GLS solutions for age at first kidding effect on cumulative milk, fat and protein yield for the combined, all breed data set.

The combined data set (all breeds) demonstrates more clearly the positive effect of age at first kidding. The plotted solutions show, that with increasing age the milk yield is up to \sim 300 kg higher when first kidding occurs at 24 months instead of at 9 months of age (Figure 14a). The trends are similar for fat and protein yield. The age effect increases fat and protein yield by approximately 10 and 9 kg respectively (Figure 14b).

All breeds combined in one data set increased not just the total animal number, but also the number of animals per age group, which might explain, why fluctuations in numbers plotted are less severe than in the single breed analyses.

4.6 Generalised Least Squares (GLS) - solutions for breed



Figure 15. GLS solutions for breed effects on cumulative milk yield, Alpine set to 0



Figure 16. GLS solutions for breed effects on cumulative fat and protein yield, Alpine set to 0

Breed effects on milk, fat and protein yield for Toggenburg, Saanen and Nubian compared to Alpine which was set to be the operational zero are presented in Figure 15 and 16. Alpine and Toggenburg had similar milk yield, but about 100 kg less production than Saanen. Nubian had lowest yield with 250 kg to 300 kg less than the other breeds. The breed differences for fat yield were much smaller than observed for milk yield, but differences in breed effects on protein yield showed similar trends than those for milk yield.

Values for fixed effects solutions and corresponding standard errors are given in Table A3 to A14 (Appendix).

4.7. Results for variance components and resulting parameter estimates

<u>4.7.1 Alpine</u>

Table 27. Variance components and resulting parameter estimates for Alpine

	Basic Model 1			Basic Model 2			
Milk yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3	
σ_{A}^{2}	26329.41	19135.87	34239.37	26934.65	15317.57	27265.23	
σ^2_M	-	6831.1	15580.47	-	10002.56	17377.03	
$\sigma_{\rm AM}$	-	-	-12445.53	-	-	-10073.01	
σ_{E}^{2}	60521.12	61128.84	51054.78	60124.84	60689.72	53543.13	
σ_{P}^{2}	86850.53	87095.81	88429.08	87059.49	86009.85	88112.38	
$h^2 = \sigma_A^2 / \sigma_P^2$	0.3032	0.2197	0.3872	0.3094	0.1781	0.3094	
s.e.	0.1253	0.1533	0.2486	0.1307	0.1616	0.2589	
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.0784	0.1762	-	0.1163	0.1972	
s.e.	-	0.0862	0.1379	-	0.0903	0.1223	
$c^2 = \sigma_{AM} / \sigma_P^2$	-	•	-0.1407	-	-	-0.1143	
s.e.	-	-	0.1577	-	-	0.1372	
Log L	-4375.160328	-5466.766474	-5466.363733 Model 3	-4167.172493	-5410.504087	-5410.229099	
<u> </u>		15.79	1 VIUUEI J		10.86	20.48	
۰ م م	21.51	[]./ /	24.0	£1.24	7.40	30.48	
O M	-	5.58	11	-	7.40	13.87	
OAM -2-	-	-	-7.47	-	-	-9.04	
0 E	93.41	92.8	87.7	90.17	89.81	84.12	
	114.92	114.16	116.03	117.41	117.08	119.43	
$n = \sigma_A / \sigma_P$	0.1872	0.1383	0.2137	0.2320	0.1697	0.2552	
s.e.	0.1282	0.1482	0.2200	0.1349	0.1587	0.2385	
$m = \sigma_M / \sigma_P$	-	0.0488	0.0948	-	0.0632	0.1162	
s.c .	-	0.0853	0.1292	-	0.0882	0.1372	
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	-0.0644	-	-	-0.0757	
s.e.	•	-	0.1366	-	•	0.1488	
Log L	-2672.44966	-3764.330634	-3764.221708	-2493.673349	-3737.612126	-3737.476646	
Protein yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3	
σ_{A}	16.98	10.1	23.01	19.34	8.89	19.84	
$\sigma_{\rm M}$	-	6.9	15.2	-	10.10	17.43	
σ_{AM}	-	-	-11.04	-	-	-9.86	
σε	63.15	62.64	54.65	61.64	61.92	54.80	
σ _P	80.13	79.64	81.81	80.98	8 0.9 2	82.21	
$h^2 = \sigma_A^2 / \sigma_P^2$	0.2119	0.1268	0.2812	0.2389	0.1099	0.2413	
s.c.	0.1258	0.1484	0.2267	0.1299	0.1539	0.2272	
$m^{-}=\sigma_{M}^{-}/\sigma_{P}^{-}$	-	0.0887	0.1858	-	0.1249	0.2120	
s.e.	-	0.0832	0.1362	-	0.0893	0.1409	
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	-0.1350	-	-	-0.1199	
S.C.	-	-	0.1453	-	-	0.1482	
Log L	-2566.942844	-3669.432024	-3669.005314	-2398.888016	-3642.029984	-3641.700594	

Estimates of (co) variance components, resulting parameters such as heritability (\mathbf{h}^2), maternal heritability (\mathbf{m}^2), genetic covariance, corresponding standard errors and maximum (log) likelihood values for each trait and basic model for Alpine are given in Table 27. Fitting basic model 2 gave, for each trait, slightly higher estimates for σ^2_A , \mathbf{h}^2 and logL, and smaller estimates for σ^2_E . Including maternal effects decreased values of logL markedly over those for model 1 and reduced estimates for σ^2_A in both cases.

Cumulative milk yield

With estimates of the maternal heritability of 7.8 and 11.6 % (BM1 and BM2 respectively), the direct heritability (h^2) was reduced from 30.3 to 22.0 % and from 30.9 to 17.8 % for cumulative milk yield. The change in likelihood values was small compared to the models ignoring σ_{AM} . Estimates of the covariance between direct and maternal effects for milk yield were -14.1 % (BM1) and -1.4 % (BM2) of the phenotypic variance. The negative covariance caused higher estimates for σ^2_A and σ^2_M compared to the models ignoring genetic covariance.

Cumulative fat yield

Heritability was reduced by 4.9 and 6.3 % from 18.7 to 13.8 % and from 23.2 to 16.9 % respectively, fitting models allowing maternal genetic effects. Estimates of \mathbf{m}^2 were 4.9 and 6.3 %. The additional random animal effect decreased the logL value significantly, while reducing σ_A^2 correspondingly. The genetic covariance between direct and maternal effects was moderate and negative and accounted for 6.4 and 7.6 % of the total phenotypic variance. Including σ_{AM} in the model reduced logL values slightly compared to a model ignoring σ_{AM} , but increased σ_A^2 and σ_E^2 , respectively. The phenotypic variance slightly increased through the negative covariance.

Cumulative protein yield

Cumulative protein yield showed similar results to cumulative fat yield. The direct heritability was reduced by 8.5 % from 21.2 to 12.7 % for BM1 and by 12.9 % from 23.9 to 11.0 % for BM2, in which maternal effect was fitted. Maternal effects were estimated

to account for 8.9 and 12.5 % of the total variance. Allowing for direct-maternal covariance yielded a negative estimate amounting to 13.5 and 12.0 %. LogL values went down by 51 % when fitting σ^2_{M} in the model, but changed very little compared to model 2 by fitting σ_{AM} in addition.

Genetic correlations (\mathbf{r}_{AM}) between direct and maternal genetic effects are presented in Table 28. Estimates of \mathbf{r}_{AM} were high and negative for all traits. BM2 gave generally smaller estimates than BM1.

These results indicate that the higher the estimates of the direct genetic variance, the smaller the estimates of the maternal genetic variance and vice versa.

Trait	Basic Model 1	Basic Model 2	
Cumulative milk yield	-0.5388	-0.4628	
Cumulative fat yield	-0.4522	-0.4397	
Cumulative protein yield	-0.5906	-0.5302	

Table 28. Genetic correlation between direct and maternal genetic effects - Alpine

The Likelihood ratio was used to test the fit for each model for the data set and to test the significance of parameters included in the model. Minus twice the difference in the logL values [-2(log₁-log₂)], has a χ^2 distribution with degrees of freedom equal to the number of parameters tested. As the smaller log likelihood values for all traits clearly demonstrate, the data was best described by fitting just direct effects (sub-model 1). Estimates for maternal effects and covariance between direct and maternal effect were not statistically significant for all traits (Appendix Table A16). The likelihood ratios, were 0.8994, 0.3496 and 1.1332 for BM1 and they were all smaller than the tabulated χ^2 values for a probability level of 10 %. Similar results were found for BM2.

4.7.2 Toggenburg

	Basic Model 1			Basic Model 2		
Milk yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3
σÂ	20323.99	18401.25	12145.09	22268.37	22126.35	15340.97
$\sigma^{2}{}_{M}$	-	1144.79	1293.72	-	9.1 3	796.39
$\sigma_{\rm AM}$	-		3961.06		-	3439.35
σ_{E}^{2}	98765.89	99017.99	101095.05	96604.73	96704.34	98724.80
σ^2_P	119089.89	118564.03	118494.93	118873.10	118839.82	118355.51
$h^2 = \sigma^2_A / \sigma^2_P$	0.1707	0.1552	0.1025	0.1873	0.1862	0.1296
s.e.	0.1530	0.1884	0.1988	0.1572	0.2068	0.2247
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.0097	0.0101	-	0.0001	0.0067
s.e.	-	0.0865	0.1002	-	0.0936	0.1099
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	0.0334	-	-	0.0295
s.c.	-	-	0.1009	-	-	0.1118
Log L	-3939.714968	-4704.760849	-4704.555165	-3797.473109	-4652.934982	-4652.811248
Fat yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3
σ_{A}^{2}	25.53	7.59	5.05	24.49	10.79	7.45
σ^2_{M}	-	13.59	8.02	-	10.62	5.71
σ_{AM}	-	•	6.36	-	-	6.50
σ_{E}^{2}	101.48	104.96	106.11	101.67	103.74	105.47
$\sigma^{2}{}_{P}$	127.01	126.15	125.54	126.16	125.16	125.13
$h^2 = \sigma_A^2 / \sigma_P^2$	0.2010	0.0602	0.0402	0.1941	0.0862	0.0595
s.e.	0.1386	0.1440	0.1231	0.1444	0.1622	0.1512
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.1078	0.0639	-	0.0849	0.0456
s.c.	-	0.0838	0.1033	-	0.0885	0.1031
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	0.0507	-	-	0.0520
s.e.	-	-	0.0777	-	-	0.0819
	-2271.514742	-3035.649869	-3035.459859	-2153.460898	-3008.412123	-3008.189177
Protein yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3
σ_A^2	15.78	10.66	7.98	17.96	17. 92	12.42
σ^2_{M}	-	3.14	1.52	-	0.061	0.60
σ_{AM}	-	-	3.48	-	-	2.73
σ_{E}^{2}	79.16	80.07	81.37	77.20	76.97	79.06
σ_{P}^{2}	94.94	93.88	94.35	95.16	94.95	94.81
$h^2 = \sigma_A^2 / \sigma_P^2$	0.1 663	0.1135	0.0846	0.1887	0.1887	0.1310
s.e.	0.1478	0.1810	0.1852	0.1516	0.1968	0.2071
$m^2 = \sigma^2_M / \sigma^2_P$	-	0.0335	0.0161	-	0.0006	0.0063
s.c.	-	0.0746	0.1008	-	0.0917	0.0976
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	0.0369	-	-	0.0288
s.c.	-	-	0.0879	-	-	0.0992
Log L	-2202.757769	-2967.734426	-2967.510865	-2086.115286	-2941.577182	-2941.471424

Estimates of (co) variance components, genetic parameters and corresponding maximum (log) likelihood values for each trait and each analysis are summarised for Toggenburg in Table 29. Results for direct variance, phenotypic variance and direct heritabilities were slightly higher for BM2 fitting phantom groups, except for fat yield. Maternal genetic variance was much lower with BM2 than with BM1.

Cumulative milk yield

The maternal effect included in BM1 decreased the direct variance by 9.5 % from 20324 to 18401. The direct heritability (h^2) went down by 1.5 % from 17 to 15.5 % and the estimate of the maternal heritability (m^2) was 0.96 %, which leads to the conclusion that almost no maternal genetic effects are present for milk yield in Toggenburg goats.

Fitting σ_{AM} increased the maternal genetic variance by just 1.1 % but decreased the direct genetic variance by 34 %. The covariance between direct and maternal effects fitted was positive and accounted for 3.3 % of the total variance.

Same trends were observed for milk yield using BM2, but values were higher for direct variance and direct heritability and smaller for σ_{M}^{2} and σ_{AM} . Fitting just additive direct variance in model 1 gave estimates for \mathbf{h}^{2} of 18.7 % of the total variance. Estimates for maternal heritabilities were 0 % and \mathbf{h}^{2} was hardly reduced fitting maternal effects in addition. Allowing for a covariance between direct and maternal effects reduced \mathbf{h}^{2} to 13 % and increased \mathbf{m}^{2} to 0.06 %. Results for σ_{AM} were smaller when phantom groups were fitted (2.9 % of σ_{PL}^{2} Log likelihood values decreased with the inclusion of maternal effects, but increased slightly when σ_{AM} was fitted in both basic models.

Cumulative fat yield

Direct heritability for fat yield was 20.1 % (BM1) and 19.4 % (BM2). Maternal heritability ranged from 8.4 % (BM1) to 8.9 % (BM2). A high estimate of the maternal genetic variance resulted in a relatively small estimate for direct genetic variance for BM1. Different starting values for the parameters always lead to the same maximum. Basic model 2 gave more acceptable results. The estimates for m² were 8.5 % and 8.6 % of the direct heritability. Adding the direct-maternal covariance as an additional random effect
(Model 3) increased σ_E^2 and decreased σ_A^2 , σ_M^2 and σ_P^2 . The covariance between direct and maternal effect counted for ~ 5 % of the phenotypic variation in both basic models.

Cumulative protein yield

Heritability for protein yield was 16.6 % in BM1 and 18.9 % in BM2. Fitting model 2 increased logL values for both basic models and gave estimates of 3.3 % (BM1) and 0 % (BM2) for maternal genetic effects. The covariance between direct and maternal genetic effects was positive but fairly small (3.7 and 2.9 % of σ^2_{P} and changed the likelihood values slightly compared to models ignoring σ_{AM} .

Genetic correlations between direct and maternal genetic effects are given in Table 30.

	Basic Model 1	Basic Model 2
Cumulative milk yield	0.9860	0.9994
Cumulative fat yield	1.00	0.9974
Cumulative protein yield	0.9960	0.9981

Table 30. Genetic correlation between direct and maternal effects - Toggenburg

Genetic correlations between maternal and direct effects are positive and very high for Toggenburg. The results indicate that increase in direct variance results in a corresponding increase in maternal variance.

As with Alpine goats, the log likelihood value of models 1 demonstrate clearly that maternal genetic effects and the covariance between direct and maternal effects do not improve the goodness of fit, and that the models only including direct effects provide the best fit for the data. Maternal effects for all traits were not statistically significant (Appendix - Table A17).

4.7.3 Saanen

	I	Basic Model 1					
Milk yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3	
σ_A^2	f	0.93	4756.31	f	1.80	21293.36	
σ^2_{M}	-	2307.97	10265.22	-	5861.61	33594.25	
$\sigma_{\rm AM}$	-	-	-6983.55	-	-	-26742.64	
σ_{E}^{2}	f	129676.75	124597.79	f	129571.76	110 796.4 6	
σ_{P}^{2}	ſ	131985.65	132635.76	ſ	135435.16	138941.44	
$h^2 = \sigma_A^2 / \sigma_P^2$	f	0.0000	0.0359	ſ	0.0000	0.1533	
S.C.	f	0.1542	0.2960	f	0.2588	0.2294	
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.0175	0.0774	-	0.0433	0.2418	
s.c.	-	0.1113	0.1386	-	0.1498	0.1496	
$c^2 = \sigma_{AM} / \sigma_P^2$		-	-0.0527	-	-	-0.1925	
S.C.	-	-	0.0375	-	-	0.0003	
Log L	ſ	-3194.788771	-3194.501076	<u>f</u>	-3139.706979	-3139.909288	
Fat yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3	
σ_{A}^{2}	14.89	0.0065	10.61	18.28	0.040	45.84	
σ^2_{M}	-	13.96	37.92	-	16.21	56.78	
σ_{AM}	-	-	-20.05	-	-	-43.99	
σ_{E}^{2}	143.6	143.91	131.21	143.75	145.05	109.65	
σP	158.49	157.87	159.69	162.03	161.30	16 8.28	
$h^2 = \sigma_A^2 / \sigma_P^2$	0.0939	0.0000	0.0664	0.1128	0.0002	0.2724	
s.c.	0.1625	0.1631	0.4608	0.1795	0.2495	0.3851	
$m^2 = \sigma^2_M / \sigma^2_P$	-	0.0884	0.2375	-	0.1005	0.3374	
s.e.	-	0.1115	0.2862	-	0.1453	0.2893	
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	-0.1256	-	-	-0.2614	
s.e.	-	-	0.2771	-	-	0.2701	
Log L	-1589.723288	-2267.572374	-2267.388625	-1561.013863	-2238.874601	-2238.400188	
Protein yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3	
σ_{A}^{2}	4.85	0.0021	11.89	12.01	0.0063	18.58	
σ^2_{M}	-	8.02	16.26	-	10.08	35.57	
σ_{AM}	-	•	-9.37	-	-	-24.93	
σ_{E}^{2}	115.28	112.47	103.54	112.43	113.57	97.61	
σ_{P}^{2}	120.13	120.5	122.33	124.43	123.65	126.84	
$h^2 = \sigma_A^2 / \sigma_P^2$	0.0404	0.0000	0.0972	0.0965	0.0001	0.1465	
S.C.	0.1842	0.2643	0.3308	0.1986	0.2737	0.2758	
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.0666	0.1329	-	0.0815	0.2805	
s.e.	-	0.1457	0.2187	-	0.1516	0.2511	
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	-0.0766	-	-	-0.1966	
s.e.	-	-	0.1962	-	-	0.1771	
Log L	-1553.210892	-2231.103234	-2230.834603	-1526.308188	-2204.259897	-2203.475363	

Table 31. Variance components and resulting parameter estimates for Saanen

Estimates of (co) variance components and resulting heritabilities (maternal and direct), together with the maximum (log) likelihood values for each trait and model are given for Saanen in Table 31. Basic model 2 (fitting coefficients for phantom groups) gave generally higher results for σ_A^2 and σ_P^2 .

Cumulative milk yield

No results for model 1 of BM1 and BM2 could be obtained, even after trying different starting values, very small step sizes and small convergence criteria. The estimates became very small and eventually negative, which caused the DFREML program to fail.

Basic model 1 gave very small estimates for h^2 and m^2 for all sub-models. Including the covariance as an effect increased the estimate of h^2 from 0 to 3.6 % and m^2 from 1.7 to 7.7 %. The covariance between direct and maternal effect (σ_{AM}) was negative and accounted for 5.3 % of the total variance. Estimates using sub-models of BM 2 were slightly higher than with BM1. Better results for direct and maternal heritabilities were obtained by fitting the covariance between direct and maternal effects in BM2. The direct and maternal genetic effect contributed 15.3 and 24.2 % to the phenotypic variance. The covariance between both effects was negative and had a value of 5 and 19.2 % of the total variance respectively. Log likelihood values changed with fitting different random effects or phantom group coefficients. The highest value was reached with model 2 of BM2.

Cumulative fat yield

Heritabilities, for models fitting just additive direct effects were 9.4 % (BM1) and 11.3 % (BM2). Fitting additive maternal effects in addition decreased h^2 almost to zero for BM1 and from 11.2 to 0.02 % for BM2. The reductions were attributed to maternal effects of 8.8 and 10.1 %. Likelihood values decreased significantly by including maternal effects. In addition the covariance increased likelihood values only slightly, but gave higher estimates of h^2 and m^2 . Neither value of h^2 (6.6 and 27.2 %) seems very. Maternal heritability estimated with sub model 3 accounted for 23.7 and 33.7 % of the total variance. The covariance between direct and maternal effects was negative and relatively high (12.6 and

26.1 % of σ^2_{P} . Changes of logL were high when fitting maternal effects in addition, but changed then very little with the covariance included.

Cumulative protein yield

Estimates of genetic parameters for cumulative protein yield showed a trend similar to fat yield. For both basic models likelihood values decreased with included maternal genetic effects and increased the covariance between direct and maternal effects. The two estimates of \mathbf{h}^2 for protein yield were 4 and 9.6 %. Fitting maternal effects gave higher results for \mathbf{m}^2 (6.7 %) than for \mathbf{h}^2 (0%) (BM1), but on the other hand, the direct heritability (9.6 %) estimated with model 2 (BM2) declined when maternal effects were included. Allowing for the covariance between direct and maternal effects gave higher results of \mathbf{h}^2 (9.7 and 14.7 %) and increased \mathbf{m}^2 estimates to 13.3 and 28 %. Covariance estimates were negative and accounted for 7.7 and 19.7 % of the total variation. The highest log likelihood value was obtained with model 1 of BM2.

Genetic correlations between direct and maternal genetic effects are given in Table 32.

	Basic Model 1	Basic Model 2
Cumulative milk yield	-0.9994	-0.9999
Cumulative fat yield	-0.9999	-0.8622
Cumulative protein yield	-0.6736	-0.9696

Table 32. Genetic correlation between direct and maternal effects

Genetic correlations between direct and maternal effects for Saanen were found to be high and negative.

Minus twice the difference between the maximum log likelihood values $[-2(log_1-log_2)]$ of the two models gives answers in how good the model fits the data set. It is demonstrated clearly that maternal genetic effects and the covariance between direct and maternal effects included in the model decrease the likelihood values, hence the 'best' models are just fitting direct genetic effects. Estimates for maternal genetic effects were not statistically

significant. All calculated χ^2 - values were smaller than the tabulated ones for a 10 % probability level (Appendix Table A18).

All results obtained for Saanen appear somewhat questionable. Maternal heritabilities were all small and close to zero, but so were the estimates for direct heritability. Reason for doubting these results are also the very high standard errors and the structure of the data set used. Saanen had the smallest pedigree, and the smallest number of sires and grand sires with progeny records. In addition, the number of dams in total was relatively small and the number of single connection parents relatively large compared to the other breeds (Table A2 - Appendix). Single connection parents are animals, which usually provide no information. The missing or small number of animal relationships might be the reason why estimates are different than expected or could not be obtained at all.

<u>4.7.4 Nubian</u>

<i></i>		Basic Model 1	sic Model 1 Basic Model			. 2		
Milk yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3		
σ_{A}^{2}	14728.37	2.64	369.29	17630.38	0.51	214.22		
σ^2_{M}	-	16991.28	14375.94	-	17889.33	16150.85		
σ_{AM}	-	-	2300.33	-	-	1857.01		
σ_{E}^{2}	42825.46	41221.79	41161.49	41786.29	41677.47	41520.21		
σ_{P}^{2}	57553.84	58215.72	58207.07	59416.67	59567.31	59742.28		
$h^2 = \sigma^2_A / \sigma^2_P$	0.2559	0.0000	0.0063	0.2967	0.0000	0.0036		
S.C.	0.2157	0.2736	0.2763	0.2276	0.2926	0.0.507		
$m^2 = \sigma^2_M / \sigma^2_P$	-	0.2919	0.2469	-	0.3003	0.2703		
s.e.	-	0.1605	0.2952	-	0.1625	0.3875		
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	0.0395	-	-	0.0311		
s.c.	-	-	0.3013	-	-	0.4715		
Log L	-2405.73259	-3193.837862	-3193.81049	-2351.909154	-3139.884979	-3139.868222		
Fat yield	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3		
σ_A^2	69.46	19.9	8.97	75.72	18.22	7.80		
σ^2_{M}	-	46.86	31.89	-	51.05	39.46		
σ_{AM}	-	-	16.91	-	-	12.51		
σ_{E}^{2}	89.36	91.84	99.56	87.75	94.46	100.11		
σ_{P}^{2}	158.82	158.6	157.33	163.47	163.74	159.88		
$h^2 = \sigma_A^2 / \sigma_P^2$	0.4373	0.1255	0.0570	0.4632	0.1113	0.0488		
s.c.	0.1973	0.2176	0.3349	0.2124	0.2482	0.3826		
$m^2 = \sigma^2_M / \sigma^2_P$	-	0.2955	0.2027	-	0.3118	0.2468		
s.e.	-	0.1289	0.3210	-	0.1354	0.3307		
$c^2 = \sigma_{AM} / \sigma_P^2$	-	-	0.1074		-	0.07 83		
S.C.	-	-	0.3547	-	-	0.3793		
Log L	-1638.97518	-2426.999572	-2426.89184	-1608.597984	-2396.421798	-2396.401496		
Protein	Model 1	Model 2	Model 3	Model 1	Model 2	Model 3		
σ_{A}^{2}	21.05	0.0001094	0.11	25.18	0.0035	0.27615		
$\sigma^{2}{}_{M}$	-	25.6	27.31		26.93	29.90		
σ_{AM}	-	-	-1.73	-	-	-2.87		
σ_{E}^{2}	62.83	59.24	59.1	61.31	59.75	59.45		
σ_{P}^{2}	83.87	84.83	84.78	86.49	86.68	86.75		
$h^2 = \sigma_A^2 / \sigma_P^2$	0.2509	0.0000	0.0013	0.2911	0.0000	0.0032		
s.e.	0.1963	0.2521	0.8702	0.2121	0.2634	0.4146		
$m^2 = \sigma_M^2 / \sigma_P^2$	-	0.3017	0.3221	-	0.3106	0.3446		
s.c.	-	0.1452	0.5777	•	0.1443	0.3876		
$c^2 = \sigma_{AM} / \sigma_P^2$	-	•	-0.0205	-	-	-0.3306		
s.c.	-	-	0.8389	-	-	0.4311		
Log L	-1563.37809	-2350.806034	-2350.79734	-1535.572401	-2322.905296	-2322.894702		

Table 33.	Variance com	popents and	l resulting	parameter	estimates	for Nubian
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Variance and covariance components, resulting parameter estimates and maximum log likelihood values for Nubian are given in Table 33. LogL values were generally smaller for basic model 2.

Cumulative milk yield

Fitting phantom groups in the model (BM2) increased estimates for σ_A^2 by 19.7% and σ_P^2 just slightly. Estimates for \mathbf{h}^2 were 25.6% without phantom group coefficients and 29.7% with coefficients included. Fitting maternal genetic effects in the model changed the importance of direct and maternal effects totally. Direct heritability had estimates close to zero and for \mathbf{m}^2 the results were 29.2 and 30% of the total variation. Including the covariance between maternal and direct effects increased direct heritability and decreased \mathbf{m}^2 slightly. The covariance was small (4 and 3.1%) and in both cases it was not statistically significant. Fitting phantom groups did not change the estimates for \mathbf{h}^2 and \mathbf{m}^2 . LogL values changed very little with the covariance included as an additional effect.

Cumulative fat yield

Results for fat yield showed similar trends than for milk yield. Direct heritability estimates of 43.7 (BM1) and 46.3 % (BM2) seem to be reasonable. Standard errors were relatively large, but were expected because of the small number of animals in the data set. Including maternal effects in the models resulted in small estimates for h^2 and large ones for m^2 . Adding σ_A^2 and σ_M^2 together yielded almost the same value for σ_A^2 when σ_A^2 fitted alone in sub-model1. The estimates for the covariance fitted in the models were positive and counted for 10.7 and 7.8 % respectively. LogL values changed very little when σ_{AM} was added as an in the model, but decreased significant when maternal effects were included in addition to direct effects.

Cumulative protein yield

The direct genetic effect accounted for 25.1 (BM1) and 29.1 % (BM2) of the total variation. Including maternal effects reduced σ_E^2 but increased σ_P^2 in both basic models. Direct heritability estimates were again close to zero, and maternal heritability estimates

were high (30.2 and 31.1 %). Including the covariance between the two effects in the model changed values for σ_A^2 and σ_M^2 very little. Estimates for the covariance were small and negative for BM1, high and negative for BM2 and were not statistically significant. Log likelihood values indicate, for all traits, that Model 1 (just fitting direct genetic effects in addition to fixed effects) described the data set best for Nubians.

Genetic correlations between direct and maternal genetic effects for cumulative milk, fat and protein yields are given in Table 34.

Table 34. Genetic correlation between direct and maternal genetic effects - Nubian

~~~~~~~~	Basic Model 1	Basic Model 2
Cumulative milk yield	0.9984	0.9984
Cumulative fat yield	0.9995	0.7131
Cumulative protein yield	-1.0000	-0.9982

The genetic correlations between direct and maternal effects were very high and positive for milk and fat yield. For protein yield, the correlations were also very high but negative. Fitting phantom group effects in the model reduced estimates just slightly.

The proportions between direct and maternal heritability, estimated for Nubians, are very unlikely and it may be that maternal and direct effects should be seen as two random effects, which are very closely related to explain these results. Genetic correlations indicated that increase in  $\sigma^2_A$ , also caused an increase in  $\sigma^2_M$ , but maybe the second random effect will be just a constant part of the other animal effect.

As with other breeds, the best models to describe the data set are, models just fitting direct effects in addition to the fixed effects. The covariance between direct and maternal effects was not statistically significant in both basic models. Estimates for maternal effects were statistically significant for milk and fat yield on a probability level of 5 % and for protein yield on a 2.5 % probability level (Appendix Table A19).

#### 4.7.5 Combined data set - ALL BREEDS

#### Basic Model 2 Basic Model 1 Model 1 Model 2 Model 3 Milk yield Model 1 Model 2 Model 3 $\sigma^2$ 17065.17 10504.21 8425.78 18721.49 12458.69 10966.02 $\sigma^2_{\rm M}$ 3853.92 5607.15 3766.79 5230.88 -2302.7 1773.81 $\sigma_{AM}$ _ - $\sigma_{E}^{2}$ 80892.71 81862.59 81332.06 82196.83 83045.66 80184.91 $\sigma_P^2$ 98397.23 98456.34 97540.93 98906.41 98582.28 98308.19 $h^2 = \sigma^2_A / \sigma^2_P$ 0.1114 0.1734 0.1068 0.0864 0.1893 0.1264 s.c. 0.1025 0.0676 0.0817 0.0976 0.0667 0.0800 $m^2 = \sigma^2 \sqrt{\sigma^2}$ 0.0386 0.0531 0.0391 -0.0570 s.e. 0.0457 0.0717 0.0697 0.0461 $c^2 = \sigma_{AM} / \sigma_P^2$ -0.0236 -0.0180 0.0684 s.e. 0.0653 --17333.49547 -17284.84849 Log L -14123.03122 -17284.88984 -14171.83008 -17333.56262 Fat yield Model 1 Model 3 Model 1 Model 2 Model 3 Model 2 $\sigma^2$ 12.12 7.37 28.84 13.94 9.15 26.82 $\sigma^2_{\rm M}$ 12.98 7.80 12.95 7.59 ---6.16 $\sigma_{AM}$ --6.23 . $\sigma_{E}^{2}$ 107.16 103.9 104.99 108.27 102.55 103.99 $\sigma^2_P$ 130.26 130.72 130.06 129.47 131.39 130.91 $h^2 = \sigma_A^2 / \sigma_P^2$ 0.2052 0.0569 0.2195 0.1065 0.0702 0.0932 s.e. 0.0818 0.0665 0.0767 0.0802 0.0664 0.0784 $m^2 = \sigma^2_M / \sigma^2_P$ 0.0599 0.0992 0.0996 0.0586 s.c. 0.0486 0.0634 0.0474 0.0638 - $c^2 = \sigma_{AM} / \sigma_P^2$ -0.0481 -0.0473 0.0528 0.5381 s.e. -11617.22787 Log L -8457.259255 -11617.56213 -8479.511076 -11639.73877 -11639.38255 Protein vield Model 3 Model 1 Model 2 Model 3 Model 1 Model 2 $\sigma^2$ 6.20 5.41 14.34 4.76 3.75 15.83 $\sigma^2_{\rm M}$ 8.35 7.18 8.29 7.30 ---1.15 $\sigma_{AM}$ 1.38 --- $\sigma_{E}^{2}$ 76.70 76.09 76.77 77.6 75.11 75.84 $\sigma_{P}^{2}$ 90.43 89.88 89.91 90.94 90.34 90.56 $h^2 = \sigma^2_A / \sigma^2_P$ 0.0597 0.0529 0.0417 0.1741 0.0687 0.1586 s.e. 0.0779 0.0907 0.0655 0.0760 0.0801 0.0678 $m^2 = \sigma^2_M / \sigma^2_P$ 0.0806 0.0799 0.0918 -0.0929 s.e. 0.0467 0.0165 0.0475 0.0696 $c^2 = \sigma_{AM} / \sigma_P^2$ 0.0153 -0.0127 -0.0618 **s.c**. 0.0511 . Log L -11311.98594 -11332.09453 -8151.420213 -11312.00710 -8171.690092 -11332.12783

# Table 35. Variance Components and resulting parameter estimates for combined data set (all breeds).

The estimates of (co) variance components and genetic parameters for cumulative milk, fat and protein yield, for the combined data set, are presented together with the maximum (log) likelihood values, for each analysis in Table 35.

For all traits and both basic models, ignoring maternal effects (Model 1) resulted in substantially higher estimates of  $\sigma_A^2$  and  $h^2$ . Estimates of  $\sigma_A^2$  and log likelihood values were higher using models in which phantom groups were fitted.

### Cumulative milk yield

Fitting maternal effect (Model 2) decreased values of logL markedly over those for Model 1. Heritabilities of 17.3 (BM1) and 18.9 % (BM2) were reduced to 10.7 and 12.6 % respectively when maternal effects were included. Maternal effects accounted for 5.7 and 5.3 % of the total variation in cumulative milk yield. Values of  $\sigma_A^2$  and  $\sigma_E^2$  were reduced correspondingly. Estimates of the genetic covariance between direct and maternal effects were very small (2.3 and 1.8 %) and changed the likelihood values little compared to models ignoring  $\sigma_{AM}$ . Fitting phantom groups resulted in a slightly smaller value for  $\sigma_M^2$  and  $\sigma_{AM}$ .

#### Cumulative fat yield

Heritabilities of 20.5 and 21.9 % were obtained with both basic models, fitting just additive direct genetic effects in addition to the fixed effects (M1). Including maternal effects resulted in 10 and 9.9 % of the total variation being contributed by  $\mathbf{m}^2$ , which reduced  $\mathbf{h}^2$  to 9.3 and 10.6 % respectively. These results suggested that  $\mathbf{m}^2$  constituted a bigger portion of the total heritability for cumulative milk yield than  $\mathbf{h}^2$ . Allowing for maternal effects (Model 2) decreased logL by 27%. On the other hand, including  $\sigma_{AM}$  in the model increased logL values just a little compared to models ignoring the covariance between direct and maternal effect.

Estimates of the genetic covariance, were small (4.8 and 4.7 %) compared to  $\sigma_{P}^2$ , but were of similar size to  $\mathbf{h}^2$  (5.6 and 7.0 %) and  $\mathbf{m}^2$  (5.9 and 6.0 %).

#### Cumulative protein yield

Results for cumulative protein yield showed same trends as cumulative milk and fat yield. Ignoring maternal effects resulted in higher estimates for  $\mathbf{h}^2$ . For both basic models, values for  $\mathbf{m}^2$  of 9.2 % of the total variation reduced  $\mathbf{h}^2$  from 15.9 to 5.3 % for BM1 and from 17.4 to 6.9 % for BM2. These results suggest again that maternal heritability contribute a bigger part to the total heritability than the direct heritability. Allowing for covariance between maternal and direct effects reduced estimates for  $\sigma_A^2$  and  $\sigma_M^2$ . The estimates for  $\sigma_{AM}$  of 1.5 and 1.3 % were just a small part of the total variance  $\sigma_P^2$ .

Genetic correlations between direct and maternal genetic effects for cumulative milk, fat and protein yields are given in Table 36.

	Basic Model 1	Basic Model 2
Cumulative milk yield	0.4087	0.2729
Cumulative fat yield	0.8330	0.7290
Cumulative protein yield	0.2653	0.1826

Table 36. Genetic correlation betwee	a direct and maternal genetic effects
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The genetic correlations between direct and maternal effects were positive and small to high for all traits.

As with the single breed analyses, the best model to describe the data set is the one fitting just direct genetic effects in addition to the fixed effects. Included maternal effects and covariance between direct and maternal effects decreased maximum likelihood values and decreased the goodness of fit. Maternal effects were not statistically significant for milk yield, but significant (p<0.05) for fat and protein yield.

#### 4.8 Best estimates

"Best estimates", are results calculated with the most appropriate model. Within models the best model to describe the data set is just fitting additive genetic variance. The F-test was used to determine if phantom groups have an effect or not. Calculated F-values are given in appendix Table A21. Results demonstrated clearly, that just for the multiple breed analysis, phantom groups had an effect and the null hypothesis could be rejected. Parameter estimates and calculated standard errors from the 'best' model for each breed and trait are summarised in Table 37.

Breed	Trait	h ² ± s.e
Alpine	Milk yield	0.303 ± 0.125
	Fat yield	$0.187 \pm 0.128$
	Protein yield	$0.212 \pm 0.126$
Toggenburg	Milk vield	0.171 ± 0.153
	Fat yield	$0.201 \pm 0.139$
	Protein yield	$0.166 \pm 0.148$
Saanen	Milk yield	f*
	Fat yield	0.094 ± 0.162
	Protein vield	$0.040 \pm 0.184$
Nubian	Milk yield	$0.256 \pm 0.216$
	Fat yield	$0.437 \pm 0.197$
	Protein yield	0.251 ± 0.196
All breeds	Milk yield	0.189 ± 0.068
	Fat yield	0.220 ± 0.066
	Protein yield	$0.174 \pm 0.068$

Table 37. Heritabilities for milk, fat and protein yield by breed

*f = estimation failed

Estimated heritabilities are well within the range of estimates from the literature summarised in Table 6. Boldman et al. (1984) found a similar heritability (0.25) using Alpine, Saanen, Toggenburg, Nubian and LaMancha goats with the BLUP method. Even Garcia (1971) who used a small number ( $\Rightarrow$  high standard errors) of the same dairy breeds, found similar results for milk (0.17 ± 0.20) and fat production (0.22 ± 0.20). Kennedy et al (1982) divided genotypes into dairy breeds and dual purpose breeds and analysed the data separately with the MINQUE method. They reported very high estimates for milk and fat yield (0.69 and 0.62) for the dairy breeds, whereas estimates for Nubian (0.32 and 0.51) correspond with the results obtained in this study. A study with Canadian dairy goats on first lactation reported heritabilities of 0.38 ± 0.10 for milk yield, 0.29 ± 0.14 for fat yield and 0.31 ± 0.10 as an average for all breed (Sullivan, 1988). A study using only Alpine does reported very high heritabilities for milk yield, fat and protein yield (0.60, 0.47 and 0.47). Possible reasons for different results can have many explainable. For example

management, number of lactation, age at first kidding and/or through different statistical methods used etc.

Alpine results are in this study the most appropriate estimates for the single breed analyses. The distribution of animals per flock was good and they had the largest number of flocks. Sire and dam distribution per flock was similar to Toggenburg, but the Toggenburg data had one weak point, i.e. that 34 % of the animals with records came out of one flock.

In general, each heritability estimate is based on the degree of resemblance among related animals vs. non-related animals in a population. If relationships exist between animals and are not recorded, estimates will be biased downwards (Massey, 1993), which explains why heritability estimates for all traits for Saanen are unexpectedly low. In this study, Saanen and Nubian data sets were very small, but Nubians had the advantage of a more complete (bigger) pedigree file, which resulted in more acceptable results than Saanen. Also a failure to account for environmental contributions might reduce the estimates of heritability.

The estimation of variance components for the combined data set was made under the assumption that all breeds have the same genetic variation. Results of the single breed analyses showed that there are differences concerning variance components, which might results in a bias, when all breeds are combined together. Genetic evaluation in Canada is undertaken with a data set containing all registered breeds, only as breed as an additional fixed effect. For future evaluations and when more data are available, we should consider analysing goat data separately by breed, similar to dairy cattle.

#### 4.9 Maternal effects

Studies show that maternal effects are present in beef cattle, and sheep. Using – twice the log likelihood value and comparing it to the tabulated  $\chi^2$  values, gave the following results; maternal effects for Alpine, for all traits and for both basic model were not statistically significant (Appendix Table A16).

Same results for Toggenburg, maternal effects for all traits and both basic models were not statistically significant (Appendix Table A17).

For Saanen maternal effects were not statistically significant for all traits and both basic models (Appendix Table A18).

Nubian is the only breed where the estimates for maternal effects on milk, fat and protein yields were highly significant (milk and fat p<0.05 and protein p<0.025) (Appendix Table A19). The significance levels found, indicate that there are maternal effects in Nubian, but the high estimates for  $m^2$  obtained in this study are doubtful and further investigation based on large numbers is warranted.

The combined data set containing all breeds indicates similar results. Maternal effects on milk yield for both basic models are not significant (p<0.25), but maternal effects on fat and protein yield are significant at a 5 % probability level (Appendix Table A20).

Van Vleck and Bradford (1966) using Holstein data found within-flock estimates of heritabilities from records expressed as deviations from flock-mate averages of 0.37, 0.30, and 0.24 from daughter-dam regressions for the first three lactations, and corresponding estimates from half-sib correlation of 0.24, 0.21, and 0.23. Differences between these results suggested a large maternal effect in first lactation, a smaller one in second lactation and almost no effect in third lactation.

A second study made by Van Vleck and Hart (1966) used covariances among first lactation milk records expressed as deviation from flock mate averages of Holstein cows related as cousins of varying degree, as daughter-dam, as full and maternal sibs, and as aunt-niece of varying degrees, to examine whether maternal genetic effects are important.

Results suggested that only additive genetic effects were important for milk yield.

Kirkpatrick et al. (1988) agreed with Reed and Van Vleck's (1987) conclusion, i.e. that there is no influence of cytoplasmic genes on lactation traits, therefore assuming no maternal influence on lactation.

More complicated models, which include maternal genetic effects and persistent environmental effects, could have been used to account more adequately for source of variation (Riska et al, 1985). Using such models would have complicated the analyses without changing the results, since maternal genetic effects would have contributed equally to heritability estimates (Kirkpatrick et al, 1988). Critical for this kind of analysis is that, the data set is large with relationships among the animals. Willham (1963) stated that cousins are the best relationships to test for maternal effects.

In this study, insufficient numbers of records were available to give reliable results. These results indicate that further research, is necessary to find reliable estimates for maternal heritability.

#### 4.10 Phantom groups

Including coefficients for phantom parents in the model resulted in general slightly higher results. This was a somewhat expected result, because of small data sets and the short-recorded time period (18 years) for Canadian dairy goats.

In this study the missing parents were grouped by assigned birth dates calculated using the four different pathways for genetic intervals. The small number of animals and the fact that milk recording for Canadian dairy goats started in 1980 and just a few phantom parents had to be assigned after 1980 led to the decision of using 4 groups for sire and 4 groups for dams. Reports of rules on how many groups should be assigned without over parameterisation were not found. Calculations done earlier with 23 sire and 23 dam groups, each of the 23 groups representing one year, gave fixed effects solutions for the groups which were unrealistically high (200000 kg for milk yield). Westell (1984) using a data set of 1,074,971 animals in her study assigned 12 sire and 12 dam groups. The point is that grouping phantom parents might be sensible only in a large population, so that groups can represent the population mean in their time period more accurately.



Figure 17. GLS solution for sire and dam phantom groups for Alpine



Figure 18. GLS solution for sire and dam phantom groups for Toggenburg



Figure 19. GLS solution for sire and dam phantom groups for Saanen



Figure 20. GLS solution for sire and dam phantom groups for Nubian



Figure 21. GLS solution for sire and dam phantom groups for all breeds

A good agreement between solutions for groups and the mean yield would indicate that the group accurately reflects the mean genetic merit of the animals constituting them. (Golden et al., 1994)

Group solutions plotted for milk yield and by breed in Figures 17 to 21 indicate no particular trend, but show the effect of each group on the milk yield. Values plotted for fat and protein yield are not included in the results, because they show a similar pattern though on a lower level (Appendix). Further studies have to be made to investigate the connection between the number of assigned groups and the accuracy of the resulting estimates.

Solutions for phantom groups with standard errors for milk, fat and protein yield are given in Table A-15 (Appendix).

#### 5. CONCLUSIONS

- Results on estimates of variance components indicate that estimates of the covariance between direct additive and maternal additive genetic effects are small and not statistically significant for cumulative milk yield, cumulative fat yield and cumulative protein yield in all Canadian dairy goats used in this study. Maternal effects seem to be significant for Nubian and the all breed combined data set. Heritabilities are moderate to high for Alpine (0.23 to 0.31) and Toggenburg, (~ 0.19), and high for Nubian. (0.29 to 0.46). Information on Saanen dairy goats was not sufficient to provide useful estimates. If maternal genetic effects exist heritability would be biased.
- □ Genetic evaluation for dairy goats in Canada is made under the assumption that genetic variances are the same in all breeds. This study reveals that the variances are different for each breed and future evaluation, when more data are collected should be done separately.
- Appropriate statistical models should be used to account for all known important environmental effects like, age at first kidding, flock-year or even genetic trends to get unbiased estimates for genetic parameters such as heritabilities. Phantom groups are an efficient way to make results more accurate through accounting for selection. For effective genetic improvement, and more reliable databases, more data on dairy goats and their pedigree has to be collected.
- Before using phantom parent grouping in genetic evaluation in dairy goats, further studies should ensure that the grouping does not introduce systematic errors (such as over parameterisation) into current analyses. Milk recording for dairy goats started 1980 and the last animal assigned to a group was from 1986. It is doubtful that there was a lot of genetic progress to account for in six years, and maybe genetic grouping based on phantom parents was not necessary.

Estimates of maternal effects and covariance between direct and maternal effects were not statistically significant. Further studies have to be carried out to investigate these uncertain trends in this study. If evidence for negative correlation between direct and maternal effects were to be found, methods of selection accounting for both direct and maternal effects would result in greater economic response to long term selection than selection based only on direct genetic effects.

### **6. LITERATURE CITE**

Agriculture Canada, 1989, Goats and their management, Publication 1820/E, Communications Branch, Ottawa K1A 0C7

Alderson, A. and Pollak, E.J., 1980, Age-season adjustment factors for milk and fat of dairy goats. Journal of Dairy Science, 63:148

Ali, A. K. A., Mohammad, W. A., Grossman, M. and Shanks R.D., 1983, Relationships among lactation and reproduction traits of dairy goats. Journal of Dairy Science, 66:1926

Amoah, A. and Bryant M.J., 1983, Gestation Period, Litter Size and Birth weight in goats. Animal Production, 36:105

Amoah E.A., Gelaye S., Guthrie P. and Rexroad C.E. Jr., 1996, Breeding Season and Aspect of Reproduction of female goats. Journal of Animal Science, 74:723

Analla, M., Jimenez-Gamero I. and Munoz-Serrano, A., 1996, Estimation of genetic parameters for milk yield and fat and protein contents of milk from Murciano-Granadina goats. Journal of Dairy Science, 79:1895

Anifantakis, E.M., and Kandarakis J.G., 1980, Contribution to the study of the composition of goat's milk. Milchwissenschaft, 35:617

Auran, T., 1976, Studies on monthly and cumulative monthly milk yield records. III. Estimates of genetic and phenotypic parameters. Acta Agriculturæ Scandinavica. 26:3

Banks, B.D., Mao I.L. and Walter J.P., 1985, Robustness of the restricted maximum likelihood estimator derived under normality as applied to data with skewed distribution. Journal of Dairy Science, 68:1785

Boldman K.G., and Van Vleck L.D., 1984, Genetic trend for milk yield from doe evaluations in the northeast United States. Journal of Dairy Science, 67:1069

BonDurant, R.H., Darien B.J., Munro G.H., Stabenfeld G.H. and Wang P., 1981, Photoperiod indication of fertile oestrus and changes in LH and progesterone concentrations in yearling dairy goats (Capra hircus). Journal of Reproduction Fertility, 63:1

Bouillon, J., Ricordeau G., 1976, Parameters genetiques des performences de croissance et de production laitiere chez les caprins en station de testage. Estimation des responses directes ou indirectes a la Selection. Ann. Genet. Selec. Anim., 8:109

Bourdon Richard M., 1997, Understanding animal breeding, Prentice-Hall, New Jersey, page 170

Casoli, C., Duranti E., Morbidini L., Panella F. and Vizioli V., 1989, Quantitative and compositional variations of Massese sheep milk by parity and stage of lactation. Small Ruminant Research, 2:47-62

Choveiri, G., 1973, The Damascus goat, the future milking animal of the Mediterranean countries. III World Conference Anim. Prod. Pre Conf. Vol. No. 1, Melbourne, Australia

Colby, B.E., Evans D.A., Lyford S.J., Nutting W.B., Stern D.N., Dairy goat breeding/ feeding/ management. Leaflet 439. American Dairy goat Association

Cue, R.I., 1998, Course notes on Estimation of Genetic Parameters, Department of Animal Science, McGill University, Macdonald Campus, Montreal, Canada

Devendra, C., 1975, Biological efficiency of milk production in dairy goats. World Review of Animal Production, 11:46

Dickerson, G.E., 1947, Composition of hog carcasses as influenced by heritability differences in rate and economy gain. Iowa Agr. Exp. Sta. Res. Bull. 354-489

Dickinson, F.N., and King G.J., 1977, Phenotypic parameters of dairy goat lactation records. Paper presented 72nd Annual Meeting Amer. Dairy Science Ass., Ames, IA. June 27

Eisen, E.J., 1967, Mating designs for estimating direct and maternal genetic variance and direct-maternal genetic covariance. Canadian Journal of Genetic and Cytology 9:13

Eler, J.P., Van Vleck L.D., Ferraz J.B.S. and Lobo R.B., 1995, Estimation of variances due to direct and maternal effects for growth traits of Nelore cattle. Journal of Animal Science, 73: 3253

Falconer, D.S., 1965, Maternal effects and selection response. In S. J. Gurts (Ed.) Genetics today., Pergamon Press, New York.

Falconer, D.S. and Mackay T.F.C., 1996, Introduction to quantitative genetics, 4th edition, Longman, Essex, England

Finley, C. M., Thompson J.R. and Bradford G.E., 1984, Age-parity-season adjustment factors for milk and fat yields of dairy goats. Journal of Dairy Science, 67:1868

Gall, C., 1973, Relationship between body conformation and milk yield in German Improved Fawn goats. L'Institute Technique de L'Elevage Ovine et Caprine : 287

Gall, C., 1996, Goat breeds of the world, Markgrafverlag, Weikersheim

Garcia, B.O., 1971, Heritabilities of milk and butter fat production in goats. III Reun. Lat. Amer. Prod. Anim.; Bogota, page113

Garcia, O., Castillo J., and Gado, C., 1972, Situacion actual de la gonaderia caprina en Venezuela, Agron. Trop., 22:239

Geerts, H.A., 1975, Is milk production Improving? Dairy Goat Journal, 53:9

Gianola, D. and Fernando, R.L, 1986, Bayesian methods in animal breeding theory. Journal of Animal Science, 63:217

Gill, G.S. and Dev, D.S., 1972, Performance of two exotic breeds of goats under Indian conditions. Indian Journal of Animal Production, 3:173

Gipson, D. and Russel W.S., 1978, The performance of cross breed dairy cattle, Animal Breeding Research Organisation, Report 29

Gipson, T. A. and Grossman M., 1989, Diphasic analysis of lactation curves in dairy goats. Journal of Dairy Science, 72:1035

Grossman, M. and Wiggans G.R., 1981, Dairy goat lactation records and potential for buck evaluation. Journal of Dairy Science, 63:1925

Guss, S.B., 1975, Dairy goat herd health problems. Journal of American Vet. Med. Ass: 167:1076

Haenlein, G.F.W., 1981, Dairy goat industry of the United States. Journal of Dairy Science, 64:1288

Haenlein, G.F.W., 1995, Nutritional Value of Dairy Products of Ewe and Goat Milk, Book: International Dairy Federation

Haenlein, G.F.W., 1996, Status and Prospects of the Dairy goat Industry in the United States. Journal of animals Science 74:1173

Harville, D.A., 1977, Maximum likelihood approaches to variance component estimation and to related problems. Journal of American Statistic Association, 72: 320

Hayden, T.J., Thomas C.R. and Forsyth, I.A., 1979, Effect of number of young born (litter size) on milk yield of goats: role of placental lactogen. Journal of Dairy Science, 62:53

Henderson, A.J., Blatchford D.R. and Peaker M., 1983, The effect of milking thrice instead of twice daily of milk secretion in the goat. Q. J. Exp. Physiol., 68:645

Henderson, A.J. and Peaker M., 1984, Feedback control of milk secretion in the goat by a chemical in milk. Journal of Physiology, 351:39

Henderson, A.J., Blatchford D.R. and Peaker M., 1985, The effect of long term thrice daily milking on milk secretion in the goat: evidence for mammary growth. Q. J. Exp. Physiology., 70:557

Henderson, C.R., 1972, Sire evaluation and genetic trends. Proc. Anim. And Genet. Symp. In Honour of Dr. Jay L. Lush ASAS-ADSA, Champaign, IL.

Henderson, C.R., 1975, Use of all relatives in intraherd prediction of breeding values and producing abilities. Journal of Dairy Science, 58:1910

Henderson D., 1985, Control of the breeding season in sheep and goat, In Practice, July 85:118

Holtmann, W.B. and Bernard, C. 1969. Effect of general combining ability and maternal ability of Oxford, Suffolk and North Country Cheviot breeds of sheep on growth performance of lambs, Journal of Animal Science, 28:155-161.

Howson, C. and Urbach, P, 1989, Scientific Reasoning: The Bayesian Approach. Open Court Publishing Co., LaSalle, IL

Iloeje, M.U. and Van Vleck, L.D., 1978, Genetics of dairy goats: A review. Journal of Dairy Science, 61:1521

Iloeje, M.U., Rounsaville T.R., McDowell R.E., Wiggans G.R. and Van Vleck L.D., 1980, Age-season adjustment factors for Alpine, LaMancha, Nubian, Saanen and Toggenburger dairy goats. Journal of Dairy Science, 63:1309

Iloeje, M. U., Van Vleck, L.D. and Wiggans G.R., 1981, Components of variance for milk and fat yields in dairy goats. Journal of Dairy Science, 64:2290

Ingraham, R.H. and Stanley, R.W., 1979, Seasonal effects of tropical climate on shaded and non-shaded cows as measured by rectal temperature, adrenal cortex hormones, thyroid hormone, and milk production. American Journal of Veterinarian research, 40:1792

Jairath, L.K., 1992, Genetic and phenotypic parameters of lifetime performance traits in Canadian Holsteins, Ph.D. Thesis, Macdonald Campus of McGill University, Montreal, Canada

Johnson, H.D., 1965, Environmental temperature and lactation (with special reference to cattle), Int. Journal of Biometerologie, 9:103

Johnson, J.C. Jr, and Southwell B.L., 1962, Interrelationships of certain climatic conditions and productive responses on lactating dairy cows, Journal of Dairy Science, 45(Suppl.1): 695. (Abstr.)

Juarez, M. and Ramos M., 1986, Physio-chemical characteristics of goat's milk as distinct from those of cow's milk. Bulletin of International Dairy Federation, 202:54

Kempthorne, O., 1955, The correlation between relatives in random mating populations. Cold Spring Harbor Symp. Quant. Biol. 22:60

Kennedy, B.W., Finley C.M., Pollak E.J. and Bradford G.E., 1981, Joint effects of parity, age and season of kidding on milk and fat yields of dairy goats. Journal of Dairy Science, 64:1707

Kennedy, B.W., Finley C.M. and Bradford G.E., 1982, Phenotypic and genetic relationship between reproduction and milk production in dairy goats. Journal of Dairy Science 65:2373

Kirkpatrick, B.W. and Dentine M.R., 1988, An alternative model for additive and cytoplasmic genetic and maternal effects on lactation. Journal of Dairy Science, 71:2502

Knowles, F. and Watkins, J.E., 1938, The milk of the goat under English conditions. Journal of Dairy Research, 9:153

Koch, R.M. and Clark, R.T., 1955, Genetic and environmental relationship among economic characters of beef cattle. III. Evaluating maternal environment. Journal of Animal Science, 14: 979

LKV (Landeskreisverband Bayern), 1993, Jahresbericht, München

Ley, W.B, 1986, Which factors figure most prominently in caprine breeding management? Veterinary Medicine, June 1986.

Maria, G.A., Boldman K.G. and Van Vleck L.D., 1993, Estimates of variances due to direct and maternal effects for growth traits of Romanov sheep. Journal of Animal Science, 71:845

Massey, J.W. and Vogt D.W., 1993, Heritability and its use in Animal breeding, Department of Animal Science, University of Missouri-Columbia, http://muextension.missouri.edu/xplor/ogguides/ansci/g02910.htm

Mavrogenis, A. P. Constantinou, A. and Louca, A., 1984, Environmental and genetic causes of variation in production traits of Damascus goats: goat productivity. Animal Production, 38:99

McDonald, M.A. and Bell J.M., 1958, Effects of low fluctuating temperature on farm animals. IV. Influence of temperature on milk yield and milk composition, Canadian Journal of Animal Science, 38:160

McDowell, R.E., Hoover N.W. and Cameons J.K., 1976, Effect of heat stress on energy and water utilisation of lactating cows. Journal of Dairy Science, 59:188

Mittal, J.P. and Pandey M.D., 1971, The yield and chemical composition of milk in Barbari and Jamnapari goats. Agra. J. Res. 20:7

Meyer, K., 1986, Between Algorithms: A "Short Cut" Restricted Maximum Likelihood Procedure to Estimate Variance Components. Journal of Dairy Science 69: 1904-1916

Meyer, K., 1989, Restricted Maximum Likelihood to estimate variance components for animal model with several random effects using derivative free algorithm. Genetic Selection Evolution 21:317

Meyer K., 1992, Variance Components due to direct and maternal effects for growth traits of Australian beef cattle. Livestock Production Science, 31:179

Meyer, K., 1993, DFREML User notes, Version 2.1

Mohammad, W. A., Grossman M. and Vatthauer J. L., 1984, Seasonal breeding in United States dairy goats. Journal of Dairy Science, 67:1813

Morand-Fehr P. and Sauvant D., 1978, Nutrition and optimum performance of dairy goats. Livestock Production Science, 5:203-213

Morand-Fehr P. and Sauvant D., 1980, Composition and yield of goat milk as affected by nutritional manipulation. Journal of Dairy Science, 63:1671

Nadarajah, N.K., 1998, Goat production: An over view of the Canadian goat industry. http://www.aps.uoguelph.ca/~gking/goat.htm

Orlyanskii, V.A. and E.B. Zaporozhtsev, 1974, Milk production - an important selection character in Don goats. Ovtsevodstvo, 1:27

Peart, J.N., 1968, Lactation studies with Blackface ewes and their lambs. Journal of Agricultural Science, Cambridge, 74:459-469

Perotto, D., 1992, Additive and non-additive genetic effects on growth and milk production traits in Holstein Ayrshire crossbreeding experimental data. Ph.D. Thesis, Macdonald Collage of McGill University, Montreal, Canada, page 69 Pollak E.J. and Quaas, R.L., 1983, Definition of group effects in sire evaluation models. Journal of Dairy Science, 66:1503

Prakesh, C., Acharya R.M. and Dhillon J.S., 1971, Sources of variation in milk production of Beetal goats. Indian Journal of Animal Science, 41:356

Quaas, R.L., 1988, Additive genetic model with groups and relationships. Journal of Dairy Science, 71:1338

Quaas, R.L. and Pollak E.J., 1981, Modified equations for sire models with groups. Journal of Dairy Science., 64:1868

Rathore, A.K., 1970, Goats- milk production and the effect of age of milk quality, Aust. Goat W., 22:1

Rastogi, R., Boylan, W.J., Rempel, W.E. and Windels, H.F., 1982, Crossbreeding in sheep with evaluation of combining ability, heterosis and recombination effects for lamb growth. Journal of Animal Science, 54:524

Reed, P.D. and Van Vleck L.D., 1987, Lack of evidence of cytoplasmic inheritance in milk production traits in dairy cattle. Journal of Dairy Science, 70:837

Renner, E., 1983, Milch und Milchprodukte in der Ernährung des Menschen, VV-GmbH Volkswirtschaft. Verlag, München, page 467

Riska, B., Rutledge J.J. and Atchley W.R., 1985, Covariance between direct and maternal genetic effects in mice, with a model of persistent environmental influences. Genetic Research Cambridge, 45:287

Robinson, D.L., 1996, Estimation and interpretation of direct and maternal genetic parameters for weight of Australian Angus cattle. Livestock Production Science, 45:1

Robinson, G.K., 1986, Group effects and computing strategies for models with groups. Journal of Dairy Science. 69:3106

Robison, O.W., 1981, The influence of maternal effects on the efficiency of selection: A review. Livestock Production Science, 8:121

Robison, O.W., McDaniel B.T. and Rincon E.J., 1981, Estimation of direct and maternal additive and heterotic effects from crossbreeding experiments in animals. Journal of Animal Science, 52:44

Rønningen, K., 1965, Causes of variation in the flavour intensity of goat milk. Acta Agriculturæ Scandinavica, 15:301

Rønningen, K., 1967, A study of genetic parameters for milk characteristics in goats. Rep. No.232, Inst. Anim. Genet. And Breed. Agric. Collage of Norway, Vollebekk, Norway.

Sambraus, H.H., 1994, Atlas der Nutztierrassen, Ulmer Verlag, Stuttgart

Shelton, M., 1978, Reproduction and breeding of goats. Journal of Dairy Science, 61:994

Shrestha, J.N.B., Rempel, W.E., Boylan, W.J. and Miller, K.P., 1983, General, specific, maternal and reciprocal effects for ewe productivity in crossing five breeds of sheep. Canadian Journal of Animal Science, 63:497

Singh, R.N., Acharya R.M. and Biswas D., 1970, Evaluation of genetic and non-genetic factors affecting some economic traits in goat. Acta Agriculturæ Scandinavica., 20:10

Singh, B.B. and Singh, B.P., 1974, Performances of Jamnapari goats. Indian Veterinarian Journal, 51:326

Skinner, J.D., 1972, Utilisation of the Boer goat for intensive animal production. Tropical Animal Health and Production, 4:120

Steel, R.G.D and Torrie, J.H, 1980, Principles and procedures of statistics - A biometrical approach, McGraw-Hill, Inc. NY, second edition, page 579 ff.

Steine, T.A., 1975, Factors affecting characteristics of economic importance in goat. Meldinger fra Norges Landblukshægskole, 54:30

Subines, J., Lava, L., Ferrando, G. and Boza, J., 1988, [Factors affecting milk yields in goats. Lactation number and type of kidding], Archivos de Zootechnia 37:145 [Dairy Science Abstracts 51(4): 162.]

Sullivan, B.P., 1988, Estimation of variance components for lactation traits of Canadian dairy goats. M.Sc. Thesis, University of Guelph

Sutton, John D., 1990, Nutrition of the lactating goat. Journal of the Science of Food & Agriculture, 53:123

Teh, T.H., Trung, L.T., Jia, Z.H. and Gipson T.A., 1994, Varying amounts of Rumeninert fat for high producing goats in early lactation. Journal of Dairy Science, 77:253

Van Vleck, L.D. and Bradford G.E., 1966, Genetic and maternal Influences on the first three lactations of Holstein Cows. Journal of Dairy Science, 49:45

Van Vleck, L.D. and Hart C.L, 1966, Covariances among first-lactation milk records of cousins. Journal of Dairy Science, 49:41

Van Vleck, L.D., 1971, Index selection for direct and maternal genetic components of economic importance. Biometrics, 27:477

Van Vleck, L.D., 1990, Breeding value prediction with maternal genetic groups. Journal of Animal Science, 68:3998

Van Vleck, L.D., Gregory K.E. and Bennett G.I., 1996, Direct and maternal genetic covariances by age of dam for weaning weight. Journal of Animal Science, 74:1801

Vesely, J.A., Kozub, G.C. and Peters, H.F., 1977, Additive and non-additive genetic effects on growth traits in matings among Romnelet, Columbia, Suffolk and North Country breeds. Canadian Journal of Animal Science, 57:233

Voutsinas, L., Pappas C. and Katsiari M., 1990, The composition of Alpine goats milk during lactation in Greece. Journal of Dairy Research, 57:41

Waite R. and Blackburn P.S., 1963, The relationship between milk yield, composition and tissue damage in a case of subclinical mastitis. Journal of Dairy Research, 30:23

Westell, R.A., 1984, Simultanous genetic evaluation of sire and cows for a large population of dairy cattle. Ph.D. Thesis, Cornell University

Westell, R.A., Quaas, R.L. and Van Vleck L.D., 1988, Genetic Groups in an animal model. Journal of Dairy Science, 71:1310

Willham, R.L., 1963, The covariance between relatives for characters composed of components contributed by related individuals. Biometrics, 19:18

Willham, R.L., 1972, The role of maternal effects in animal breeding: III. Biometrical aspects of maternal effects in animals. Journal of Animal Science, 35:1288

Williams, J. C., 1993, Influence of farm, parity, season and litter size on the lactation curve parameters of white British dairy goats. Animal Production, 57:99

Zygoyiannis, D., and Katsaounis, N., 1984, The influence of number and genotype of lambs suckled on milk yield and milk composition of Karagouniko ewes. Scientific Yearbook of Veterinary Faculty, Aristotelian University of Thessaloniki, 22:365

Zygoyiannis D., 1994, A note on the effect of number and genotype of kids on milk yield and composition of indigenous Greek goats (Capra prisca). Animal Production, 58:423

# <u>APPENDIX</u>

# I. Phantom Groups

#### Table A1. Number of animals assigned per phantom group per breed

Phantom group Number of Animals assigned per group						
Sire	Year	Alpine	Toggenburg	Saanen	Nubian	All breeds
PS1*	1957-1965	202	153	142	164	462
PS2*	1966-19 <b>7</b> 0	307	197	194	214	<b>76</b> 3
<b>PS3</b> *	1971-1975	205	88	128	184	541
PS4*	> 1975	60	37	72	101	277
Dam	Year	Alpine	Toggenburg	Saanen	Nubian	All breeds
PD1*	1957-1965	249	199	177	191	587
PD2*	1966-1970	315	186	210	230	779
PD3●	1971-1975	162	73	102	166	<b>48</b> 3
PD4*	> 1975	84	48	91	101	331

*PS = Phantom Sire groups

*PD = Phantom Dam groups

# II. Data Structure

# Table A2: Data Structure: Offspring-Parent combinations

	ALPINE	TOGGENBURG	SAANEN	NUBIAN	ALL BREEDS
NO. OF RECORDS IN DATA	691	641	439	433	2204
NO. OF PEDIGREE RECORDS	3871	2642	2472	2901	11012
AVERAGE INBREEDING COEFFICIENT	1.01873	1.02117	1.01819	1.01986	1.02059
NO. OF INBRED ANIMALS	1378	1095	863	832	4120
WITH AVERAGE INBREEDING	1.05262	1.05108	1.0521	1.06924	1.05504
COEFFICIENT					
no. of animals with great-grand parent(s)	2548	1876	1576	1 <b>74</b> 6	7549
no. of "base" animals	772	471	532	663	2034
no. of animals in the data	691	641	439	433	2204
no. of sires :					
in total	1166	722	783	967	3223
with progeny in the data	180	153	123	150	606
which are also grand sires	1099	662	738	904	2988
which also have grandprogeny records	106	90	73	76	345
paternal half-sib record pairs	5422	5048	2586	1686	14742
av. no. progeny rec.s/sire	3.84	4.19	3.57	2.89	3.64
<u>no. of dams :</u>					
in total	2135	1440	1340	1584	6040
with progeny in the data	517	455	328	320	1617
with own record as well	106	146	78	75	405
dam-offspring record pairs	134	177	122	108	542
which are also grand dams	1802	1184	1141	1388	5061
which also have grandprogeny records	162	190	118	109	579
with own record as well	19	29	24	18	90
maternal half-sib record pairs	474	524	312	304	1626
av. no. progeny rec.s/dam	1.34	1.41	1.34	1.35	1.36
no. of paternal grand sires :					
in total	547	334	386	462	1514
with progeny in the data	111	93	80	101	385
quarter-sib record pairs	7628	8276	3882	2312	22098
av. no. progeny rec.s/grand sire	6.23	6.89	5.46	4.29	5.72
no. of maternal grand sires :					
in total	811	517	526	655	2276
with progeny in the data	257	197	179	204	831
quarter-sib record pairs	3166	4200	1506	1318	10244
av. no. progeny rec.s/grand sire	2.68	3.19	2.39	2.12	2.62
<u>no. of grand sires together :</u>					
in total	1358	851	912	1117	3790
with progeny in the data	368	290	259	305	1216
quarter-sib record pairs	10794	12476	5388	3630	32342
av. no. progeny rec.s/grand sire	3.75	4.38	3.34	2.84	3.6

# Table A2: Data Structure: Offspring-Parent combinations .....continuation

no of naternal grand dame -					
in total	727	455	421	608	2047
with programs in the data	150	155	108	137	573
with sup around as well	7	140	108	137	343
with own record as wen	20	14	9	•	30
grand dam-olispring record pairs	58	42	44	16	140
quarter-sib record pairs	6196	5770	2924	1928	16818
av. no. progeny rec.s/grand dam	4.61	5.01	4.05	3.16	4.21
no. of maternal grand dams :					
in total	1279	870	790	940	3649
with progeny in the data	417	350	257	264	1283
with own record as well	27	32	28	21	108
grand dam-offspring record pairs	36	48	41	36	161
quarter-sib record pairs	906	1172	554	444	3162
av. no. progeny rec.s/grand dam	1.61	1.78	1.55	1.54	1.64
no. of grand dams together :					
in total	2006	1325	1271	1548	5691
with progeny in the data	567	478	365	401	1806
with own record as well	34	44	37	29	144
grand dam-offspring record pairs	74	90	85	52	301
quarter-sib record pairs	7102	69 <b>42</b>	3478	2372	19980
av. no. progeny rec.s/grand dam	2.4	2.64	2.29	2.09	2.38

#### Table A3: GLS - solutions for age at fist kidding - Alpine - cumulative milk, fat and

filk yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTIONS	ST. ERROR
9	7	444.71	4.6879	-395.3713302	261.4500880
10	17	533 24	11 593	-187.283100	238.5301443
11	45	630.38	31.6005	-230.4385425	227.2719430
12	116	685 42	71 8o5L	-240 9850469	219 5724469
13	123	768.83	75.7636	-234.4358537	223.8126935
14	55	670.47	36 1687	-227 8119984	227 330-4079
15	49	751.53	33.326	-230 4119588	227 0000405
10	1.1	8.11 18	23 2520	-180-5309312	779 6779607
10	17		35 2073	100 400 511	778 2020034
10	31	828.00	17 - 127	244.4/3373	131 900 5111
10		731.12	1/813/	-247 0088723	231 8094311
19	17	731.76	12.0243	-102.9783631	23/99/20/4
20	10	722.06	11.2347	-108 2006034	238 5937709
21	19	700.74	13.7305	-77 82426799	235 727549
22	35	850 29	24 5031	-104 6272978	229 9617658
23	-48	768.31	33 0324	-106.4811983	228 1710899
24	25	780 08	17.9658	-95 1100-1387	232 9697945
25	13	870.40	9.1318	-122.8316063	243 760182
20	1	1296.50	2,9953	380.9430242	275 0751 (0)
27	3	997.00	2 2041	202 3209974	284,3025362
78		868 50	1.588	51 53754644	305 2117524
70	-	762.00	0 7722	-327 634300	375 0676253
				-22,05050	5.700.0155
L YICH	NPEC	MEAN	NACONAL		67 Fasca
<u>AUE</u>	7	NEAN	UIAGONAL	GLS - SOLUTIONS	0.110-1090
10	17	1343	11 (2)(1	-(7 /4/300)4	9 575020//
10	17	18.50	13.4211	-10.2715227	8.3/39/944
11	45	21 27	35 8982	-11 70527527	8 15770-18
12	llö	24.11	83.1511	-11 94196197	7 85977535
13	123	20 35	87 5301	-11 94559162	8 0267812
14	55	23.67	41.5074	-11 40210596	8.1577091
15	-19	20 87	38 (	-11 39971706	8 16878188
lo	34	28.62	26.587	-10 19154965	8 24090638
17	37	29 51	28 7883	-11 01670658	8 21206788
18	25	27.62	20 01 19	-9 4405736	8 32445286
19	17	74 55	137105	- 51 583819	8 50-17775-1
20	16	17 55	13 8311	1 85370351	8 57077735
20	10	23 47	12.0211	-185210351	0.010-0-01
21	19	23 /4	13 5243	1 0000334	8 4/600004
22	33	30.47	21.9402	-/ 0200//63	8.24002029
25	-18	27 02	37 /913	14011561 0-	8 19229111
24	25	20 20	20.3751	-5.93870996	8.37244737
25	13	29 08	10 409	-8 44075294	8 77858309
20	4	45 61	3.3598	10.56020476	9.96467602
27	3	33.29	2.4945	3 00795901	10 31850325
28	2	28.70	1.669	-5 77122763	11.13226954
29	1	29 78	0 8264	-9 -13263-182	13 622 403 42
otein vield					
AGE	NREC	MEAN	DIACONAL	CLS. SOLUTIONS	ST. ERROR
9	7	12.34	5 2807	-13.43500944	7 88364869
10	17	10.10	130215	-7 0n67771R	7 18061511
11		18 75	1.1 04.72	-R BAI 15607	6 8177 6417
17		10.0	20 - 11	-90047480	A 5877394.4
11	110	20 80	00 021	** UBG/0085	C 75 42 440
13	123	22.23	54.574/	-6.8023011	0 /243337/9
14	22	20.35	-40 3244	-8 200-988888	6 83334212
15	49	23.54	37 0592	-8.23551034	6.84273498
10	34	25 82	25 858	-0 57821148	o 90289674
17	37	26.23	28.0017	-7 64833442	6.87851
18	25	23.10	19 4941	-8 05949004	o 97176825
19	17	21.53	13.3501	-5.3268962	7.16960352
20	10	21 32	12.4777	-5 07297026	7 1831 1081
21	19	20.45	151384	-1 62632921	7 09844803
22	15	25 47	77 105.1	-1 78213421	N 908 448 41
22		10/	34 7547	-4 701 44 443	6 8611000
23		44.98	30 / 342	~4./01.40402	0.0012983/
24	<i>D</i>	23 /0	19 85/5	-3 40/40913	101063044
25	13	25.56	10.132	-6.05914741	7 34755629
20	4	39.48	3 2819	10 07607002	8.3300569
27	3	28.93	2.4326	2.98790789	8.62198312
28	2	25.00	1 6304	-3 05995109	9 29178288
29	1	25,72	0.8043	-5 73870644	11.38057579

protein yield (Basic Model 1).

# Table A4: GLS - solutions for age at fist kidding - Toggenburg - cumulative milk, fat

# and protein yield (Basic Model 1).

Milk yield					
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTIONS	ST. ERROR
9	2	807.00	1.7114	-391.2917858	427 2726539
10	7	639 57	5 6385	-210 8592323	381.1738299
11	30	632.10	24 1932	-171.2867344	300.375272
12	48	633.85	35 61 42	-97 55257611	355 5260637
13	<b>6</b> 5	598.28	49 7293	-202.8812033	153.5664385
14	56	675.95	43.4956	-170 0475173	353 2833735
15	44	912.70	34.5434	-160.9377487	354.4610172
lo	45	812.07	35 1264	-165 4415125	353.3062552
17	39	857.95	28 8624	-205 4611471	355 3688512
18	33	832.33	26.5323	-278 1145338	350.8031578
19	36	890.33	29.0794	-129.4420734	355 6885568
20	27	788.48	22.3065	-187 4022325	357.9367071
21	32	932.47	26.1804	-10.15030822	350 4198732
22	36	1024 25	29 1283	91 1895712	355 1936884
23	46	844.33	36.5231	12.62989447	343 8535307
24	51	857 71	39 1412	-06 00583568	353 7896046
25	19	1031.95	15 4544	-3 43046683	360.5655829
20	11	088.09	9.1433	-94 6 <b>48484</b> 61	370 2136553
27	8	767.13	6.7478	-264 1949931	372.552618
28	4	773.00	3_3915	-81.9657411	405 901 437
29	1	892.00	0.8576	te.50028588	508 4169387
30	1	640 00	0.856	-150.0453375	500 034657
Fat yield					
ACE	NEEC	MEAN	DIACONAL	CLS. SOL LITIONS	ST FREOR
<u></u> 0	2	25.48	1 004	+139721067	13.9214216
10	7	19.77	5.4466	_8 R0039945	17 47558751
10	10	20.00	21 2012	-7 11347075	11 75456797
17	50	19.57	14 2677	-5 80175001	11 507/15735
13		18.05	.08.0.41	-8 80770164	11 \$3060131
1.1	5	72.00	47 01 41	-0.00170.000	11 51756890
15		20.74	11 3906	-7.75665466	11.558668.4.1
15	**	25.24	11.0550	-7 <u>1</u> 005-000	11 57179020
10	ده ۲۵	2014	33 7337	-6 0000.3482	11.32128929
17	39	27.63	27.7031	-6.33670733	11.39049197
16	33	27 43	23 0617	-10 33282029	11 03020390
19	30	30.21	28.2507	-5.50532848	11 3992/18/
20	27	23 83	21.629	-/ 31391406	11 0/443354
21	32	29.53	23.3793	-2-0701331	11 02101/09
22	96	31.15	28.1706	-1 05/55286	11.38248420
23	40	26.59	35.2678	-1.75107627	11 21 483 489
24	51	27 47	37 7423	-3 95300936	11 53721027
25	19	34.00	14.9/64	-0.56568946	11.75558701
20	11	21 99	8.8349	-1 0100.3809	12 07131673
27	8	24.51	6.0003	-11.17993540	12.14423691
28	4	23 88	3 2902	-3 06927401	13.23355939
29	1	27.96	0 8344	-2_32057773	10.57859491
30	I	20 84	0.833	-8.04314596	16.51178228
Protein yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTIONS	ST. ERROR
0	2	21.60	1.7183	-12.57193180	12.06799371
10	7	19 30	5 667	-7 35610799	10.76522011
11	30	18.62	24.3113	-5.97455332	10 17694814
12	48	17 04	35.8167	-4 81005841	10 04051547
13	o5	16.91	49 9809	-7.43585151	9.98490376
14	50	19 48	43 710	-a 56370894	9 977 40782
15	-14	25.96	34 7139	-5.94915636	10 01031519
10	45	22.57	35 2997	-0 47961079	9 9776767
17	39	24.21	29.0273	-7.47895286	10.03565144
18	33	23.00	26.6576	-9 02415000	10 07630162
19	36	25 76	29 2033	-5.09757265	10 04490892
20	27	22.44	22.4059	-7 09504547	10 10816055
21	32	26.58	26.2984	-2.13325497	10.06575112
22	30	29.32	29 2694	0 85265574	10.03101800
23	46	23.75	36.7087	-1.01453819	9.71049246
24	51	24 58	39.3495	-3 08958062	9.99129246
25	19	29.12	15.5248	-1.42864983	10 18297595
20	11	19 75	9.1856	-3.79103227	10.45530706
27	8	21_57	6.7758	-9.13496411	10.52178241
28	4	21.88	3 4063	-3 73262437	11.46335225
29	1	22.58	0.861	-2.00950095	14,35806199
30	1	19 78	0 8593	I diability to	14 30886775

# Table A5: GLS - solutions for age at fist kidding - Saanen* - cumulative milk, fat and

lilk yield	· · · · · · · · · · · · · · · · · · ·	······································			
AGE	NREC	MEAN	DEAGONAL	GLA-SOLUTION	SLEARORS
9	4	637.75	3.9126	-191.6243593	334.7012749
10	10	727 69	15.5423	-172.483-4385	281 7392481
п	36	668.39	34.368	-54.58685765	277.6362111
12	71	692.17	oc.9168	-176 7976161	207 0023693
13	51	792_47	48.4931	-37.3331260	261.5196254
14	54	755 81	51 8728	-81 92905627	201 0555731
15	31	839.32	29 9959	30 83975445	272 3701625
10	18	926 83	17 3541	-64 18783139	286.8697998
17	76	1030 97	25 1631	45 70205881	279 0112375
18	15	895.80	Lala	56 90171 497	284 6747853
10	8	681.50	7 8736	-175 3080.402	300 3517607
10	6	017.05	18 3800	0.5419557-	281 3705533
20	19	246.7J 875.41	17.54%	118 4110308	783 2300-75
21	18	873.44	17.5420	-118.0119798	282.2390073
22	24	1034 42	23 1778	101 8/6029/	280.1128314
23	24	963.42	23.235	-84.36206718	281.0090128
24	12	1050 33	11 7216	32.64741.472	301 2256958
25	9	984.22	8.7904	-122.1978501	308.5368664
20	2	553.00	1.9607	268.0487125	583 9288689
36	1	1245.00	0.9797	59 46333148	486 1482219
t yield					
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTION	SLERBORS
9	4	19.62	3.0119	-4.78672295	11.62430401
10	lo	22.81	14 1958	-5 46150324	9 8320126
11	36	21.29	30.2252	-2.21758370	9 681 69831
12	71	21.00	57 9952	-0 29838604	9 34850956
13	51	25.77	42.9019	-2.06024643	91493856
14	54	23 79	40 1307	-3 13862331	9 16405204
15	31	28.62	27.1948	2.18423904	9 52726863
10	18	32.63	15 9607	-2.06072532	10 01 37 61 42
17	26	33.73	27.585	2 84057353	9 75202191
18	15	29.18	13,4068	1 907538	9 9396867.1
19	8	77 79	7 2035	-3 1400567	10 47406098
20	10	20.18	10 7351	0.05803531	0 23080198
21	18	22 10	16 (13)-4	3 10234137	0 84676805
21	10	26:10	10.0847	-3 10220137	9.84070803
22	24	33.56	20 8997	0.198/0411	- 193-9114
23	24	33.13	21.0307	0.0096412	9 80009414
24	12	36 00	10 7216	3 40941088	10 4986/032
25	9	31.22	8.0673	-2.69796411	10.73844725
20	2	21 01	1 8169	9 5462500-4	20 05168089
27	<u>l</u>	35.45	0.9179	-2.62220599	16.91558782
otein yield					
AGE	NREC	MEAN	DIAGONAL	GLA-SOLUTION	JO LIGHT
10	4	18.50	3.6166	-3.32341329	0.11001437
10	16	21 12	13 0434	-4 25015200	8 328733-44
11	30	2010	32.9572	-9.58368724	8.40347357
12	71	20.47	03 0250	-4 27922-486	8 10.99478
13	SI	24.01	40.5074	-0.3554086	7 92485625
14	54	22.38	19 7954	-1 45712098	7 93046669
15	31	25.57	29 0509	2.50791614	8.2510104
lo	18	27 54	16.965	-0.36653934	8 67974777
17	26	30.73	24.2739	2.80614532	8.44589875
18	15	26 95	14 2094	2 48063334	8 61994035
19	8	21.17	7 6269	-1.55585832	9 06840754
20	19	27.47	17 8097	1 36305458	8 51 432739
21	18	25.53	17 0523	-2 60803944	8.54173164
22	24	3/2 9n	77 .1.17	\$ 35651350	8 48540861
23	24	29.00	77 491.4	-0.61229371	8 50207418
7.1	17	31 %	11 3063	707769744	0 11103411
14	14	31.00	11 3703	1.01000.423	7 11173311
40	*	17.07	6.000.6	+L91999423	y_32776/23
20	2	17.59	1.9188	10 4813359	17.55149018
77	1	274 I W	0.0677	8450(3733	1.4 607375977

#### protein yield (Basic Model 1).

* values of GLS solutions and standard errors for milk are obtained using model 2, because model 1 did not yield in estimate results.

# Table A6: GLS - solutions for age at fist kidding - Nubian - cumulative milk, fat and

Milk yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SLERRORS
10	4	\$95.25	3.1401	-120 1052615	205.5797454
11	20	485.55	15 0486	76 41095554	134.2023422
12	58	508.14	38.5328	94.24617987	120.3776677
13	64	454.05	43.401	43 47458442	119.1200348
14	52	476.58	30.0050	49.31971185	119 100063
15	28	471.89	20.441	78.34586074	127 2282248
10	20	537 20	14.8756	123 2163902	123 0556708
17	17	464.24	11.9666	16 281 53801	146.9993114
18	13	503.31	9 62 42	167 0878245	146 6781865
19	15	575.13	11.20-12	145 1071024	137 2691073
20	20	674.50	14.9041	220 4232811	132.7517816
21	25	435.08	18 5749	21 7179338	121 8455785
22	28	471.86	20.3162	117 423428	113 0029703
23	27	573.90	19.5223	144.0694328	124.657946
24	17	455 88	12 0820	40:05512149	131 4243271
25	9	573.89	6.9017	192 1162400	149 9640843
26	8	496.63	0.1421	187.0434402	150.1045263
27	5	570.00	38108	-43 77476015	214 2703268
29	3	507 07	2.1283	209 3485604	203 482590
Fat vield	····				
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SEPROPS
10	1	78.00	7 5668	-10514536	10.75400976
11	20	27.12	17 1330	4 35752091	7 1/10/308
12	58	2411	29 9604	4.44607150	0 77765913
13	64	22.07	33 7404	2 99915789	6 05735844
14	57	23.40	28.6313	3 78006447	5 0 50 B00 F3
15	28	77.99	16 2023	1 97 40 40 15	7.03716065
io	20	25.09	11 7742	6.75817085	A 807313
17	17	23.05	9 1501	1 07078817	8 02735026
18	13	24 37	7 5211	9 1678796	801555361
19	15	27.05	0.0543	7 10131205	7 515.00178
20	20	30 38	11 8474	6 73 50 1731	7 20313380
71	25	18.61	1.4 73.41	0.49720276	A 78057777
22	78	77.85	16 1092	0.427360039	0.41552036
23	20	27.30	15 7817	8.00-00558	6.91-05330
24	17	21.24	10 0274	2 15962165	7 7.4.158080
25	9	28.46	5 5910	13 1077878	8 11033361
20	8	7418	4 8973	10 10184201	8 30365324
27	š	21.75	3,0307	-5 37610110	11.250.42085
29	1	20.11	1.6257	8.4813.4055	13 85378796
Protein vield			1.0257	6.46134035	13.83328170
	NACO			410 40110001	
10	NREL	MEAN	DIAGONAL	GES-SOLUTION	SLEKKURS
10	4	21.74	3.1330	-7 04087033	7846(1325
17	20	1711	10 1280	U /123540	3.11000809
12	<b>J</b> 6	17.00	35.7841	1.3808U3.59	4.2600 [.39
13	04 (1)	17 07	43 0806	-0 10/5/403	4.3388,5703
14	32	17.80	36 KP3	0.28301220	4.54030962
15	20	1/ 8.3	20.3782	1 06 27 3270	4.849/21/20
10	20	19 /U	14.9013	5.36000287	4 09019045
17	17	16.27	12.04/9	0 29210092	00-189841
10	د،	17 28	7.0525	> 0/334102	5.39243910
20	20	26 74	11.3233	3 28283004	5.23341179
20	1U 1U	24.20 16 CM	14.9891	3 / 3628237	5 00007452
22	20 79	133/	18 0812	+LU/100430	4 04311844
23	20	17.31	20.4345	2.09097900	4.32090244
24	17	41.46 16.37	19 041/	4 14133083	4.75070364
25	0	10.27	12./338	-0.83392082	500935153
20		20.10	0.93/3	3 39233431	5 / 193/241
20	•	10.70	01/0	4.98001409	2.9346320
27	3	ZU 84	1868 6	-2.08418200	8.1793900
	د	18.23	2.1428	0.93052274	10.05736717

#### protein yield (Basic Model 1).

# Table A7: GLS - solutions for breed - All breeds - cumulative milk, fat and protein

yield (Basic Model 1)

Milk yield					<u>.</u>
BREED	NREC	MEAN	DIAGONAL	<b>GLS - SOLUTION</b>	SLERRORS
Alpine	691	740,01	164.83	-373.799	250.8489
Saanen	439	825.68	118.56	-278.0114	254.0316
Toggenburg	641	796.29	124.24	-350.0127	254.9251
Nubian	433	500.61	149.52	-600.0804	252.89
Fat yield					
BREED	NREC	MEAN	DIAGONAL	<b>GLS - SOLUTION</b>	St.ERRORS
Alpine	691	25.74	142.85	-14.96744	9.173017
Saanen	439	<b>26.74</b>	103.65	-12.24453	9.289337
Toggenburg	641	25.46	107.3	-16.03895	9.321726
Nubian	433	23.74	131.89	-16.74097	9.245225
Protein yield					
BREED	NREC	MEAN	DIAGONAL	<b>GLS - SOLUTION</b>	SLERRORS
Alpine	691	22.48	177.14	-12.10844	7.592172
Saanen	439	24.65	126.83	-9.275795	7.688577
Toggenburg	641	22.64	133.81	-12.05526	7.715746
Nubian	433	18.48	159.12	-16.0279	7.655001
## Table A8: GLS - solutions for age at fist kidding - All breeds - cumulative milk, fat

## and protein yield (Basic Model 1).

ilk yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SL ERRORS
9	13	559.85	0 35 7115	0	0
10	131	620.50	35.7115	140.2-31.520	103.10007533
12	101	617.11	103.937	140_3021079	103.0077322
12	293	643.32	214.9620	129 1342407	101.339911
13	303	009.73	225 7473	27 8323484	102.0810757
14	217	0-10.00	167.9759	124.4440885	102.9691289
15	152	764.58	120.4933	169.8494735	103.8779308
10	117	791.84	92.9539	180 0802200	104.3342298
17	119	830.18	91.8389	158.3141177	104.848675
18	86	764 42	69 714	140 8613215	107 1389257
19	76	770.67	61.8537	210.1651496	107.3668934
20	82	783 51	66 8938	719 3012113	107 6762711
71	94	747 43	77.0638	220 2304077	105 9690547
	123	960.00	06 7301	317 3030134	101 51 6393
	125	200 78	76.7371	317 3079124	104 0110363
25	143	788.33	113.3372	239/1114203	104.4394617
24	105	790_32	83 3922	240.4025032	105.5647202
25	50	898.92	40 8675	280.9080583	111 928-1396
20	25	713.36	20.9008	340.2520680	124.9464313
27	ló	748 81	13.4331	141.627033	134 8038670
28	0	804.83	5.0731	238.721936	162.9980251
29	5	702.20	4 0687	321.9942030	180.7336956
30	2	766.00	1 6726	310 8628847	269 8264602
Fat stald					
	NBSC	MEAN	DIACONAL		
9	13	17.19	0	OLS-SOLOTION	0
10	44	71 19	34 4576	1.09831773	41110276
11	111	21.12	100 3400	116797657	1 7673737
11	101	21.27	100.240.5	4.10/8/05/	3 (04333330
12	293	22.61	206.4049	348/4/203	3.68/3/068
13	303	23.76	217.183	3.5417/658	3.70091947
14	217	23.21	161.804	3 95902816	3 73917288
15	152	27.20	116.301	5.07720860	3 77224128
ló	L17	27.30	89 7539	5 00794915	3 78895203
17	119	28.95	88.357	5 68751762	3 80718478
18	80	27.33	67.3703	5 0 1 0 8 2 4 0 7	3.89049607
[9	70	27 49	59 8716	7 39885970	3.89805596
20	87	27.25	61 6408	6 081 55083	3 91022031
21	04	2, 12	14 6620	4 2008 4580	3 8 13060 14
	123	20 70	14.3329	0.3078-087	3 30477,0044
22	123	29 70	45 2793	10 04113191	3 79333233
25	145	21.95	111.4109	9.25316735	3.79242523
24	105	27.15	80.440-4	8 289 433 57	3 83390077
25	50	31.22	39.521	10.00844443	4 06465528
20	25	20 39	20.2239	12.27367613	4.53575303
27	ló	25.29	13.0149	3 37041837	4.89399118
28	0	25.49	49173	7 02010687	5 91393271
29	5	25.11	3.9176	9 95741502	6.57597795
30	2	24.40	1 6772	8 72019153	9 80608316
Protein vield				6,201,133	/ 00000310
	NREC	MEAN	Dia CODE A L		
Q	13			GLA-SOLUTION	
10	4	18 94	36 3407	2 831 47787	3 .4351 3551
11	131	18 84	105 7154	1 73683457	3 14617171
17	303	10.70	105.7154	3.12063451	3 0931 6300
12	273	1979	219 18/3	3 1 3 9 30	3 0631 3299
13	303	20.75	229 9325	3.17359485	3 099/1295
14	217	20.02	170.9002	3 33028305	3 12007482
15	152	23.60	122.5069	4.87926287	3 15423757
lo	117	23.79	94,4901	5.34296322	3 16803117
17	119	25.41	93.5269	4.84233652	3.18383293
18	80	23.41	70.8323	4 42626892	3.253328%
19	76	23.36	62.7993	0.03236752	3 26053289
20	82	23 83	67.9661	0 05131743	3 2695396
21	94	22 21	78 7611	5 60017191	3 7178171
27	123	25 97	100 1077	9 13300707	3 173415.03
 73	دي. ايا	12 04	117 617	7 11170171	3 17344343
20 14	140	23.94	117.517	1.308.59437	3.17137118
24	105	23 84	84 8092	7 17500028	3.20534135
25	50	26.57	41.5098	7.94070541	3.39865976
20	25	22.48	21.2212	10.20255649	3 79463806
27	16	22.72	13.6309	3.58108251	4.09381283
78	Ó	22.92	51465	6.89250461	4 951 470 43
29	5	24.73	4,1407	11 75953103	5 4830237

## IV. GLS - solutions and standard errors for age at first kidding BM2

#### Table A9: GLS - solutions for age at fist kidding - Alpine - cumulative milk, fat and

	NPOO		BIA CONTA I		
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTIONS	SI. ERROR
y 10	, 'a	444.71	4.0457	-61.9812552	264.3391377
10	17	533.24	11.4975	-185 8872128	241.2184201
11	45	630.38	31 3725	-213.8213583	230.1831548
12	116	085.42	71.2899	-223.0562144	222.5263652
13	123	768.83	75.1627	-217.742924	226 6005094
14	55	670.47	35 893	-207 8221231	229 9095706
15	49	751 53	33.0748	-221.8616074	230.3472105
lo	34	8-03.38	23.078	-172.0185182	232.1583685
17	37	828.00	25 0204	-207 616946	231.0304041
18	25	731.12	17.4866	-233 2326285	234.7640456
19	17	731.76	11.9348	-85.49590134	241 1609843
20	Io	722.06	11 1506	-92.08110502	241 5556591
21	19	700 74	13.6347	-44.25577289	238 2931 54
22	35	850.29	24.3211	-79 00307900	232 0227979
23	-58	768 31	12 7821	-90 8984461	230 817999
24	25	780 08	17 837	-87 7.59030.59	715 13698.19
25	ii ii	870.46	9.0647	107 7718877	7.46.6957066
20	.5	1206 50	3 0757	-107 7218827	277 403-381
17	1	1250 50	2,51,51	336 603 404	277 4430361
27	3	997.00	2.1883	233.3974040	287.3178830
28	2	808.50	1.4782	-45 43428-804	308.0013990
	l	743.00	0 7165	-1/0.4512446	380 8609519
t yield				· · · · · · · · · · · · · · · · · · ·	
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTIONS	ST. ERROR
6	7	13.0	51400	-18 51761492	9 65833015
10	17	18 50	12.7007	-10.72430566	8 80227678
11	45	21.27	34.218	-11.7853604	8.39151435
12	110	24.11	78.6224	-12.1171674	8 09849979
13	123	20.35	82.8127	-12.13655007	8 25690094
14	55	23 67	39.3836	-11.42272205	8.38150786
15	49	26 87	36.2238	-11.00110048	8 39564848
lo	34	28 62	25.2734	-10.6328847	8 46237223
17	37	29 51	27 3729	-10 92396754	8 44606778
18	25	27 62	19 0752	-9 \$\$079387	8 Sol 37008
19	17	74 55	13 05 47	-0.85100231	8 80,490905
20	lo	23.47	12 2008	-8 01627729	8 8177867
21	10	22.71	149343	-0.01027729	8 -0799-9-
21	17	20.74	14.8203	1 48009830	8 07/88080
22		30 47	26.3933	-7.03117912	8 4/406302
23	*8	27.02	35.9213	-0 490/0000	8.41529431
24	25	20.20	19 4.985	-0 100-12982	8.58911579
25	13	29 08	9.9089	-8.7080-1697	9 00667896
20	4	45 61	3.2186	10.30796123	10 16998853
27	3	33 29	2.3823	3.36837662	10.55213614
28	2	28.70	1.599	-0.45190486	11.34430489
29	1	29.78	0 7862	-9 09010041	13 988 49673
Protein yield					
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTIONS	ST. ERROR
9	7	12.34	5.1015	-14.06333225	8 02519839
10	17	16.10	12.5923	-7 42822585	7 31469518
11	45	18 75	33.9638	-8.85233125	o.97392118
12	llo	20 86	77 952	-8.98974735	6 731 46185
13	123	23.33	82.1142	-8.71497386	6 86234967
14	55	20 35	39 0669	-8 04213771	6 96 56 7 1 18
15	49	23 54	35 9.11 1	-8 32089487	6 9775733
10	34	25 82	75 0750		7 (1170.125.)
17	37	26 71	77 1607	-7 05830017	7 0101 4070
18	2.	72 10	190333		7 1 1 49000 4
10	17	10 ED	10 7334	5 30000000	7 11467074
10	17	21.33	12.9545	• 7 JUUE9837	/ 51051342
20	10	21.52	12 1069	-4.98/02983	/_32693432
21	19	20.45	14.7203	-4.05635163	7 227756-44
22	35	25 67	26 391 9	~4 46844855	7 0435289
23	48	22.98	35.6395	-4 66724881	6.99364764
24	25	23 76	19 2961	-3 43260756	7 13771215
25	13	25.56	9 8333	-6.1162289	7 48423839
26	1	39 48	3 1971	10 1491612	8 44782775
27	3	28.93	2.3652	3.55669314	8 76375427
	•	36.00	1 (197)	-3 3035574	0.410.41604
28	4	2300	1 366.3		7 91 74 8793

## protein yield (Basic Model 2)

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## Table A10: GLS - solutions for age at fist kidding - Toggenburg - cumulative milk,

Milk yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTIONS	ST. ERBOR
9	2	807.00	1.6853	-311.6846997	428.6998724
10	7	639.57	5 5323	-101.59-13068	383.3615281
11	30	632.10	23.7513	-105 1188357	362.4880319
12	48	633.85	34.8652	-36.99047433	357.7302825
13	05	598.28	48.7934	-1.49.0928848	355 3108252
14	50	675.95	42,6748	-107 681022	354.4786436
15	-1-1	912 70	33 9001	-98 11557715	350 1827552
10	15	812.07	21 (79)	-130.007#707	355 1728705
10	10	812.07	29.9771	-120,3978702	167.0-14010
17	39	837 93	28.2514	-145.81/31.32	337 0014223
18	33	832.33	20.0627	-219 8999287	358.7652491
19	30	890 33	28.0145	-06.19069385	357 o121982
20	27	788.48	21.9331	-lol 01 <b>882</b> 51	359.283163
21	32	932.47	25.7382	39.54502555	357 8523709
22	30	1024.25	28.5995	153.269577	357.199099
23	40	844.33	35 8293	o5 8o124908	345 6336915
24	51	857.71	38.3658	5.17119403	356.0859474
25	19	1031 95	15 1905	67 87441176	362 235270
20	11	688.00	8 0844	7 21 22 20 20 5	37.4 3303070
10		068 07	8.38**	1/11/22/303	374.3343723
27	5	/0/ 13	0.0424	-108./32.9-79	3/3 4311019
28	4	773.00	3.3358	-48.21035389	408.4759554
29	2	766 00	1 6504	88 96317008	450.3398412
Fat yield					
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTIONS	ST. ERRON
9	2	25.48	1.6747	-11.11202903	13.95839038
10	7	19.72	5-1896	-5 04138904	12.48340131
11	30	20 cici	23 5732	-5 23010R12	11.80517403
17	.08	19.57	11 5077	-1 18791006	11 53979393
13		19.05	69 4197	7 139 (407-	11 57110-00
15	6	18.93	46.4187	-7 128-4078	11.57110009
14	00	22.00	42.3408	-4 8221 3585	11 04330070
15	44	29.24	33.6497	-524832939	11.59927833
10	45	25.14	34 2188	-7 45278619	11 500-1998
17	39	27 63	28.0079	-6 72813745	11.62830635
18	33	27 45	25 8734	-8.78211658	11.68350252
19	36	30.21	28.4268	-3 37804422	11 64588083
20	27	2583	21 7821	-6 53-533761	11 70075702
21	12	20 51	15 5508	96722108 0	11.653.0003
**	14	25 33	20.0078	0 763010 43	11 + 11293
		31.15	28.3803	0 73301042	(1 3238370
25	40	20 59	35 55	-0 21102309	11 2002 49/1
24	51	27 47	38 0553	-1 07375454	11 59613423
25	19	34 00	15.0841	1.70545333	11.79007100
20	11	21 99	8.9201	-1.39857084	12 18986839
27	8	24 51	0.5997	-8.30245407	12.22552777
28	4	23 88	3.3132	-3 63325003	13 30162913
29	2	24.40	1 6 3 9 6	-0 9R066685	14 861 50636
notein stald					
	NREO				
<u>AUE</u>	3	11.60	DIAGONAL	GLS-SULUTIONS	31. LKRUR
10	7	10.00	1.0634	100033801	14.14/10406
	ý	19,30	5 5437		10 04240023
11	<b>SU</b>	18.62	23.7152	-4 12050308	10 25513175
12	-18	17 0-1	34 8048	-3 23242752	10 12036431
13	05	16.91	48.7174	-5.9626461.5	10.05202402
14	50	19 48	42.6082	-4 73421597	10 02836853
15	44	25.96	33 8542	-4 19857607	10.07665258
io	45	22.57	34 4204	-5 [968ci993	[0 04812375
17	39	24 21	28 202	-5 87841.4Q1	10 10158.00
18	11	13 AA	76 07 11	.7 4517 4712	10 1.077.09
10		23 00	20 0244	1 361000.	10 11 71 01 00
17	30	10 /6	28.5760	-3_33(9/96	1011/10172
20	1/	22.44	21 9026	-0.29503268	10 16445515
21	32	26.58	25.7021	-0.84968423	10.12384841
22	30	29 32	28 5564	2.39831597	10 10540682
23	46	23.75	35.7728	0.37706169	9.77830116
24	51	24 58	38.3029	-1 13062973	10 07391558
25	19	29.12	15 160	0.49717708	10 2478285
20	11	10.74	20714	1 1541140	10 59000944
	0	17 /3	5.7/14	-1.524/2007	10 (01777433
21	5	21.57	0.0338	-0.24121268	10.02155221
28	4	21.88	3.3312	-2.37520351	L1.5559581
29	2	20.93	16482	0.67500988	12.91026848

## fat and protein yield (Basic Model 2).

## Table A11: GLS - solutions for age at fist kidding - Saanen - cumulative milk*, fat

# and protein yield (Basic Model 2).

<b>Lilk yield</b>					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SL ERRORS
9	4	637.75	3 880	-160.0833005	372,9799019
10	10	727.69	15.4139	-151.2908533	323.0575719
11	30	008.39	33 9501	-00 4783-4972	306 5343156
12	71	692.17	65.8906	-168.0520666	304.5238606
13	51	792 47	47 8562	-37 54138785	300.343294
14	54	755.81	51.3098	-00.00464515	301 3446187
15	31	839_32	29 7223	34.97574906	309 9055441
to	18	926 83	17.1793	-80 60845059	319 6323591
17	76	1030 92	24 92 47	1975901849	314 6219084
18	15	896 80	14 5073	68 50468674	373 3716175
10	15	681.50	7 7713	-107 8553063	145 3638307
19		0.01.00	18 2021	4 13-00337	310 0501 747
20	19	742.93	18.2021	4.13000227	317 0301/4/
21	18	8/3 44	174124	-11/ 3900302	317.7190090
22	24	1034.42	22.9409	101.7354561	317.5941259
23	24	903 42	23 0139	-88.74824768	319 5350977
24	12	1050.33	11.637	31.95328076	337.5201918
25	9	984.22	8 7351	-108.8056678	340 9243224
20	2	553.00	1.9484	276.5150718	599.6846001
29	I	1245 00	0 9732	109 8960994	519 7762999
ıt vield	· · · · · · · · · · · · · · · · · · ·				
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SL ERRORS
9	4	19 62	3 544	-7 07577368	12.91032227
10	10	77.81	13 9109	-8 44990741	11 22934799
11	10	21 29	79 JOR	-5 26285036	10 66 590 808
17		71.00	56.11.56	-0.011738.17	10 00305003
12	<i>.</i>	35.77	41 8603	514738707	10.1.036710
13	31	23.77	41.8003	*J.14736474	10 4031003
14	54	23.79	45 0437	-3 /0220803	10 4731073
15	31	28.02	20 0101	-0 /0932458	10./8/8/923
10	18	32.03	15 6531	-5 11556252	11 102/45/8
[7	26	33.73	22.059	-0.33358007	10 94334437
18	15	29 18	13 1487	-0 671 4084	11.24046876
19	8	22.29	7.0634	-5.78808251	11.975903
20	19	29-18	lo 3777	-2.98699274	11 0723575
21	18	28.10	15.774	-0.73454398	11 04915892
22	24	34 38	20 4203	3 42099391	11 03460075
23	24	33.15	20.5730	-3 18625706	11.06864296
24	17	30.00	10 4997	0 33736321	11.72257076
25	9	31 72	7 9023	-5 29R52846	11 83486781
20	2	21.01	1 7821	R 12180831	20 56350-888
20		35.45	0.9076	.3 95619873	18 00465075
		33.45	4.5440	-3.73017673	13.00-0030-5
ouen yren					A. 50000
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTION	SL LIGHUNS
9	4	18.56	3.6026	-3 86032849	11.32010117
10	10	21.12	14.1509	-5 12319912	983827408
11	30	20.10	30 1124	-2.10379776	9.34384207
12	71	20 47	57 7749	-5 271 47300	9 287660L
13	51	24.01	42.7575	-1 70238312	9.15380568
14	54	22.38	45 9859	-2.28061815	9 19002775
15	31	25.57	27.1155	1.24503161	9 44946337
to	18	27 54	15 9238	-1 93034471	9 72009680
17	20	30.73	22.5129	0.90241045	9.58600076
18	15	20.95	13.3716	1.30623396	9 8-18-1803
19	R	21 17	7 1845	-2.49481713	10 49593849
20	10	27.47	16 6265	0.0817.4681	9 707 5970
20	17	21 97	10 0000	3 77060717	0.6700.41.43
21	18	20.00	10.0423	-3.//00721/	0.07774143
22	24	34/90	20 8.54	4 13/03/08	700/0//13
23	24	29.00	20.9681	-1.96298738	9.71713123
24	12	31.50	10 0915	1 8368603	10 271 19066
25	9	29 09	8 0449	-2.50040563	10.37123369
20	2	17 59	1 8122	9 99400815	18.0509649
27	1	43.18	0.9158	9.22618735	15.78395342

* values of GLS solutions and standard errors for milk are obtained using model 2, because model 1 did not yield in estimates.

# Table A12: GLS - solutions for age at fist kidding - Nubian - cumulative milk, fat

## and protein yield (Basic Model 2).

Milk yield			-		
AGE	NREC	MEAN	DEAGONAL	GLS-SOLUTION	SLERRORS
10	4	595.25	3.0133	-150.014924	233.654053
11	20	485.55	14.396	45 21 1997 57	172 9114376
12	58	508.14	36.5208	58.94827095	157 1638844
13	64	454.05	41.1509	9 00808908	154 7520419
14	52	476.58	34.8192	17 85259782	157 0593474
15	28	471.89	19 4899	44 47184112	165 650001
10	20	537 20	14.1772	91.63980707	157 8315571
17	17	464.24	11.3112	-18 29851952	181.8179742
18	13	503.31	9 1 473	115.9760754	182.8288572
10	15	575 13	10 76-48	108 1949182	172.7632053
20	20	674.50	14 21 18	186.9190549	100 4082175
21	25	435 08	17 7084	-3 00267317	159 6851364
27	78	471.80	19.3566	84 251 39687	152.9388747
11	20	\$71.95	18 554	110 2031 391	157 4509158
25	17	455 99	12 08.48	8 54018304	109 0001 534
24		433.00	6 6101	190 0750570	18.4 0770087
25	ý	573.87	5 9-3-	174 0018933	187 2719158
20	8	490.03	3.8636	174.00167.55	210 1067736
27	5	570 60	3.0420	-39 4830033	237 4702/30
29	3	507 67	20110	140.1//802/	482.332/831
Fat yield					·······
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SLERBORS
10	4	28.00	2.4817	-6.54046045	12.19471665
11	20	22.12	11.7069	1 70782851	9 27513195
12	58	24.11	28.7771	1.72801179	8 541 40889
13	64	22.07	32.3985	0 3301 0017	8 41103158
14	52	23 40	27.4986	1.28267754	8 51601928
15	28	22.99	15 0045	2.31577197	8.93571085
10	20	25.09	11.3255	4.37561587	8 5259395
17	17	23.45	8 7616	-0 46045953	9 73705502
19	13	24 37	7 7207	5 596-40617	9 82227212
18	13	24.37	8 73.65	4.4723705	9 2891 5800
19	13	20 38	11.4083	6 61347787	8 9761 2345
20	20	30.36	14 1203	1 00338070	80000113
21	25	1861	14 1793	3 99133-0	8 1 4007948
22	28	22.85	13.30%6	3.8813200	9 6 19- 107
23	27	27 30	14.6798	6.004.5171	8 3480-8//
24	17	21.24	9 6-137	-0 4386556	911972284
25	9	28 40	5 3987	11 09107440	982009793
20	8	24.18	4 7093	8.70201911	95824732
27	5	21.75	2.9265	-7 0100007	12.51979329
29	3	20 1 1	1.5563	4 08044058	14.78500139
rotein vield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SLERRORS
10	4	21.71	3.0308	-7 13082694	8.91538c01
11	20	17.11	14.480	0.84099863	o 59141441
12	58	18 00	36 7941	1,28814369	598823198
13	Â	17.07	41 4575	-0 21707505	5 89633556
14		17 20	15 0719	0.39655841	5 98480385
14	24	17 87	19 6704	1 730.40888	0 3133129
15	48	10.20	14 2733	3.45703413	00151120
10	20	19 70	11 4007	0 10 10 00 00	691140124
17	17	18.27	11-00/	0.303000.36	6 0=0000044
18	13	19 28	9.2129	1.44//2317	0.70702200
19	15	20.83	10.8335	3 00307839	0.36511987
20	20	24.26	14.307	5 676822-19	0.34401595
21	25	15 57	17 8277	-0 83606799	6 08 477977
22	28	1731	19 4882	2.76199540	5 82625301
23	27	21 48	18 6869	4 38703801	5 99942649
24	17	lo.27	12.1671	-0.7649315	6 44113093
25	9	20.1o	6 6 503	o 50212738	7 0182947
20	8	18.96	5.9021	5 09971282	7.14115031
27	5	20 84	3 0007	-2.09772001	9 13793896
70	1	18.25	2 0276	5 10034403	10 77964873

## Table A13: GLS - solutions for breed - All breeds - cumulative milk, fat and protein

Milk yield					
BREED	NREC	MEAN	DIAGONAL	GLS - SOLUTION	St.ERRORS
Alpine	691	740.01	0	0	0
Saanen	439	825.68	110.6955	86.105529	52.36280324
Toggenburg	641	796.29	115.257	17.90404726	59.57778728
Nubian	433	500.61	140.2825	-236.1517426	50.42568071
Fat yield					
BREED	NREC	MEAN	DIAGONAL	GLS - SOLUTION	St.ERRORS
Alpine	691	25.74	0	0	0
Saanen	439	26.74	<b>97.88</b> 3	2.56282649	1.95832369
Toggenburg	641	25.46	100.861	-1.18516168	2.2203795
Nubian	433	23.74	124.9537	-1.82685791	1.88578661
Protein yield					
BREED	NREC	MEAN	DIAGONAL	GLS - SOLUTION	St.ERRORS
Alpine	691	22.48	0	0	0
Saanen	439	24.65	118.221	2.61136517	1.56598083
Toggenburg	641	22.64	123.8515	-0.08740305	1.78521144
Nubian	433	18.48	149.1345	-4.16368184	1.50804139

## yield (Basic Model 2).

## Table A14: GLS - solutions for age at fist kidding - All breeds - cumulative milk, fat

## and protein yield (Basic Model 2).

filk yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SLEERORS
9	13	559.85	0	0	0
10	+4	020.50	35 0812	108.7309431	113.3493898
11	131	019.11	102.0755	147.5725565	103.8166543
12	193	043.32	210.0245	134.0908347	101 73/6267
13	303	009.75	221.4130	133.5370406	102.2844494
14	217	040 00 7-: ( 69	104 8009	133 0992033	103 1970395
15	132	701.94	01.3632	177 6396401	104 0414641
10	117	771.84	91.3400	193.2473943	104.248479
19	80	761.07		107.5275703	107 2040100
10	30 7a	770 47	0.9587	147.5430802	107 549109
70	70 97	793 51	00 0.382 a \$ 76\$1	210 8031310	107 2061129
20	04	763 51	75 9009	227 17/0713	107.9033091
21	74	/42.43 950.09	13 8028 (h) 908h	22/ 8/42303	100 1793301
** 77	145	6.JU.70 7897 53	113 4608	323.1241963	104 /111348
20	140	706.33	91.0066	1 40 2347939	104.3799901
24	(U)	790.32 908.00	81 9000	247.2207838	105 7704821
25	30	676.72	40 1912 20 Sela	280 41/0/23	120944742
28	15	713.30	20 3010	334.8100012	125 3509054
27	10	/46.61	13.22.90	130.0180070	135 101308/
28	0	804 83	4.9931	241.30906	103 1203-198
29	,	102.20	3.5528	347.3871902	181.8080634
		/00 00	[ 64/3	330,4391041	2/1.09//3/8
Pat yield					
AGE	NREC	MEAN	DIAGONAL	GLS-SOLUTION	SLERBORS
9	13	17.19	0	0	0
10	44	21.19	33 8994	3.63054296	4 11805999
11	[3]	21.27	98.0045	4.49914418	3 76985178
12	293	22.01	202.683	3.70260071	3 69 476231
13	303	23.76	213.4393	3 8514002	3.71-632283
14	217	23 21	159 0871	4.32130217	3.74751565
15	152	27 20	114-4417	5 989386	3 778212
16	117	27.30	88.334	5 85904403	3 79670511
17	119	28.95	80 8264	6 04310162	3 81 477809
18	86	27.33	66.3252	5 93513204	3 89988293
19	70	27 49	58.9871	7 0808-191	3 90540464
20	82	27.25	63 6452	7 24504695	3 91850-131
21	94	25 18	73 4323	0.01720785	3.85558829
22	123	29.70	93.743	10.37039108	3 80261109
23	145	27 95	109 5803	9 47378374	3 79762100
24	105	27.15	79.1311	8.6081592	3 84126337
25	50	31.22	38 9204	10 20940273	4 07077563
26	25	20.39	19.9201	12.70022944	4.55054291
27	16	25 29	12.8269	3 87523-884	4.9047283
28	0	25.49	4 8472	7.59335	5 9189808
29	5	25 11	3 8502	10 80951234	o ó1379342
30	2	24.40	1.5997	9 689078	9 85139041
Protein yield					
AGE	NREC	MEAN	DIAGONAL	GLS - SOLUTION	SLERRORS
9	13	15.68	0	0	0
10	-14	18.94	35 686	2.93267864	3 44228266
11	131	18.84	103.8615	3.9293164	3 15356893
12	293	19 79	214 7854	3 28376420	3 09023289
13	303	20.75	225 5706	3.38081261	3 10697669
14	217	20.02	167 8493	3 55969729	3 13467205
15	152	23 00	120.4078	5.10246492	3 16030441
lo	117	23 79	92.8887	5 55220479	3.17566429
17	119	25.41	91.7674	5 093-43039	3 19126509
18	80	23 41	1000 90	4 651 13 665	3 26221345
19	76	23.36	61.8135	6 221 562 48	3 26780162
20	82	23 83	00 8482	6.28298067	3 27768586
21	94	22.21	77.0128	5 81 43691	3 22536068
22	123	25 87	98 0080	9.35656868	3 18061302
23	145	23.94	115.4531	7.71543791	3 17672184
24	105	23.84	83.332	7 41892842	3 21271601
25	50	26.57	40.8401	8.10085621	3 40486046
20	25	22.48	20 8872	10.59563896	3,80826237
27	16	22.72	13.4246	3 99208342	4 10436501
28	6	22.92	5 0699	6.98501367	4.957097.17
29	5	24 71	40656	12 49366801	5 51774581
30	2	20 91	1.6716	9 44670004	8 211 40824
**	-	av 73		/	وتعدرابه

	SOLUTION	S.ERROR	SOLUTION	S.ERROR	SOLUTION	S.ERROR
ALPINE	Milk		Fat		Protein	
Sire						
1	-1602.093072	2506.879034	-41.16992118	90 8057028-1	-60.99999799	75 51367336
2	-935.1360739	2315.296949	-25.53587055	83.7938149	-44.62911917	09.68885095
3	-1103 023638	2202.385294	-17.86803235	79 807023	-38 94172454	00 36526876
4	-902.8376413	1819.399941	-28.71312126	65.62991219	-24.80064851	54.59890362
Dam						
1	-1008 9777 52	2155.708152	-14.08188356	78 0472261c	-8 272808534	o4.90579954
2	-1898.864016	1967.219027	-12.61388948	71.13282421	-32.42439023	59 16210309
3	-1223 514111	1840.490634	-32.44457875	00 02417357	-23 85008172	55 4065179
4	-1128.444935	1524.639344	-22.70524711	54.92718616	-28.96422607	45.70028557
TOGGENBURG	Milk		Fat		Protein	
Sire						
1	1709.799224	2685 010907	47 72472249	87 63078845	38 83305120	75 99457193
2	1334.204675	2360.138268	31.83506671	77.19444011	33.95156975	66.96-1262-48
3	1407 593547	2260.270674	38.18691873	73 72423411	40 43881441	03 96521965
4	111.4322655	1881.973395	20.77796209	61.39153039	19 19100447	53 26059684
Dam						
1	-715 6802882	22.47 286-133	11.46415101	73 300-48-409	13 03852804	03 00930524
2	-784.9204342	1826 440113	10.26131331	59.59582518	-1.440212032	51.69176299
3	-23 50374905	1773.00334	24 45859744	57 800796	13 77130174	50 19445254
4	1504 831934	1799.387575	60.18898327	58.69726118	49 2128124	50.92333754
SAANEN	<u>Milk*</u>		Fat		Protein	
Sire		_				
1	524.6190275	3263 744931	-36.80054714	115 193069	-8.008273888	100 5905034
2	159 7959565	2729.359486	-49 09035783	96 57310604	-11.03503544	84.28537429
3	-071 2082011	2539 652034	-80.80362014	89 88335995	-38 74239077	78.47918671
4	-250.7677239	2279 604735	-41.54442495	81.04780349	-26.99154767	70.72246465
Dam						
1	-672.3801494	2801.045039	-25 30228358	100 0520451	-31.51849157	87 20093278
2	-606.54452	2430.252938	-16.39247049	86.49459318	-33.64081123	75.40331317
3	100.8302978	232/ 243/78	2.200789949	82.54632379	-13.33224/62	12.21390872
4	-367.3930315	2040 78377	-20.03284283	/302924326	-23.9349837	03.06078193
NUBIAN	Milk		Fat		Protein	
Sire						
1	610 8402887	3235.161708	53 21143114	172.3504045	35 77820072	123.3503438
2	-327 60[11]	2703.633834	1.307011673	142.8683736	-17 6080816	103 1179356
3	-400.9499221	2340 808144	-23.87237014	123 6235888	-21 04381487	89 281800-48
4	-209 7564226	1791 113187	-18.02793061	94.29281347	+14.15720094	08.32395857
Dam						
1	-022.070035	1832.20317	-00 00344404	98 9541813	-47 20338233	09 82.987914
2	115.6/06275	1449.778372	-24.03684186	77.22009528	-9 29 3968 340	55.28001506
3	098.9050377	1302.351005	27.37591382	72.06204049	15 12322006	51.95971067
4	85 //920409	1023 103631	-0.174030068	828181754	-8.763709815	01.90202964
ALL BREEDS	Milk		Fat		Protein	
Sire	MA 140	1944 001 004	6 6 . D		12 600-0-0	37
1	490 3990334	1240 091435	2010/06/06	45 8113/414	12.9592568	3/0113343
Z	031 3841334	1102.548028	-2.552332555	40.47983426	10 10081004	33_3U288U39
3	636 7813238	995.8321011	3.297007884	36 52220975	19 64948472	30 09702336
4	352.3785303	//1.1481401	-0.513637582	28.51062322	12.52595621	23.29408896
Dam	110	M30 70 4 73 m	16 19-0 90	24.000.0000	0.00000000	30 30 400. 4
I	112.0028674	938./U44249	15.17090079	34 0000 5229	4 20/002376	28.29436803
2	-108.0403338	/51.5016205	8 107310340	27.67939787	1 878906008	22.00112706
3	112.7823534	047.9724379	8 280917502	23 86053873	7 045847755	19.54183-06
4	506 1819854	688 9707307	10.13090414	25.32848450	16.43769912	20.79674857

## Table A15. GLS - solutions for group coefficients by trait and breed

* Estimates for age effects on milk are taken from model 2 of basic model 1 because model 1 did not yield in estimates.

		<b>Basic Model</b>	1		Basic Model 2	2
Milk yield	Model 1	Model 2 (H _A )	Model 2 (He) $\sigma^2_{N}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed
σ²A	26329.41	19135.87	26329.32	26934.65	15317.57	26934.66
σ² _M	-	6831.1	0	-	10002.56	0
$\sigma_{E}^{2}$	60521.12	61128.84	60521.19	60124.84	60689.72	60124.83
σ² _P	86850.53	87095.81	86850.51	87059.49	86009.85	87059.50
b²	0.3032	0.2197	0.3032	0.3094	0.1781	0.3094
<b>m</b> ²	-	0.0784	0	-	0.1163	0
Log L	-4375.160328	-5466.766474	-5467.216161	-4167.172493	-5410.504087	-5411.382603
$-2[H_0 - H_A]$			0.8994			1.7572
Fat yield	Model 1	Model 2 (H _A )	Model 2 (He) ^{σ²_M} fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed
σ²A	21.51	15.79	21.52	27.24	19.86	27.24
σ² _M	-	5.58	0	•	7.40	0
$\sigma_{E}^{2}$	93.41	92.8	93.41	90.17	89.81	90.17
$\sigma_{P}^{2}$	114.92	114.16	114.92	117.41	117.08	117.41
h²	0.1872	0.1383	0.1872	0.2320	0.1697	0.2320
m²	-	0.0488	0	-	0.0632	0
Log L	-2672.44966	-3764.330634	-3764.505499	-2493.673349	-3737.612126	-3737.883459
$-2[H_0 - H_A]$			0.3496			0.5426
Protein yield	Model 1	Model 2 (H _A )	Model 2 (H ₀ ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{M}$ fixed
σ²A	16.98	10.1	16.98	19.34	8.89	19.34
σ² _M	-	6.9	0	-	10.10	0
$\sigma_{E}^{2}$	63.15	62.64	63.15	61.64	61.92	61.64
σ²p	80.13	79.64	80.13	80.98	80.92	80.98
h ²	0.2119	0.1268	0.2119	0.2389	0.1099	0.2389
m ²	•	0.0887	0	-	0.1249	0
Log L	-2566.942844	-3669.432024	-3669.998677	-2398.888016	-3642.029984	-3643.098125
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			1.1332			2.1364

## Table A16. Estimates for Model 2, $\sigma^2_M$ set to zero. Significance test – Alpine

$\chi^2_{1,0005} = 7.88$		[from Steel and Torrie, 1980]
$\chi^2_{1,0050} = 3.84$	$\chi^2_{1,0.025} = 5.02$	$\chi^2_{1,0010} = 6.63$
$\chi^2_{1,0500} = 0.455$	$\chi^2_{10250} = 1.32$	$\chi^2_{1,0100} = 2.71$
$\chi^2_{1,0.950} = 0.00393$	$\chi^2_{1,0.900} = 0.0158$	$\chi^2_{1,0.750} = 0.102$
$\chi^2_{1,0.995} = 0.0000393$	$\chi^2_{1,0990} = 0.000157$	$\chi^2_{1,0.975} = 0.000982$

		<b>Basic Model</b>	1		<b>Basic Model</b>	2
Milk	Model 1	Model 2 (H _A )	Model 2 (H ₀ ) $\sigma^2_{\rm N}$ fixed	Model 1	Model 2	Model 2 σ ² _M fixed
$\sigma_{A}^{2}$	20323.99	18401.25	20323.98	22268.37	22126.35	22268.31
$\sigma^2_{M}$	-	[144.79	0	-	9.13	0
$\sigma_{E}^{2}$	98765.89	99017.99	98765.91	96604.73	96704.34	96604.78
σ² _P	119089.89	118564.03	119089.89	118873.10	118839.82	118873.09
h²	0.1707	0.1552	0.1707	0.1873	0.1862	0.1873
m²	-	0.0097	0	-	0.0001	0
Log L	-3939.714968	-4704.760849	-4704.768404	-3797.473109	-4652.934982	-4652.934864
$-2[H_0 - H_A]$			0.0152			0.0002
Fat	Model 1	Model 2 (HA)	Model 2 (He) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed
$\sigma^{2}_{A}$	25.53	7.59	25.53	24.49	10.79	24.49
$\sigma^{2}{}_{M}$	-	13.59	0	-	10.62	0
$\sigma_{E}^{2}$	101. <b>48</b>	104.96	101.48	101.67	103.74	101.67
σ² _P	127.01	126.15	127.01	1 <b>2</b> 6.16	125.16	126.16
h²	0.2010	0.0602	0.201	0.1941	0.0862	0.1941
m²	-	0.1078	o	-	0.0848	0
Log L	-2271.514742	-3035.649869	-3036.568178	-2153.460898	-3008.412123	-3008.922656
$-2[\mathbf{H}_0 - \mathbf{H}_A]$		<u>.                                    </u>	1.8366			1.021
Protein	Model 1	Model 2 (H _A )	Model 2 (H _e ) σ ² _M fixed	Model 1	Model 2	Model 2 <u> </u>
$\sigma_{A}^{2}$	15.78	10.66	15.78	17.96	17.92	17.96
$\sigma^{2}_{M}$	-	3.14	0	-	0.061	0
$\sigma_{E}^{2}$	79.16	80.07	79.16	77.20	76.97	77.20
$\sigma_{P}^{2}$	94.94	93.88	94.94	95.16	94.95	95.16
h²	0.1663	0.1135	0.1663	0.1887	0.1887	0.1887
m²	•	0.0335	о	-	0.0006	0
Log L	-2202.757769	-2967.734426	-2967.811205	-2086.115286	-2941.577182	-2941.577044
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			0.1536			0.0002
χ ² 1,0 995 χ ² 1,0 950	= 0.0000393 = 0.00393	χ ² χ ²	1.0990 = 0.0001 1.0900 = 0.0158	57	$\chi^2_{1,0.975} = 0$ $\chi^2_{1,0.750} = 0$	.000982
X ² 1.0 500	= 0.455 = 3 84	x ²	$1 \circ 250 = 1.32$		$\chi^2_{1,0100} = 2$	.71

 $\chi^2$  1.0.005 = 7.88

Table A17. Estimates for Model 2,  $\sigma^2_M$  set to zero. Significance test – Toggenburg

 $\chi^2_{1,0\,010} = 6.63$ 

[from Steel and Torrie, 1980]

	Basic Model 1				Basic Model 2			
Milk	Model 1	Model 2 (HA)	Model 2 (H _e ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed		
σ²A	۴	0.93	r.	f	1.80	f		
$\sigma^2_{M}$	-	2307.97	f	•	5861.61	f		
$\sigma_{E}^{2}$	ſ	129676.75	ſ	f	129571.76	ſ		
$\sigma_{P}^{2}$	ſ	131985.65	ſ	f	135435.16	f		
h²	ſ	0.0000070	ſ	ſ	0.000013	ſ		
m²	-	0.0174	f	-	0.0433	ſ		
Log L	f	-3194.788771	f	ſ	-3139.706979	f		
$-2[H_0 - H_A]$			ſ			ſ		
Fat	Model 1	Model 2 (H _A )	Model 2 (H ₀ ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed		
$\sigma_{A}^{2}$	14.89	0.0065	14.89	18.28	0.04	18.28		
$\sigma^2_{M}$	-	13.96	0	-	16.21	ο		
σ²ε	143.6	143.91	143.6	143.75	145.05	143.75		
σ² _P	158.49	157.87	158.49	162.03	161.30	162.03		
h ²	0.0939	0.00004	0.0939	0.1128	0.0002	0.1128		
m²	-	0.0884	0	-	0.1005	0		
Log L	-1589.723288	-2267.572374	-2267.8223	-1561.013863	-2238.874601	-2239.112875		
$-2[H_0 - H_A]$			0.5			0.4764		
Protein	Model 1	Model 2 (H _A )	Modei 2 (H ₀ ) $\sigma^2_{\rm M}$ fixed	Modei 1	Model 2	Model 2 σ ² _M fixed		
$\sigma_{A}^{2}$	4.85	0.0021	4.85	12.01	0.0063	12.01		
$\sigma^2_{M}$	-	8.02	0	-	10.08	0		
$\sigma_{E}^{2}$	115.28	112.47	115.28	112.43	113.57	112.43		
$\sigma_{P}^{2}$	120.13	120.5	120.13	124.43	123.65	124.43		
h ²	0.0404	1.77167E-05	0.0404	0.0965	0.00005	0.0965		
m²	-	0.0666	0	-	0.0815	0		
Log L	-1553.210892	-2231.103234	-2231.309903	-1526.308188	-2204.259897	-2004.407200		
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			0.4134			0.2948		

Table A18. Estimates for Model 2,  $\sigma^2_M$  set to zero. Significance test – Saanen

## * = estimation failed

$\chi^2_{1,0005} = 7.88$		[from Steel and Torrie, 1980]
$\chi^2_{1,0050} = 3.84$	$\chi^2_{1,0.025} = 5.02$	$\chi^2_{1,0.010} = 6.63$
$\chi^2_{1.0500} = 0.455$	$\chi^2_{10250} = 1.32$	$\chi^2$ 1.0 100 = 2.71
$\chi^2_{1,0950} = 0.00393$	$\chi^2_{1,0.900} = 0.0158$	$\chi^2_{1,0,750} = 0.102$
$\chi^2_{1,0.995} = 0.0000393$	$\chi^2_{1,0990} = 0.000157$	$\chi^2_{1,0.975} = 0.000982$

	Basic Model 1			Basic Model 2			
Milk	Model 1	Model 2 (H _A )	Model 2 (H _e ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed	
σ²A	I <b>4728.3</b> 7	2.64	14727.83	17630.38	0.51	17630.22	
$\sigma^2_M$	-	16991.28	0	-	17889.33	0	
$\sigma_{E}^{2}$	42825.46	41221.79	42825.87	41786.29	41677.47	41786.41	
σ² _P	57553.84	58215.72	57553.71	59416.67	59567.31	59416.63	
h²	0.2559	0.000045	0.2559	0.2967	0.000009	0.2967	
m²	-	0.2919	0	-	0.3003	0	
Log L	-2405.73259	-3193.83786	-3195.972899	-2351.909154	-3139.884979	-3142.149467	
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			4.27			4.529	
Fat	Model 1	Model 2 (H _A )	Model 2 (He) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	$\frac{\text{Model 2}}{\sigma^2_{\text{M}} \text{ fixed}}$	
$\sigma^2_A$	69.46	19.9	69.46	75.72	18.22	75.72	
$\sigma^2_{M}$	-	46.86	0	-	51.05	0	
$\sigma^2_E$	89.36	91.84	89.36	87.75	94.46	87.75	
σ² _P	158.82	158.6	158.82	163.47	163.74	163.47	
h²	0.4373	0.1255	0.4373	0.4632	0.1113	0.4632	
m ²	-	0.2955	0	-	0.3118	0	
Log L	-1638.97518	-2426.99957	-2429.21549	-1608.597984	-2396.421798	-2398.838296	
$-2[\mathbf{H}_0 - \mathbf{H}_{\mathbf{A}}]$			4.4218			4.8344	
Protein	Model 1	Model 2 (H _A )	$\begin{array}{c} \text{Model 2 (H_{0})} \\ \sigma^{2}_{M} \text{ fixed} \end{array}$	Model 1	Model 2	Model 2 $\sigma^2_{M}$ fixed	
$\sigma^2_A$	21.05	0.0001094	21.05	25.18	0.0035	25.18	
$\sigma^2_{M}$	-	25.6	0	-	26.93	0	
$\sigma_{E}^{2}$	62.83	59.24	62.83	61.31	59.75	61.31	
$\sigma_{P}^{2}$	83.87	84.83	83.87	86.49	86.68	86.49	
h²	0.2509	1.289E-06	0.2509	0.2911	0.00004	0.2911	
m ²	-	0.3017	0	-	0.3106	0	
Log L	-1563.37809	-2350.80603	-2353.618399	-1535.572401	-2322.905296	-2325.812714	
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			5.6248			5 819	

Table A19. Estimates for Model 2,  $\sigma^2_M$  set to zero. Significance test – Nubian

$\chi^2_{1,0005} = 7.88$		[from Steel and Torric. 1980]
$\chi^2_{1,0050} = 3.84$	$\chi^2_{1,0.025} = 5.02$	$\chi^2_{1,0010} = 6.63$
$\chi^2_{1,0500} = 0.455$	$\chi^2_{10250} = 1.32$	$\chi^2_{1,0100} = 2.71$
$\chi^2_{1,0950} = 0.00393$	$\chi^2_{1,0.900} = 0.0158$	$\chi^2_{1,0.750} = 0.102$
$\chi^2_{1,0.995} = 0.0000393$	$\chi^2_{1,0990} = 0.000157$	$\chi^2_{1,0.975} = 0.000982$

		Basic N	Model 1		Basic Model 2		
Milk	Model 1	Model 2 (H _A )	Model 2 (H ₀ ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed	
σ²A	17065.17	10504.21	17065.14	18721.49	12458.69	18727.77	
$\sigma^2_{M}$	•	5607.15	0	-	5230.88	0	
$\sigma_{E}^{2}$	81332.06	82196.83	81332.09	80184.91	80892.71	80179.96	
σ² _P	98397.23	98308.19	98397.23	98906.41	98582.28	<b>98907.73</b>	
h ²	0.173431	0.10685	0.1734	0.0.1893	0.1264	0.1893	
m²	-	0.0570365	0	-	0.0531	0	
Log L	-14171.83008	-17333.56262	-17334.39784	-14123.03122	-17284.88984	-17285.59898	
$-2[H_0 - H_A]$			1.67			1.418	
Fat	Model 1	Model 2 (H _A )	Model 2 (He) $\sigma^2 M$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed	
$\sigma^{2}{}_{A}$	26.82	12.12	26.82	28.84	13.94	28.84	
$\sigma^2_{M}$	-	12.95	0	-	12.98	0	
σ² _E	103.9	104.99	103.9	102.55	103.98	102.55	
σ² _P	130.7 <b>2</b>	130.06	130.72	131.39	130.91	131.39	
h²	0.205166	0.0931989	0.2052	0.2195	0.1065	0.2195	
m²	-	0.0995728	0	-	0.0992	0	
Log L	-8479.511076	-11639.73877	-11642.07884	-8457.259255	-11617.56213	-11619.82702	
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			4.68			4.53	
Protein	Model 1	Model 2 (H _A )	Model 2 (H ₀ ) $\sigma^2_{\rm M}$ fixed	Model 1	Model 2	Model 2 $\sigma^2_{\rm M}$ fixed	
$\sigma^{2}{}_{A}$	14.34	4.76	14.34	15.83	6.20	15.83	
σ² _M	-	8.35	0	-	8.29	0	
$\sigma_{E}^{2}$	76.09	76.77	76.09	75.11	75.84	75.11	
σ² _P	90.43	89.88	90.43	90.94	90.34	90.94	
h²	0.158552	0.0529395	0.1586	0.1741	0.0687	0.1741	
m²	-	0.0929023	0	-	0.0918	0	
Log L	-8171.690092	-11332.12783	-11334.25785	-8151.420213	-11312.00710	-11313.98797	
$-2[\mathbf{H}_0 - \mathbf{H}_A]$			4.26			3.96	

$\chi^2_{1,0.995} = 0.0000393$	$\chi^2_{1,0.990} = 0.000157$	$\chi^2_{1,09^{-5}} = 0.000982$
$\chi^2_{1,0950} = 0.00393$	$\chi^2$ 1.0.900 = 0.0158	$\chi^2_{1,0.750} = 0.102$
$\chi^2_{1,0.500} = 0.455$	$\chi^2_{10,250} = 1.32$	$\chi^2_{1,0,100} = 2.71$
$\chi^2_{1,0\ 050} = 3.84$	$\chi^2$ 1,0.025 = 5.02	$\chi^2_{1,0.010} = 6.63$
$\chi^2_{1,0005} = 7.88$		[from Steel and Torrie. 1980]

Breed / Trait							
Alpine	SSR (F)	SSR (R)	<b>€</b> ² E(F)	€ ² _A (F)	đſ	residual df	F _c -value
Milk	30543419.84	31228898.39	60521.12092	26329.40902	8	508	0.964209946
Fat	45806.23777	48199.89475	93.41064873	21.513487	8	508	2.548420904
Protein	31311.69951	32584.74617	63.14873288	16.97686676	8	508	1.965058269
Toggenburg	SSR (F)	SSR (R)	<b>€</b> ² z (F)	€ ² _A (F)	đf	residual df	F _c -value
Milk	46370310.44	48098989.52	96604.72943	22268.37369	8	480	1.81777776
Fat	48801.64868	49419.03832	101.6700898	24.48886366	8	480	0.611718018
Protein	37055.89465	38550.02347	77.19976538	17.95579254	8	480	1.962745073
Samen	SSR (F)	SSR (R)	● ² _E (F)	€ ¹ _A (F)	df	residual df	F _c -value
Millik	ſ	f	f	f	ſ	f	ſ
Fat	38380.85248	39490.02438	143.7485112	18.27790747	8	267	0.855702969
Protein	30017.79798	31700.83237	112.4262096	12.00739238	8	267	1.690695241
Nubian	SSR (F)	SSR (R)	€ ¹ L (F)	€ ² _A (F)	dſ	residual df	Fevalue
Milk	10446573.36	11048969.51	41786.29342	17630.37503	8	250	1.267313041
Fat	21936.4908	23055.27553	87.7459632	75.72180305	8	250	0.85550867
Protein	15328.66025	16209.43941	61.31464098	25.17604029	8	250	1.272939389
All breeds	SSR (F)	SSR (R)	€ ¹ E(F)	€ ¹ _A (F)	đſ	residual df	F _c -value
Milk	137036018	139647150.8	80184.91399	18721.49429	8	1709	3.300004577
Fat	175261.3301	178398.8928	102.5518789	28.84236029	8	1709	2.984874679
Protein	128363.1557	130651.1136	75.11009697	15.83008755	8	1709	3.144866475

#### Table A21. Sums of Squares for residual, additive variance, error variance and calculated F-values

 $F_c$  - value = calculated F - value

 $F_{0\,100,8,\infty} = 1.67$   $F_{0\,050,8,\square} = 1.94$   $F_{0\,025,8,\infty} = 2.19$   $F_{0\,010,8,\infty} = 2.51$  $F_{0\,005,8,\infty} = 2.74$ 

[from Steel and Torrie, 1980]